

Separation of internal and forced variability of climate using a U-Net

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Key Points:

- We present a new method to separate the forced and internal variability of the surface air temperature.
- We utilise a U-Net trained with global climate models outputs and implement a noise to noise methodology to eliminate internal variability.
- The results are assessed through the utilisation of very large ensemble simulations of two distinct climate models.

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Abstract

The internal variability pertains to fluctuations originating from processes inherent to the climate component and their mutual interactions. On the other hand, forced variability delineates the influence of external boundary conditions on the physical climate system. A methodology is formulated to distinguish between internal and forced variability within the surface air temperature. The noise-to-noise approach is employed for training a neural network, drawing an analogy between internal variability and image noise. A large training dataset is compiled using surface air temperature data spanning from 1901 to 2020, obtained from an ensemble of Atmosphere-Ocean General Circulation Model (AOGCM) simulations. The neural network utilized for training is a U-Net, a widely adopted convolutional network primarily designed for image segmentation. To assess performance, comparisons are made between outputs from two single-model initial-condition large ensembles (SMILEs), the ensemble mean, and the U-Net's predictions. The U-Net reduces internal variability by a factor of four, although notable discrepancies are observed at the regional scale. While demonstrating effective filtering of the El Niño Southern Oscillation, the U-Net encounters challenges in areas dominated by forced variability, such as the Arctic sea ice retreat region. This methodology holds potential for extension to other physical variables, facilitating insights into the enduring changes triggered by external forcings over the long term.

Plain Language Summary

To comprehensively grasp future climate change, it becomes imperative to differentiate between forced variability and internal climate variability. Internal variability refers to the climate's variations driven by the chaotic nature of geophysical fluids. Conversely, forced variability denotes changes prompted by external forcings, predominantly alterations in radiative forcing, primarily due to anthropogenic activities. Here, a novel approach is introduced for filtering internal variability through the utilisation of a convolutional neural network. This neural network is trained using a noise-to-noise methodology, targeting the filtration of internal variability from surface air temperature outputs of climate models or observational data. Internal variability is treated analogously to noise within an image, which is removed to restore the "true image," corresponding to forced variability in our case. This method capitalises on the data generated by state-of-the-art climate models through the coupled model intercomparison project (CMIP). To val-

47 idate this methodology, we assess its performance using very large ensembles of climate
48 model simulations, enabling precise estimation of forced variability. Our findings demon-
49 strate a reduction in internal variability by a factor of four, accompanied by notable re-
50 gional variations.

51 **1 Introduction**

52 The phenomenon of climate warming is characterized by an elevated surface air tem-
53 perature, notably reaching a pivotal juncture during the latter half of the twentieth cen-
54 tury (Eyring et al., 2021). Nevertheless, the observed anomalies in surface air temper-
55 ature arise from a dual spectrum of variabilities. The first source of variability is due to
56 the effect of the external forcings, such as the increase in the greenhouse gases concen-
57 tration, the variations of concentration in anthropogenic and natural aerosols, the fluc-
58 tuations in solar variability or volcanic eruptions and the land-use changes. The related
59 variability is designated as the forced variability. The second source of variability is com-
60 ing from processes internal to the atmosphere, oceans, cryosphere and land or the inter-
61 actions between them (Cassou et al., 2018). Subsequently, this form of variability is re-
62 ferred to as 'internal variability,' encapsulating its inception within the climate system
63 and its persistence even without alterations in external forcings. Despite the overarch-
64 ing dominance of forced variability in shaping the broad-scale and long-term trajectory
65 of surface air temperature across the 1900-2020 timeframe (Deser et al., 2012; Kay et
66 al., 2015), a comprehensive understanding of the distinct contributions of internal and
67 forced variability remains elusive. Internal variability takes center stage in briefer tem-
68 poral scales and smaller spatial dimensions. For instance, the leading mode of internal
69 variability in global air surface temperature manifests as the El Niño Southern Oscilla-
70 tion (ENSO), characterized by significant anomalies in the equatorial Pacific Ocean, ac-
71 companied by distant teleconnections, and a prevailing cycle spanning two to seven years
72 (Wang & Picaut, 2004). Additionally, the interdecadal Pacific variability (Newman et
73 al., 2016) and the Atlantic Multidecadal variability (Zhang et al., 2019) wield the capac-
74 ity to influence climate dynamics across the decadal to multidecadal spectrum. A no-
75 table example involves the deceleration in the global warming rate experienced during
76 2002-2012, commonly referred to as the global warming hiatus, which has been robustly
77 linked to Interdecadal Pacific Variability (Meehl et al., 2013; Kosaka & Xie, 2013; Eng-
78 land et al., 2014). Lastly, internal variability exercises influence even over centennial and

79 multi-centennial spans (Jiang et al., 2021; S. Li & Huang, 2022) exerting substantial im-
80 pact on trends within the 1900-2015 interval (Bonnet et al., 2022).

81 The distinction between forced variability and internal variability is essential for
82 conducting detection and attribution studies, enabling accurate estimation and simula-
83 tion of the climate’s reaction to alterations in radiative forcing. Moreover, this differen-
84 tiation aids in recognizing and comprehending internal climate variability. Nevertheless,
85 the availability of instrumental observations is limited to the period since 1850, and the
86 relatively brief duration of these observations presents challenges in effectively and con-
87 fidently discerning internal variability.

88 For identifying both internal and forced variability, linear trends (Swart et al., 2015;
89 Vincent et al., 2015) or quadratic trends (Enfield & Cid-Serrano, 2010) have been em-
90 ployed to characterize forced variability. However, linear or quadratic trends inadequately
91 capture the temporal evolution of temperature, particularly failing to account for the abrupt
92 cooling subsequent to significant volcanic eruptions, which hold significant climate im-
93 pact (Schmidt et al., 2018). Additional approaches include the application of Empiri-
94 cal Orthogonal Functions (EOF) analysis (Parker et al., 2007), low-frequency pattern
95 filtering (Wills et al., 2020), and linear inverse models (Marini & Frankignoul, 2014). These
96 techniques deconstruct forced variability into a combination of modes featuring distinct
97 patterns and corresponding time series. Regression analysis of the global mean surface
98 temperature (GMST) has also been employed, although this may inadvertently estab-
99 lish misleading links between the Atlantic and Pacific basins (Frankignoul et al., 2017;
100 Deser & Phillips, 2023). However, a comprehensive and systematic examination of these
101 methodologies remains notably absent.

102 Climate model simulations have been employed to overcome the limitations of sparse
103 observation sampling. Conducting an ensemble of climate model simulations with diverse
104 initial conditions enables estimation of forced variability via the ensemble mean. This
105 approach effectively mitigates the variance linked to internal variability by a factor of
106 n , where n signifies the ensemble’s size (Harzallah & Sadourny, 1995; Hawkins & Sut-
107 ton, 2009; Ting et al., 2009; Solomon et al., 2011; Deser et al., 2014; Frankcombe et al.,
108 2015). As a result, modeling centers have undertaken substantial ensembles with over
109 20 or 30 ensemble members (Jeffrey et al., 2013; Rodgers et al., 2015; Sun et al., 2018;
110 Deser et al., 2020). These large ensembles are commonly referred to as Single-Model Initial-

111 Condition Large Ensembles (SMILE; Deser et al. (2020)). Multiple SMILE initiatives
112 have been undertaken using models such as CCSM3 (Collins et al., 2006), CCSM4 (Gent
113 et al., 2011), CESM (Kay et al., 2015), MPI-ESM (Maher et al., 2019), FGOALS-g3 (Li
114 et al., 2020), CanESM2 (Chylek et al., 2011), and IPSL-CM6A-LR (Bonnet et al., 2021),
115 among others. This offers a valuable dataset for crafting methodologies dedicated to the
116 disentanglement of forced and internal variability. Notably, employing members of a large
117 ensemble model as surrogate observations allows for a comparison of results with the en-
118 semble mean. Differences primarily mirror residual internal variability or limitations in-
119 herent in the method.

120 Nevertheless, the forced variability estimated through an ensemble mean remains
121 contingent upon the specific climate model employed. These climate models carry sub-
122 stantial uncertainties, particularly in terms of their climate sensitivity (Sherwood et al.,
123 2020), often attributed to factors like uncertain cloud retroaction which significantly im-
124 pact equilibrium climate sensitivity (Zelinka et al., 2016). Additionally, significant un-
125 certainties surround historical emissions and the linked radiative forcing from aerosols
126 (Menary et al., 2020; C. J. Smith & Forster, 2021). Moreover, the internal variability ex-
127 hibited by different models also varies significantly (Parsons et al., 2020).

128 Several methodologies have been devised to harness data from diverse climate mod-
129 els, as employing a multi-model approach holds the potential to alleviate the uncertain-
130 ties inherent in individual climate models. Multi-model ensemble means are widely adopted
131 for estimating the forced signal (Steinman et al., 2015). Notably, techniques such as the
132 signal-to-noise-maximizing empirical orthogonal functions (Ting et al., 2009; Wills et al.,
133 2020) and the discriminant analysis and maximization of the average predictability time
134 (DelSole et al., 2011) have been put forth to extract forced variability with superior ef-
135 ficacy compared to ensemble means. Furthermore, scaling techniques that adjusts the
136 forced signal from models using observational data have been proposed. Among these
137 methodologies are fingerprinting methods grounded in linear regression, commonly ap-
138 plied for detecting and attributing climate change with a unified forcing that encapsu-
139 lates the influence of all external forcings (Hasselmann, 1993; Allen & Tett, 1999; Allen
140 & Stott, 2003). More recently, the use of scaling factors was also proposed by Frankcombe
141 et al. (2015).

142 This paper introduces an alternative approach to distinguishing internal and forced
143 variability using climate model data, employing a non-linear method that takes into ac-
144 count the spatio-temporal data covariances. This method is rooted in a neural network
145 trained on data from Atmosphere-Ocean General Circulation Models (AOGCMs). Among
146 the areas where neural networks have excelled is image analysis (Egmont-Petersen et al.,
147 2002). One of the prominent applications of neural networks in image processing is im-
148 age denoising, involving the elimination of noise from an image to restore its true form
149 (Ilesanmi & Ilesanmi, 2021; Tian et al., 2020). In this context, internal variability is treated
150 as noise. It is demonstrated that machine learning image denoising methodologies can
151 subsequently isolate forced variability. The internal variability is eliminated, leaving be-
152 hind a quantifiable residue. This method leverages the temporal and spatial information
153 inherent in climate models to establish the weights and biases of a neural network. With
154 these parameters in place, the neural network is also employed with observations to delve
155 into and attribute the progression of climate change since 1905 to 2016. To the best of
156 our knowledge, this represents the pioneering application of a dedicated neural network
157 for the purpose of disentangling internal and forced variability.

158 The structure of this paper is as follows: Section 2 outlines the data utilized. Sec-
159 tion 3 introduces the method anchored in a neural network. Section 4 assesses the method's
160 performance. In Section 5, the neural network method is applied to observations. Lastly,
161 Section 6 offers the conclusion and discussion.

162 **2 Data**

163 **2.1 Observations**

164 The gridded monthly Surface Air Temperature anomaly (SAT) from 1901 to 2020,
165 as provided by GISS Surface Temperature Analysis version 4 (GISTEMP; Hansen et al.
166 (2010); Lenssen et al. (2019)), is employed in this study. GISTEMP amalgamates me-
167 teorological station data over land (NOAA GHCN v4) with sea surface temperature (SST)
168 estimates from ERSST v5. This data is available on a consistent $2^\circ \times 2^\circ$ grid. The monthly
169 values are aggregated to calculate annual means, and the SAT anomalies are determined
170 using the reference period 1950-2014.

2.2 Climate model simulations

The monthly SAT data is sourced from historical simulations within the Coupled Model Intercomparison Project Phase 5 (CMIP5; Taylor et al. (2012)) and the Coupled Model Intercomparison Project Phase 6 (CMIP6; (Eyring et al., 2016)), along with several Single-Model Initial-Condition Large Ensembles (SMILEs) from distinct models: MPI-ESM (Maher et al., 2019), CSIRO-Mk3-6-0 (Collier et al., 2011), EC-Earth (Döscher et al., 2021), and FGOALS-g3 (Li et al., 2020). For the historical simulations, spanning 1901 to 2005 (2014 for CMIP6), all external forcings are integrated. These forcings encompass the effects of historical greenhouse gas concentrations, anthropogenic and natural aerosols, stratospheric ozone, solar activity, and land-use changes. Each climate model delivers multiple realizations referred to as ensemble members, generated through distinct initial conditions. From 2005 (2014 for CMIP6) until 2020, the outputs under the pessimistic Representation Concentration Pathway 8.5 (RCP8.5) scenario for CMIP5 (Van Vuuren et al., 2011) and the intermediate Shared Socio-economic Pathway 2 4.5 (SSP2-4.5) for CMIP6 (Tebaldi et al., 2020) are employed. These simulations utilize socio-economic assumptions to project future external forcing patterns. Additionally, several SMILEs are incorporated, employing distinct historical forcings or scenario simulations of CMIP5 or CMIP6 (elaborated in Table S3). While minor differences are anticipated in external forcing between CMIP5 and CMIP6 simulations, notable uncertainties arise in aerosol emissions (C. J. Smith et al., 2020; Fyfe et al., 2021). Modest differences may also emerge between the RCP8.5 (strong) and SSP2-4.5 (moderate) scenarios, particularly until 2020, where actual forcings mirror observed forcings to a considerable extent (Masson-Delmotte et al., 2021).

The count of members accessible for scenario simulations is fewer compared to the historical counterparts. Therefore, we extended the outputs from historical experiments using the scenario ensemble member of the same model with the same number identification. In case the number identification is lacking, we select randomly an scenario ensemble member of the same climate model.

All monthly data are aggregated into annual means. Subsequently, the SAT anomalies are computed for each ensemble member using 1950-2014 as a reference period. This furnishes a multi-model ensemble comprising 801 members derived from 47 AOGCMs. Subsequently, the concatenated historical and scenario members are harnessed within

203 the 1901-2020 timeframe. All model data is regridded using bilinear interpolation on the
204 horizontal grid from GISTEMP. The details pertaining to the climate model names, en-
205 semble sizes, and the names of the employed scenario simulations are elucidated in Tabs.
206 S1, S2, and S3.

207 **2.3 Validation of the data set**

208 The forced variability simulated within the multi-model ensemble is succinctly ex-
209 amined for two specific data subsets. We investigate the MPI-ESM and FGOALS-g3 cli-
210 mate models from SMILE, as they have a very large size of 100 and 115 members, re-
211 spectively, which largely exceed the size of other model ensembles. Anticipatedly, the es-
212 timated forced variability derived from the ensemble mean for each of these models is
213 expected to be accurate, as the reduction in variance attributed to internal variability
214 reaches 100 and 115, respectively. For instance, Deser et al. (2012, 2014) demonstrated
215 that identifying regional climate responses on time scales of several decades may neces-
216 sitate between 10 to 40 members. Specifically, to detect a change in SAT between the
217 decades 2005-2014 and 2028-2037 on a global scale, the use of 3 to 6 members is requi-
218 site. This requirement can surge beyond 10 for local analyses such as in North Amer-
219 ica. Subsequently, the data originating from these two models is subsequently employed
220 to appraise the outcomes of the neural network model in section 4.1.

221 We utilize the ensemble mean to characterize the forced variability and employ the
222 standard deviations from the ensemble members for evaluating the internal variability.
223 Figure 1 illustrates the standard deviation of the SAT deviation from the ensemble mean
224 for FGOALS-g3 and MPI-ESM. The variability in SAT is more pronounced over land
225 surfaces ($\sim 0.3^\circ\text{C}$) compared to oceans ($\sim 0.1^\circ\text{C}$), consistent with the lower thermal in-
226 ertia of land. Notably, substantial variability (ranging from approximately 1.5°C to 2.5°C)
227 is observed over regions coinciding with the sea ice edge, such as the Bering Sea and Nordic
228 Seas in the Northern Hemisphere, as well as the Amundsen and Weddell Seas in the South-
229 ern Hemisphere. Additionally, a marked variability is observed in the equatorial Pacific
230 Ocean, with a standard deviation of 0.8°C , and this variability is more prominent in MPI-
231 ESM compared to FGOALS-g3. A localized peak of variability is situated over the sub-
232 polar North Atlantic, especially notable for FGOALS-g3 (reaching up to 2°C). These out-
233 comes coherently reflect a significant internal variability stemming from extratropical weather
234 fluctuations over land surfaces, exhibiting local maxima around regions adjacent to the

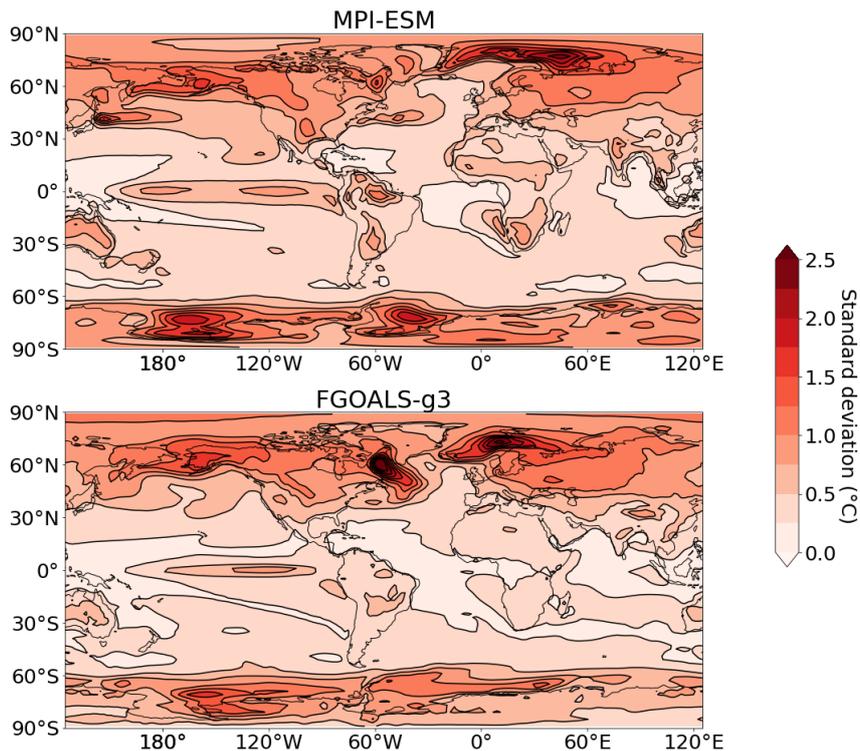


Figure 1. Standard deviation of the SAT deviations from the ensemble mean for (top) MPI-ESM and (bottom) FGOALS-g3.

235 sea ice edge. Moreover, the variability observed in the equatorial Pacific mirrors the phe-
 236 nomenon of El Nino Southern Oscillation (Neelin et al., 1998).

237 The forced variability is estimated through the ensemble mean of each model. Sub-
 238 sequently, the multi-model mean (MMM) is computed by averaging the ensemble means
 239 across all models, ensuring equal weight for each model. Nonetheless, MPI-ESM and FGOALS-
 240 g3 are excluded from this computation, as the intention is to later compare them to the
 241 MMM. To assess the prominent impact of greenhouse gas forcing, Figure 2 (a, c, e) il-
 242 lustrates the ensemble mean SAT anomaly for MPI-ESM, FGOALS-g3, and the MMM
 243 throughout the 2010-2020 interval. Furthermore, Figure 2 (b, d, f) presents the tempo-
 244 ral standard deviation of the ensemble means across the period from 1901 to 2020. As
 245 anticipated, all climate models project more substantial warming over land (up to 0.8°C)
 246 than over oceans (approximately 0.3°C). Notably, the Arctic exhibits an amplification
 247 of global warming, with warming exceeding 2°C north of 60°N. The MMM showcases an
 248 average warming of 0.8°C for the 2010-2020 period, surpassing MPI-ESM (0.64°C) and

249 FGOALS-g3 (0.69°C). This aligns with the comparatively lower equilibrium climate sen-
250 sitivity (ECS) of these two models (3.6°C for MPI-ESM and 2.8°C for FGOALS-g3) when
251 compared to other models employed in this study (Zelinka et al., 2020). Within the sub-
252 polar Atlantic, the SAT anomalies exhibit a minimum, with negative temperatures anom-
253 alies observed in FGOALS-g3 over the Labrador Sea, or in MPI-ESM over the subpolar
254 gyre. This phenomenon, known as the North Atlantic warming hole (Keil et al., 2020),
255 is associated with a deceleration of the Atlantic meridional overturning circulation (He
256 et al., 2022). It is worth noting that such a minimum is less pronounced in the MMM,
257 presumably due to considerable uncertainties regarding the precise location of this warm-
258 ing hole and the linked processes. An equivalent spatial pattern can be derived using stan-
259 dard deviations, revealing values of approximately 0.3°C for the majority of global re-
260 gions and higher values over land ($\sim 0.6^{\circ}\text{C}$). Grid points located north of 60° also exhibit
261 elevated values, peaking at around 2°C in the Barents Sea for MPI-ESM or the Labrador
262 Sea for FGOALS-g3.

263 The forced variability exhibited by MPI-ESM and FGOALS-g3 diverges from that
264 of the MMM, revealing a comparatively weaker global warming trend and standard de-
265 viation pattern. This divergence is particularly evident north of 60°N , where the warm-
266 ing exhibits greater amplification (refer to Fig. 2), amounting to 1.54°C for MPI-ESM
267 and 1.45°C for FGOALS-g3. Local variations are also observed in regions such as the Labrador
268 Sea, Barents and Kara Sea, the Canadian archipelago, and the Bering Sea in the case
269 of FGOALS-g3. Notably, MPI-ESM similarly presents notable differences in the Barents
270 Sea. These discrepancies may arise from biases related to sea ice representation. Specif-
271 ically, FGOALS-g3 depicts an excessive extent of Arctic sea ice (Li et al., 2020), which
272 in turn leads to inaccuracies in simulating the location of the sea ice edge. This discrep-
273 ancy can account for spurious SAT variability attributed to the misplaced sea ice edge
274 within the Labrador Sea. The mean standard deviation of the ensemble mean registers
275 as 0.34°C for MPI-ESM and 0.43°C for FGOALS-g3, exceeding the mean standard de-
276 viation of the SAT deviations of the members to the ensemble mean which is of 0.51°C
277 for MPI-ESM and 0.46°C for FGOALS-g3. This underscores that the internal variabil-
278 ity is marginally more pronounced than the forced variability.

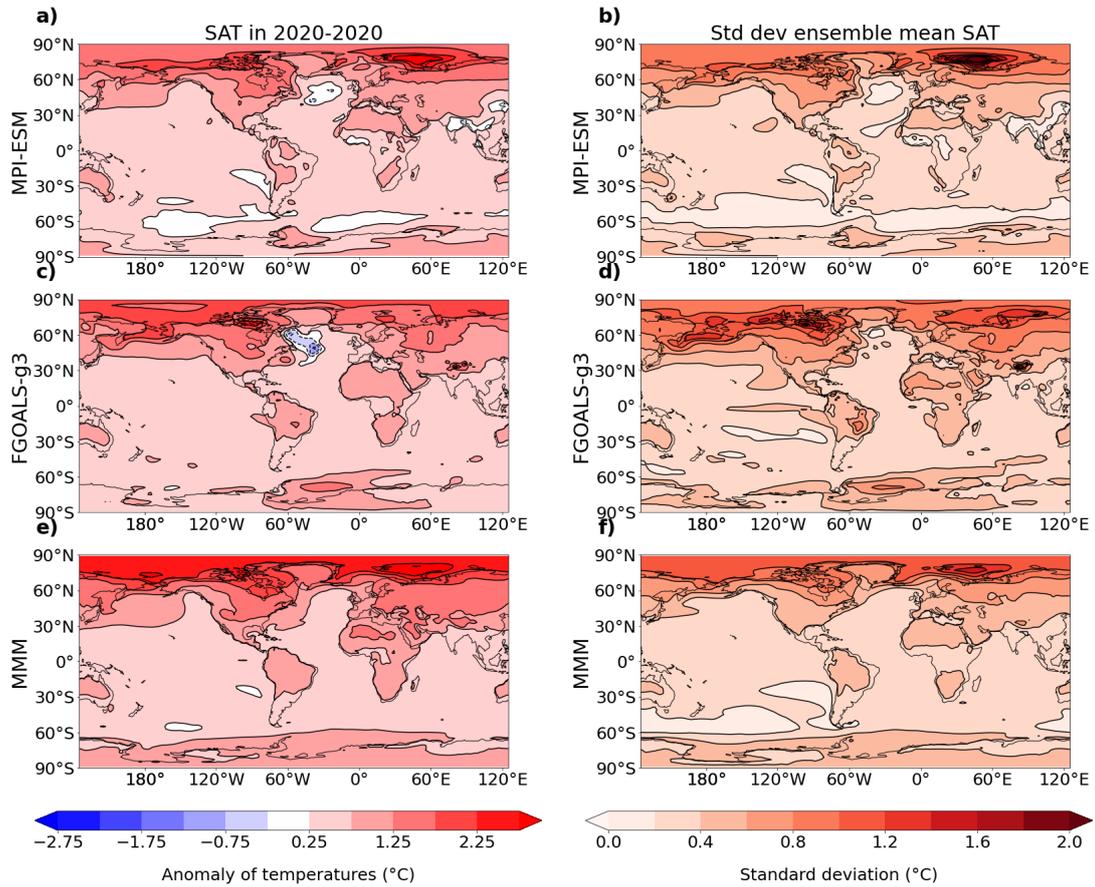


Figure 2. a) Ensemble mean of the air surface temperature ($^{\circ}\text{C}$) in MPI-ESM in 2010-2020. c) Same as a) but for FGOALS-g3. e) Same as a) but for the MMM. b) Standard deviation of the ensemble mean surface air temperature ($^{\circ}\text{C}$) in 1901-2020 for MPI-ESM. d) Same as b) but for FGOALS-g3. f) Same as b) but for the MMM.

3 Methods

3.1 Neural network

We design a neural network to remove the internal variability from the SAT. The input data is structured with dimensions (120, 90, 180), corresponding to time spanning from 1901 to 2020, latitude, and longitude, respectively. On the other hand, the output holds dimensions of (112, 90, 180), encompassing the years 1905 to 2016, while maintaining the latitude and longitude dimensions intact. Notably, the output’s temporal span is truncated compared to the input, by excluding the initial and final four years. This reduction addresses the substantial uncertainty typically observed at the dataset’s endpoints, an aspect that will be elaborated upon later.

A neural network’s characteristics are shaped by its hyperparameters, which dictate both its architecture and training process. Our approach involves utilizing three distinct datasets, each composed of input and desired output pairs. The training dataset serves the purpose of establishing the neural network’s weights and biases. Meanwhile, the validation dataset comes into play for estimating the hyperparameters. Finally, the test dataset is employed to assess the neural network’s performance.

3.2 Constitution of the database

To construct the training dataset, we adapt a noise-to-noise methodology originally introduced in Lehtinen et al. (2018). This approach was initially designed to train a neural network in denoising images. In this method, the network is exclusively trained on noisy images depicting various objects. Each object has more than one noised image depicting it. In the noise to noise method, we create an input/output training database that comprises pairs of noisy image combinations for identical objects. It’s essential to note that the network cannot effectively learn to transform a random noise realization into another. Instead, the configuration is designed to approximate the mathematical expectation of all noisy images associated with the same object, culminating in an estimate that closely resembles the noise-free image.

For our application, we consider the forced spatio-temporal SAT anomalies from each climate model as distinct objects. These anomalies, inherent to each member, can be likened to noisy images, where the internal variability introduces the noise compo-

309 ment. The ensemble members' mathematical expectation equates to the forced variabil-
310 ity, which can be approximated through the ensemble mean.

311 To create the training dataset, we follow a procedure wherein we compute pairs of
312 members for each climate model, except for MPI-ESM, FGOALS-g3, and MIROC6, which
313 are reserved for testing and validation purposes. Adopting an approach similar to Lehtinen
314 et al. (2018), we augment the dataset by introducing the ensemble mean of the climate
315 model's members as an additional member. This inclusion serves to expedite the train-
316 ing process without introducing any other influences. In this process, each pair of mem-
317 bers becomes an input/output pair. If we denote the number of ensemble members ob-
318 tained from a specific climate model as n , this approach yields $n(n+1)$ input/output
319 pairs per model. By accumulating such pairs from all models, the resulting training dataset
320 primarily comprises simulations characterized by the most extensive ensemble sizes (namely
321 IPSL-CM6A-LR, CanESM5, CNRM-CM6-1, and ACCESS-ESM1-5).

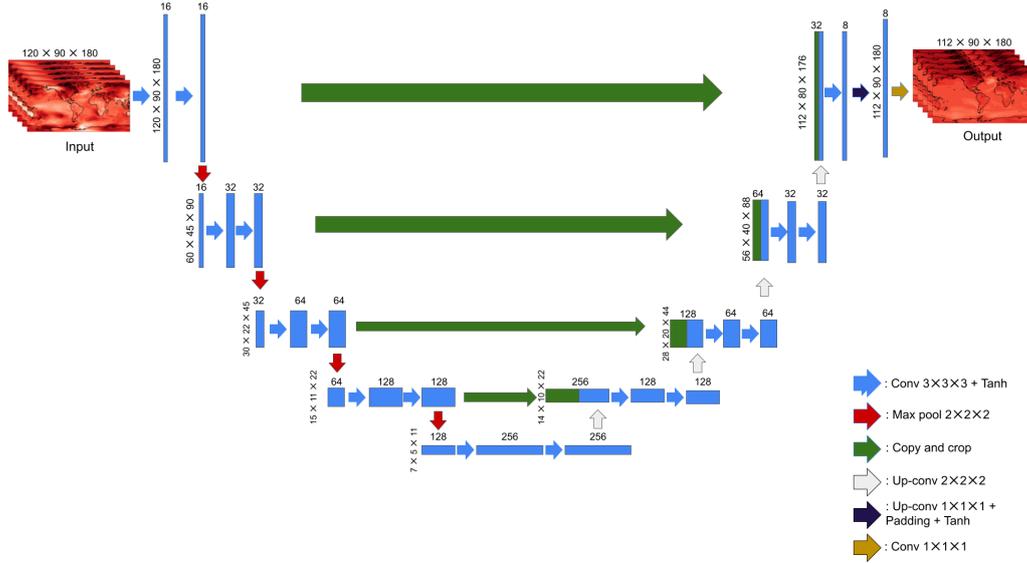
322 To create the validation set, we employ the ensemble simulation data from the MIROC6
323 model, which ranks as the third-largest ensemble in terms of size (with $n = 50$ mem-
324 bers). For this purpose, we designate the ensemble members as inputs, while the ensem-
325 ble mean spanning the period from 1905 to 2016 serves as the desired output.

326 To form the test dataset, we draw upon data derived from the FGOALS-g3 and
327 MPI-ESM models, leveraging their extensive ensemble sizes of $n = 110$ and $n = 100$
328 respectively. Subsequently, we proceed to make comparisons between the outputs of the
329 neural network obtained from ensemble members and their corresponding ensemble means
330 for both of these models.

331 The conclusions drawn from these tests and validation processes may exhibit some
332 dependence on the specific model being analyzed, as alternative models could yield vary-
333 ing outcomes. Nevertheless, this approach has been chosen due to its simplicity and its
334 potential to mitigate the impact of any remaining internal variability.

335 **3.3 U-Net**

336 Convolutional neural networks (CNNs, Yamashita et al. (2018)) constitute a cat-
337 egory of non-linear neural networks, notably applied in tasks related to imagery (O'Shea
338 & Nash, 2015). A distinctive attribute of CNNs is their utilization of convolutional lay-
339 ers, which incorporate a trainable kernel that slides across the input data.



prising 112 time steps. The neural network is comprised of a total of 5,659,009 trainable parameters.

A batch size of 8 is chosen, and the optimization process employs the Adam optimizer with a learning rate of 0.001. To ensure proper application of the CNN to the data, padding is introduced. This involves extending the image by appending zero values at its edges. For the longitudinal dimension, which is periodic, the zero padding only results in a slight discontinuity at 180°E, the edge of the data. Indeed, due to the nature of convolutional layers, a U-Net has more difficulty processing information located at the edge of the data. This is the reason why we excluded the initial and final four years (1901-1904 and 2017-2020) in the U-Net's outputs. The chosen cost function is the root mean squared error (RSME), calculated using an area-weighted mean of the gridded data.

The validation dataset is utilized to determine the optimal values for two key hyperparameters: the number of epochs and the number of filters used in the convolutional layers. The term "number of filters" pertains to the thickness of the convolutional layers. The number of epochs refers to how many times the training dataset is processed during the training phase. These hyperparameters are selected to minimize the root mean squared error (RMSE) using the validation dataset. Examination of the validation RMSE for different values of epochs and layer thickness reveals a consistent pattern (see Fig. S1): a significant reduction in RMSE occurs in the initial epochs, followed by a gradual increase. As a result, we settle on a layer thickness of 16 for the first layer (as shown in Fig. 3) and a total of 32 epochs.

3.4 Example

Figure 4 provides an illustrative example featuring two randomly selected ensemble members from MPI-ESM and FGOALS-g3. The comparison focuses on the SAT at the year 2016, depicted in the top panels, as well as the resulting output generated by the neural network in 2016 (centre panels), juxtaposed against the ensemble mean anomaly for the same year (bottom panels). The anticipated impact of elevated greenhouse gas concentrations in 2016 is evident in the SAT of both MPI-ESM and FGOALS-g3 members, which exhibit warm anomalies. However, the internal variability introduces anomalies that surpass those of the ensemble mean in numerous regions, accompanied by some negative anomalies in other areas. To elaborate, an instance of cooling is simulated across

389 the Equatorial Pacific Ocean, possibly linked to a La Niña event in the case of MPI-ESM.
390 The same ensemble member displays cooling over land in equatorial Africa, South-Eastern
391 Asia, and Australia, as well as in extratropical zones like the North Atlantic Ocean and
392 the Weddell Sea. In the example from FGOALS-g3, cold anomalies emerge over the Nordic
393 Seas and the Labrador Sea. Such cooling diverges from the ensemble average, which ex-
394 hibits a relatively uniform warming pattern across the globe, with a more pronounced
395 effect over landmasses. Notably, the Arctic and its environs experience heightened warm-
396 ing compared to other global regions, due to polar amplification. Conversely, minimal
397 warming is observed in the Southern Ocean and the subpolar North Atlantic Ocean, and
398 even a cooling tendency is noted in the Northern Atlantic warming hole.

399 The SAT obtained from the U-Net’s output, utilizing the same ensemble member
400 as input, exhibits a pattern strikingly similar to that of the ensemble mean (compare cen-
401 tre and bottom panels). In both instances, the pattern is relatively uniform, albeit with
402 heightened warming observed over land areas, coupled with an Arctic Amplification phe-
403 nomenon. This suggests that the internal variability—such as the influence of ENSO events
404 or the effects of prolonged weather patterns over continents—has been successfully elim-
405 inated. The regions displaying subdued warming or cooling tendencies are replicated,
406 although the exact positioning and intensity might not precisely match those of the en-
407 semble mean in certain areas, particularly the Southern and subpolar North Atlantic.
408 It’s worth noting a minor discontinuity at 180°E resulting from the padding process.

409 The performance of the method is quantified more systematically in the next sec-
410 tion.

411 **4 The U-Net as an internal variability filter**

412 The U-Net was applied to every member of FGOALS-g3 and MPI-ESM. We then
413 compare the results obtained with the respective ensemble mean of these two climate mod-
414 els.

415 Figures 5a and 5b illustrate the root mean squared error (RMSE) between the out-
416 comes generated by the U-Net and the corresponding ensemble mean for the time pe-
417 riod of 1905-2016. Notably, the discrepancies in U-Net’s predictions are not uniformly
418 distributed across space. The RMSE values fall within the range of 0.05°C to 0.5°C. The
419 discrepancies generally remain below 0.2°C in tropical regions, except for instances over

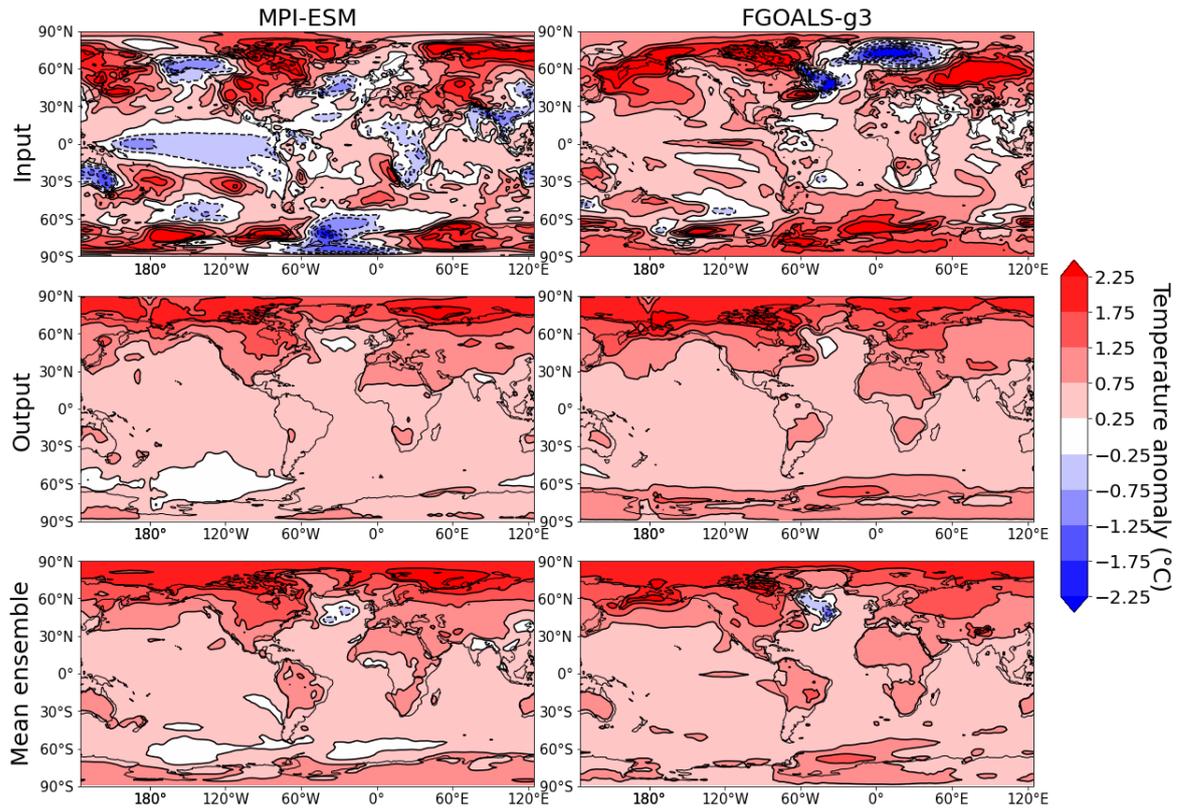


Figure 4. (First column) Anomalies of SAT in a randomly chosen member of MPI-ESM, the associated U-Net output and ensemble mean in 2016. (Second column) Same as the first column but for a randomly chosen ensemble member for FGOALS-g3.

420 Western Africa in the MPI-ESM model. In contrast, the largest errors are concentrated
421 in polar areas, encompassing the Nordic Seas, Labrador Sea, and Bering Sea. Moreover,
422 sizable errors are also evident over the Southern Ocean and the continents of the North-
423 ern Hemisphere situated above 45°N. These high-error regions correspond to locales char-
424 acterized by substantial internal variability (refer to Figure 1). Nevertheless, it is note-
425 worthy that the errors produced by the U-Net are approximately five times smaller than
426 the actual internal variability. Between the years 1996 and 2016, both ensemble results
427 exhibit a warming trend that is roughly 0.1°C lower in the U-Net results when compared
428 to the ensemble mean (as observed in Figs. 5cd). This difference is indicated by the nearly
429 consistent negative divergence situated between latitudes 45°N and 45°S.

430 The prevailing trend of systematic underestimation is, however, disrupted by an
431 exception involving the subpolar Atlantic and the Southern Ocean, where an overesti-
432 mation of warming is observed. This overestimation is particularly conspicuous in the
433 FGOALS-g3 model, with warming anomalies extending to approximately 1°C over the
434 Labrador Sea and 0.5°C over the Bering Sea. This divergence from the ensemble mean
435 highlights the limited capacity of the neural network to accurately predict forced changes
436 within the subpolar North Atlantic, which is a region that exhibits inconsistent surface
437 temperature shifts across models (Swingedouw et al., 2021). The neural network’s per-
438 formance is restricted due to this discrepancy among models, which hampers its abil-
439 ity to discern the specific features of each climate model. For example, in the case of FGOALS-
440 g3, the extensive anomalies in the Labrador and Bering Seas are not mirrored in the multi-
441 model mean (see Figure 2). It’s also plausible that the substantial internal variability
442 observed in these regions poses a challenge for accurate removal by the neural network
443 (refer to Figure 1). This underestimation extends to the continents, with a greater im-
444 pact on South America, Africa, and Australia in the tropics, as well as North America
445 and Northern Siberia in boreal regions. The degree of underestimation reaches 0.15°C
446 for MPI-ESM and 0.13°C for FGOALS-g3 in these regions.

447 Figures 6c and 6d illustrate the temporal evolution of the global surface air tem-
448 perature (GSAT) for both the MPI-ESM and FGOALS-g3 models, before and after ap-
449 plying the U-Net correction. The range of data variability is portrayed by a 90% con-
450 fidence interval assuming a Gaussian distribution. The forced variability’s temporal trend
451 extracted via ensemble mean (depicted by the red line) is effectively captured by the U-
452 Net outputs (represented by the blue line and blue shading).

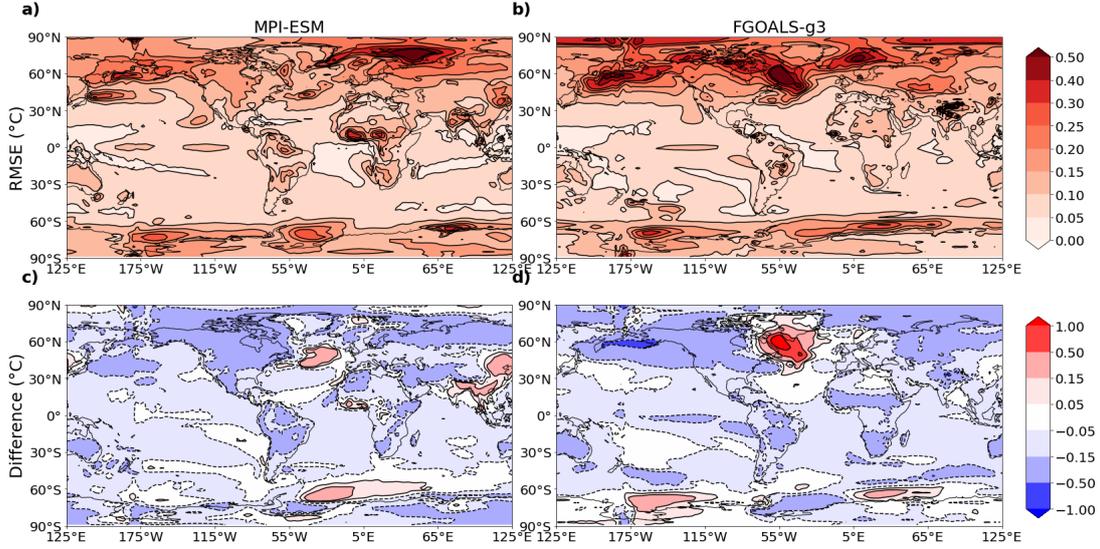


Figure 5. a) Root mean square difference of the surface air temperature, in $^{\circ}\text{C}$, between the outputs of the U-Net and the mean ensemble in MPI-ESM, calculated across the members and all years in 1905-2016 b. b) Same as a) but for FGOALS-g3 c) Difference of the time mean SAT anomaly during 1996-2016, in $^{\circ}\text{C}$, between the mean output of the U-Net and the corresponding ensemble mean, for MPI-ESM. d) Same as c) but for FGOALS-g3

453 From 1905 to 2016, a GSAT rise is observed, aligning with the anticipated shifts
 454 in radiative forcing (Gulev et al., 2021). Additionally, a cooling pattern emerges a few
 455 years subsequent to the significant volcanic eruptions of Agung (1963), El Chichón (1982),
 456 and Pinatubo (1991), a phenomenon accurately estimated by the U-Net. This outcome
 457 aligns with expectations based on climate models incorporating volcanic aerosol emis-
 458 sions. Impressively, the U-Net’s outputs exhibit a marginal spread, reduced approximately
 459 tenfold, indicating a substantial removal of internal variability.

460 Nonetheless, the U-Net results exhibit anomalies with a slightly diminished am-
 461 plitude compared to the ensemble mean. The spread of the U-Net outputs is also ap-
 462 proximately twice as wide at the time series’ beginning and end. The distribution of spa-
 463 tially averaged RMSE values within 90°S - 90°N , comparing all U-Net outputs to the en-
 464 semble mean (depicted in Fig. 6a and 6b as blue histograms), reveals errors of around
 465 0.12°C in MPI-ESM and 0.13°C in FGOALS-g3. Additionally, we examine the RMSE
 466 values when averaging within 60°N - 90°N , as Fig. 5ab suggests that errors are most pro-
 467 nounced in this region (illustrated in Fig. 6ab as red histograms). Errors north of 60°N

468 are approximately twice as substantial as global averages, with an average error of around
469 0.23°C in MPI-ESM and 0.26°C in FGOALS-g3. In Fig. 6ef, the internal variability ob-
470 served when averaging the SAT north of 60°N (as depicted by the red shading) is con-
471 siderable in the raw model outputs (around 0.8°C). The ensemble mean SAT anomalies
472 in this region increase from approximately -1°C in the early twentieth century to about
473 1.2°C in 2010. The temporal evolution of the SAT north of 60°N demonstrates notable
474 similarity between the ensemble mean and the ensemble mean of U-Net outputs, with
475 a roughly 10-fold reduction in spread. However, the amplitude of the anomalies is slightly
476 underestimated, with a reduction of around 0.3°C in negative anomalies in the U-Net
477 output between 1905 and 1930 in MPI-ESM. For FGOALS-g3, the SAT is underestimated
478 by around 0.2°C during 1970-1990.

479 In Figure S2, the quadratic errors between the mean ensemble members and the
480 U-Net output are presented for each year, with global (90°S - 90°N) and north of 60°N av-
481 erages considered for both MPI-ESM and FGOALS-g3. Notably, the RMSE exhibits el-
482 evated values during the initial and final years, characterized by peaks around the years
483 1975-1985 in both models. This pattern underscores the presence of substantial uncer-
484 tainties at the data's onset and conclusion. When applying the 1900-2020 period for the
485 output (without excluding the first and last four years), the errors actually surpass those
486 portrayed in Figure S2, a fact that elucidates the rationale for excluding the endpoints
487 in the ongoing analysis, as detailed in the methods (section 2). Moreover, the notable
488 error peak during 1975-1985 lacks a definitive explanation, although it's plausible that
489 this discrepancy could be linked to uncertainties associated with the implementation of
490 aerosol forcings, notably CMIP5 for MPI-ESM and CMIP6 for FGOALS-g3.

491 The errors exhibited by the U-Net in relation to data from FGOALS-g3 are more
492 prominent compared to those arising from the use of MPI-ESM data. This discrepancy
493 can be attributed to the fact that MPI-ESM's simulated forced variability aligns more
494 closely with the training data's characteristics, on average. Specifically, the training data's
495 forced variability is in line with that of the MMM, and MPI-ESM demonstrates a smaller
496 root mean squared difference from the MMM compared to FGOALS-g3 (as illustrated
497 in Fig. 2).

498 To assess the reduction in internal variability achieved by the U-Net, we can quan-
499 titatively measure the number of ensemble members needed to surpass the U-Net's in-

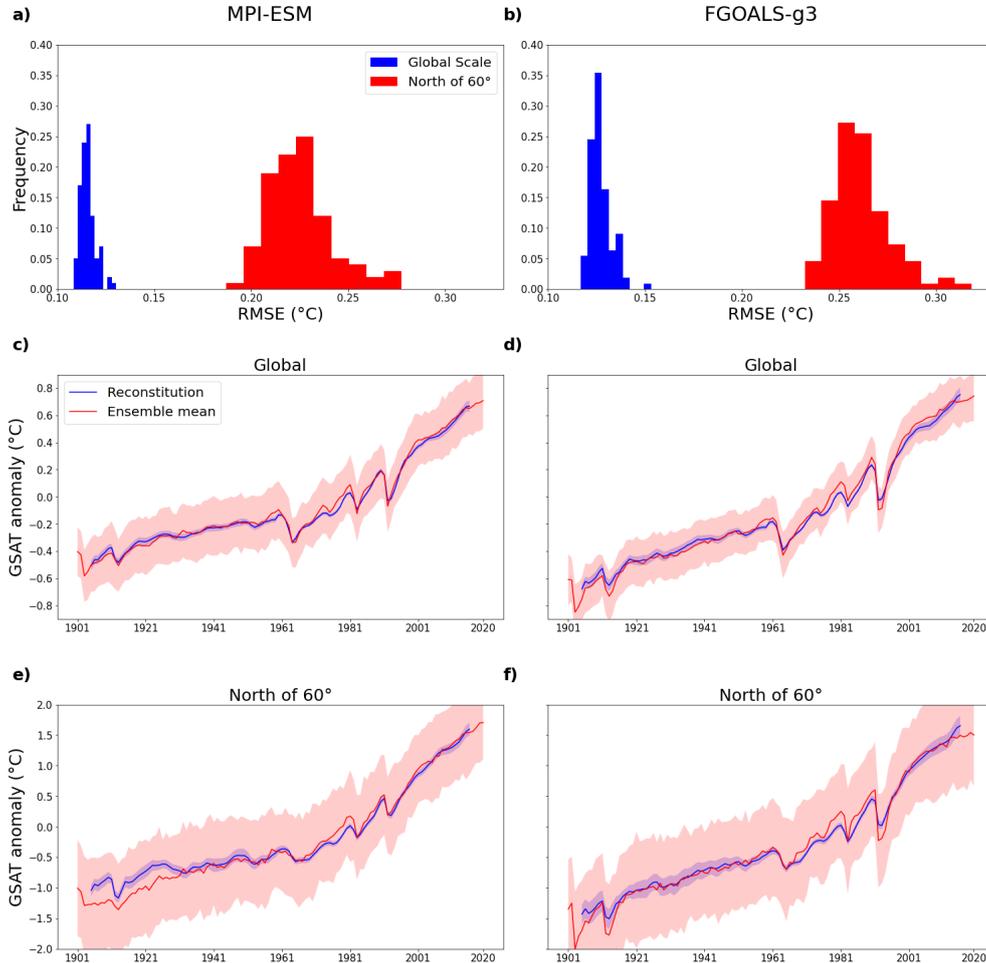


Figure 6. a) Histogram showing the distribution of the RMSE between the mean ensemble and the U-Net outputs of MPI-ESM. b) Same as a), but for FGOALS-g3. c) Time evolutions of the global mean surface air temperature, in $^{\circ}\text{C}$, for the ensemble mean and the mean U-Net outputs for MPI-ESM. Color shade shows the spread of the time series, with 90% the ensemble members uncertainty assuming a gaussian distribution. d) Same as c) but for FGOALS-g3. e) and f) are the same as c) and d) but when averaging the SAT, in $^{\circ}\text{C}$, north of 60°N .

500 individual member results using a basic ensemble mean approach. This evaluation is con-
 501 ducted through a random subsampling process involving 500 sets of m members, where
 502 m varies from 1 to 40, for both the FGOALS-g3 and MPI-ESM ensembles. Within each
 503 subset, ensemble means are calculated. The RMSE between these subsample ensemble
 504 means and the actual ensemble mean obtained from all members is then determined (de-
 505 picted by vertical red and blue lines in Figure 7). This RMSE computation is performed
 506 across all grid points and is spatially averaged. The 90% intervals, assuming an Gaus-
 507 sian distribution, of the 500 subsamples are also illustrated. This analysis is done for both
 508 the MPI-ESM and FGOALS-g3 ensembles across distinct geographical regions: global
 509 (90°S-90°N), North Atlantic (60°W-0°E, 0°N-60°N), North Pacific (120°E-100°W, 20°N-
 510 60°N), Niño3 (5°N-5°S, 150°W-90°W), as well as polar regions north of 60°N and south
 511 of 60°S. These chosen regions exhibit considerable forced and internal variability, as vi-
 512 sually demonstrated in Fig. 1 and Fig. 2. Additionally, this evaluation is extended to
 513 encompass both oceanic and terrestrial areas in the 60°S-60°N band, allowing for a more
 514 comprehensive understanding of the U-Net’s performance. The horizontal lines in the
 515 illustration correspond to the same RMSE values but for the U-Net output from each
 516 individual member. The accompanying color shade represents the spread of 90% uncer-
 517 tainty assuming an Gaussian distribution.

518 Figure 7a visually illustrates the progression of errors within the subset of mem-
 519 bers as the size of the subset increases. This pattern aligns with expectations, as a larger
 520 subset size leads to better estimations of forced variability and a corresponding reduc-
 521 tion in residual internal variability by a factor of \sqrt{n} . The distribution of U-Net outputs
 522 mirrors the histograms presented in Figure 6, showing a high degree of similarity across
 523 both climate models. The U-Net effectively diminishes internal variability in GSAT by
 524 approximately a factor of slightly more than four, which is analogous to the residual vari-
 525 ability observed within subsets containing around 17 members for FGOALS-g3 and 20
 526 members for MPI-ESM. When focusing on regions spanning oceans and land between
 527 60°N and 60°S, the outcomes remain largely consistent, showcasing a reduction in error
 528 magnitude by a factor of approximately four. This reduction corresponds closely to that
 529 achieved by using a subset of 15 to 20 members.

530 The U-Net’s efficacy stands out prominently over the equatorial Pacific region, as
 531 depicted in panel 7f. This region is known for being heavily influenced by the ENSO, which
 532 dominates internal variability. The U-Net achieves a substantial reduction in variabil-

533 ity, amounting to a factor of 5.5. This reduction is akin to the outcome of utilizing an
534 ensemble mean derived from around 30 members for both MPI-ESM and FGOALS-G3.

535 In other regions, the variability reduction is quite similar to that found globally.
536 For instance, this consistency is observed in the North Pacific and polar regions, where
537 the required number of members for equivalent outcomes remains relatively steady. How-
538 ever, in terms of removing internal variability, the U-Net showcases higher efficiency in
539 the context of MPI-ESM for most scenarios. This pattern holds true except for the North
540 Atlantic, where a notable deviation is observed: a set of 15 members is necessary in MPI-
541 ESM to achieve results equivalent to the U-Net (~ 4 -fold reduction in residual variabil-
542 ity), while merely 5 members suffice for FGOALS-g3 (halving of the residual variabil-
543 ity).

544 The variation in performance between FGOALS-g3 and MPI-ESM might arise from
545 dissimilarities in their internal variability, particularly over multi-decadal timescales, or
546 due to differences in forced variability compared to the training data. Having completed
547 this method evaluation, our focus now shifts to examining the outcomes when the U-Net
548 is employed with observational data.

549 **4.1 Filtering of the observations**

550 The U-Net is now employed to process SAT observations derived from GISSTEMP.
551 By utilizing observed data as input, the U-Net provides an estimate of the forced vari-
552 ability. In the interval from 1996 to 2016, the U-Net-derived forced SAT (depicted in Fig-
553 ure 8a) illustrates a fairly uniform warming, with amplified warming evident over the
554 Arctic region, consistent with Arctic amplification. Furthermore, this warming effect is
555 slightly more pronounced over land compared to oceans. Conversely, the Southern Ocean
556 experiences less warming in comparison to other global regions. The spatial distribution
557 of standard deviations (Figure 8b), computed from 1905 to 2016 using U-Net output,
558 mirrors the anomalies observed in the 1996-2016 period. This agreement indicates the
559 prevailing influence of increasing anthropogenic forcing. Notably, this pattern closely re-
560 sembles the changes observed in the multi-model mean (MMM) (as depicted in Fig. 2).
561 This underscore the significant contribution of the training dataset in determining the
562 identified forced changes.

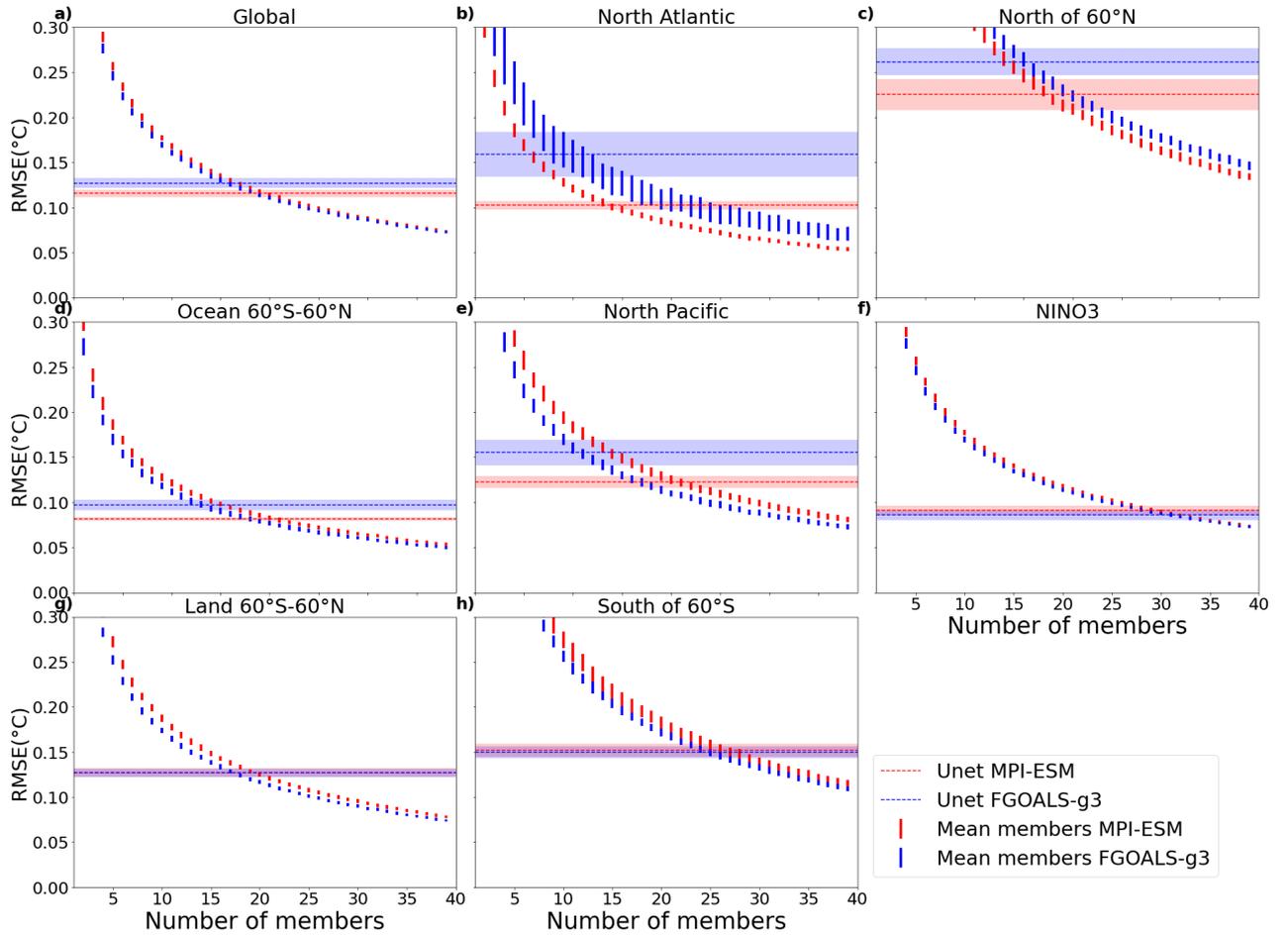


Figure 7. Spatial average of the RMSE for the forced variability estimated with the U-Net outputs obtained from each ensemble member, and the forced variability obtained with ensemble averages subsampling ensemble of size 1 to 40; for (red) MPI-ESM and (blue) FGOALS-g3. The RMSE calculated from the U-Net and each ensemble member is given by (color shade) the interval including 90% of the distribution, assuming a gaussian distribution, and (horizontal dashed line) the mean RMSE. The RMSE calculated from 500 subsample of size between 1 to 40 is illustrated with (vertical lines) the intervals including 90% of the ensemble member distribution, also assuming a gaussian distribution.

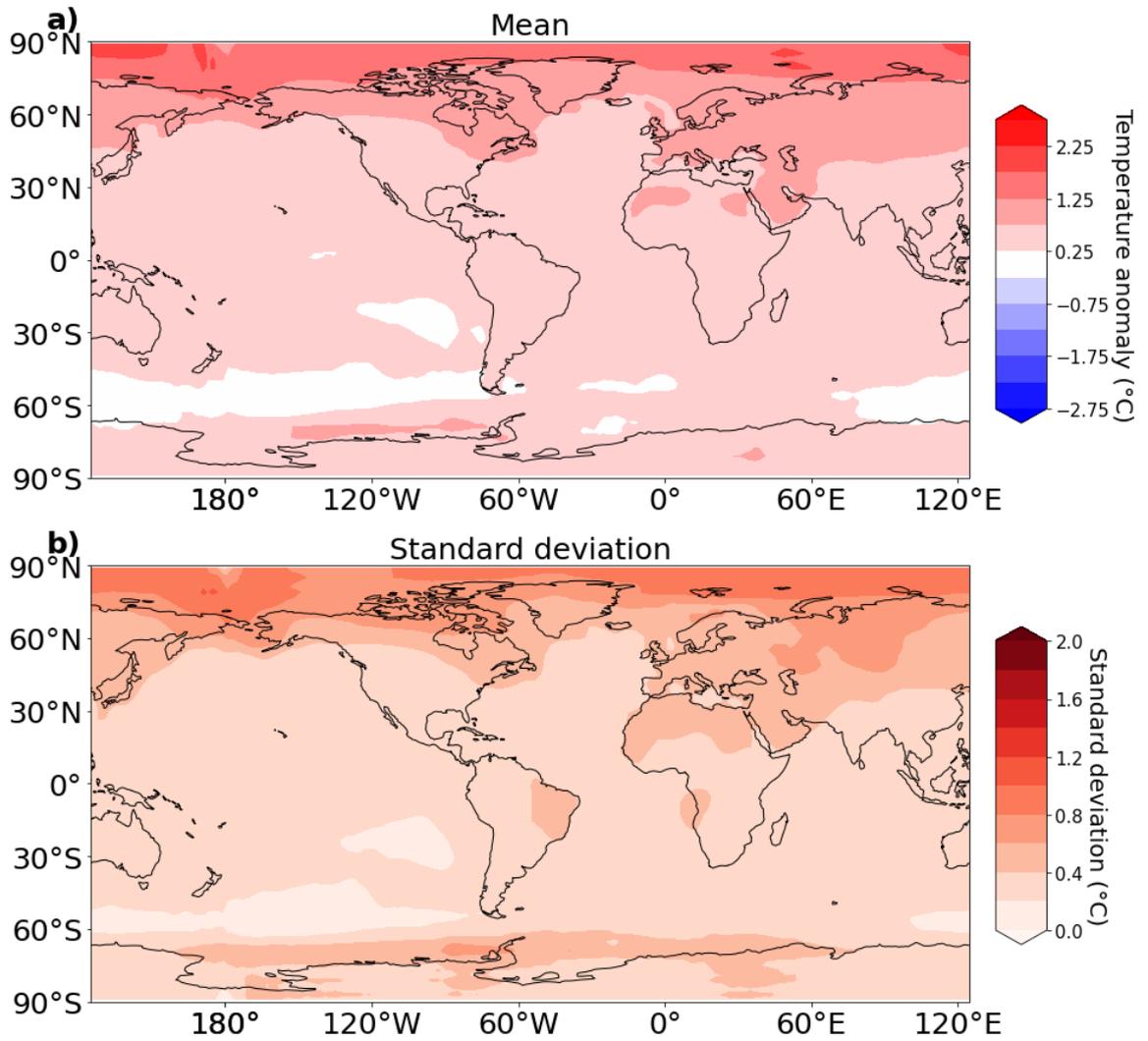


Figure 8. Forced surface air temperature (in °C) anomaly when applying the U-Net to GIS-STEMP observation : a) time average in 1996-2016; b) standard deviation in 1905-2016.

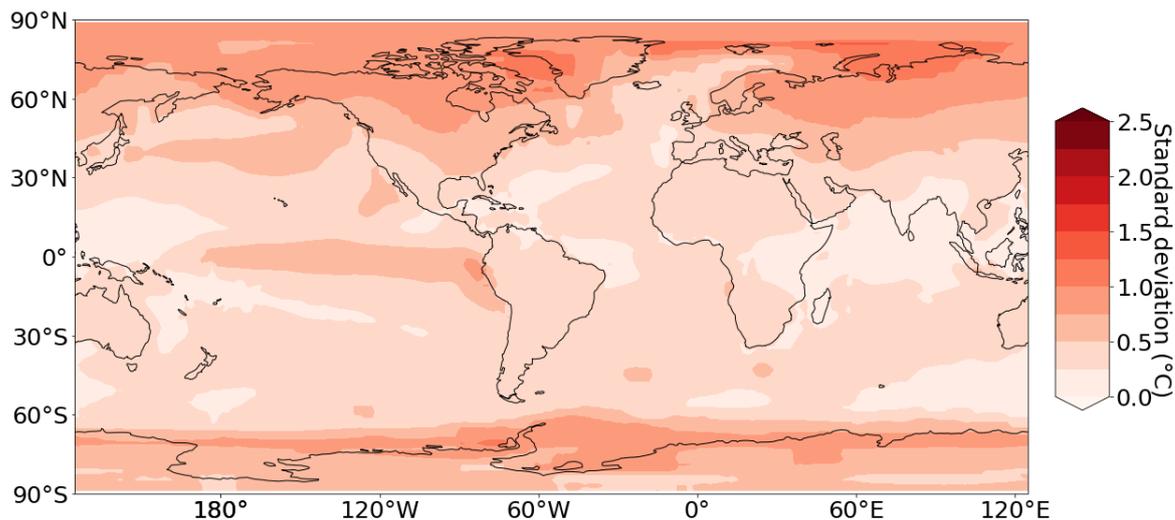


Figure 9. Standard deviation of the SAT deviations from the forced SAT, as estimated using the U-Net, in 1905-2016.

563 To quantify internal variability within the observations, we compute the deviations
 564 of observed SAT anomalies from the estimated forced changes. The resulting internal
 565 variability pattern, illustrated by the time standard deviation of these deviations shown
 566 in Figure 9, mirrors the model-derived pattern (Fig. 1). Higher internal variability val-
 567 ues are observed over land areas, as well as regions near the boundaries of sea ice, such
 568 as the Labrador Sea and the Nordic Seas in the Northern Hemisphere, and the South-
 569 ern Ocean. Notably, a local maximum of internal variability emerges in the equatorial
 570 Pacific, corresponding to the El Niño-Southern Oscillation region. This similarity in the
 571 spatial distribution of internal variability between observations and models underscores
 572 the consistency of our findings.

573 We now shift our focus to the GSAT and the Niño 3.4 region (5°N-5°S, 170°W-120°W),
 574 with a particular emphasis on Niño 3.4 due to its notably improved performance in our
 575 study. In the global context (Figure 10a), the forced variability reveals a consistent warm-
 576 ing trend, which becomes more pronounced during the 1960s. Notably, the major vol-
 577 canic eruptions of Agung (1963), El Chichón (1982), and Pinatubo (1991) are associated
 578 with temporary cooling patterns. By 2016, the GSAT anomaly reaches 0.7°C. As expected,
 579 the forced variability time series exhibits a significant reduction in inter-annual variabil-
 580 ity. This reduction is particularly striking within the Niño 3.4 region (Figure 10b), where
 581 variability at 2 to 7 years is almost entirely eliminated. The U-Net estimates the Niño

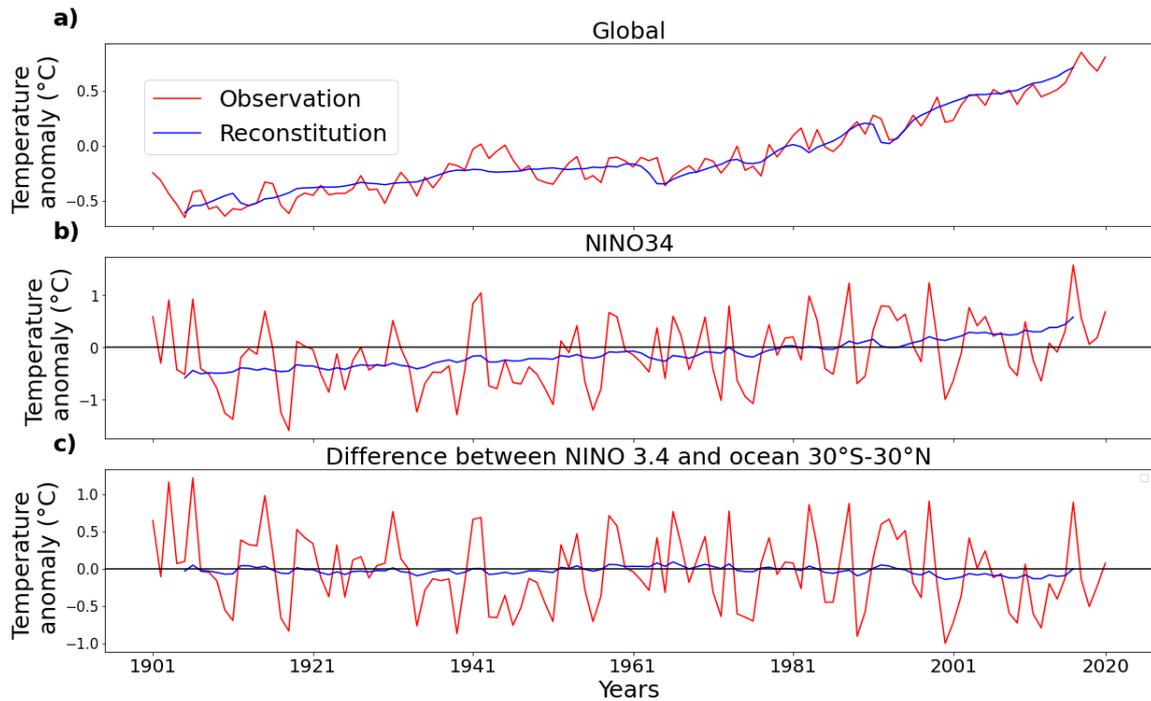


Figure 10. Time series of (red) the observed SAT anomaly and (blue) the forced SAT anomaly estimated by the U-Net for a) the global mean b) NINO 3.4 and c) the relative SAT, calculated as the difference between the averaged SAT in Niño 3.4 region and the tropical ocean SAT (30°S-30°N).

582 3.4 forced variability, depicting a steady warming trend. To quantify the changes of SAT
 583 in Niño 3.4 relative to the tropics, we calculate also the relative SAT, defined as the dif-
 584 ference between the average SAT on the NINO 3.4 region and the average SAT on ocean
 585 grid between 30°S-30°N. The relative SST shows that the warming over the Niño 3.4 fol-
 586 lows that of the tropics, so that no clear El Niño-like reponse is found, unlike climate
 587 models (Fig. 2). Some authors (Clement et al., 1996; Heede et al., 2020) have suggested
 588 that a forced cooling could exists in the relative SAT, called thermostat effect. Here the
 589 relative SAT shows a very small cooling (see Fig. 10c). In addition the SAT in the Niño
 590 3.4 region are not affected by the forcing from the main volcanic eruptions. Therefore,
 591 no evidence of a Niño-like response to volcanic eruption (as in Khodri et al. (2017)) is
 592 found.

5 Conclusion

A novel approach is introduced in this study to effectively eliminate internal variability from a time-evolving two-dimensional dataset, specifically focusing on surface air temperature. The method employs a U-Net neural network and draws inspiration from the noise-to-noise technique. This framework treats internal variability as an analogous noise superimposed on the underlying forced variability. The U-Net model is trained using outputs from a diverse ensemble of climate models obtained from the CMIP simulations. Subsequently, this trained network is applied to observational data to unveil the forced variability signal by attenuating internal variability. The validation of this method involves utilizing large ensemble simulations from individual models, specifically the MPI-ESM and FGOALS-g3, to gauge its effectiveness. The forced variability derived from the ensemble mean is then contrasted with the outcomes from the U-Net application. To quantitatively assess the U-Net's efficacy in reducing internal variability, an "equivalent ensemble size" is computed. This metric indicates the ensemble size that would be required to achieve the same level of precision in capturing forced changes as the U-Net which is applied to a single member. The U-Net outputs for these two climate models' test data exhibit an error equivalent to an internal variability reduction of a factor of more than 4. This magnitude corresponds to the internal variability one could expect from an ensemble averaging 17 to 20 members. Furthermore, when the U-Net is applied to surface air temperature observations, the inferred forced changes align closely with the multi-model mean in terms of spatial patterns. The U-Net's results do not suggest an El Niño-like response to global warming. We observe that the U-Net encounters greater challenges in accurately estimating forced variability over the Arctic region. This discrepancy can be attributed to the significant forced and internal variability associated with changes in sea-ice extent in that area. Additionally, the U-Net's performance in capturing forced variability in the North Atlantic is less successful for the FGOALS-g3 model. This limitation might be linked to uncertainties stemming from the multi-decadal variability prevalent in these regions (Menary & Wood, 2018; Zhang, 2007).

In the pursuit of enhancing the U-Net methodology, several avenues for future improvements have been identified. One potential approach is to address the U-Net's sensitivity to the multi-model consensus of future variability by employing neural network regularization techniques, such as weights penalisation. Additionally, preprocessing methods like data augmentation could be explored to potentially mitigate such impacts. Im-

626 proving the evaluation process of the U-Net’s performance is also on the horizon. This
627 could involve testing the U-Net on a broader range of climate models to assess its gen-
628 eralizability. Comparing its outcomes with results from alternative methods, such as signal-
629 to-noise filtering, could offer a comprehensive evaluation of the U-Net’s effectiveness. To
630 broaden the scope of application, the U-Net’s performance might be further investigated
631 using additional climate variables beyond surface air temperature (SAT). Variables such
632 as sea level surface pressure and precipitation could be explored, capitalizing on poten-
633 tial correlations among these variables to provide more comprehensive insights. Lastly,
634 the proposed method holds the potential for wider applications, including its deployment
635 on simulations from projects like the Detection and Attribution Model Intercomparison
636 Project (Gillett et al., 2016) or the Large Ensemble Single Forcing Model Intercompara-
637 tion Project (D. M. Smith et al., 2022). By leveraging transfer learning, the U-Net trained
638 on historical simulations could be adapted to these datasets. This adaptation could fa-
639 cilitate the evaluation of specific forcing effects in individual climate models, offering a
640 valuable tool for studying the impact of different external factors on the climate system.
641 Such extensions of the method could contribute significantly to our understanding of cli-
642 mate attribution and variability.

643 **Acknowledgments**

644 We acknowledge the support of the SCAI doctoral program managed by the ANR with
645 the reference ANR-20-THIA-0003, the support of the EUR IPSL Climate Graduate School
646 project managed by the ANR under the "Investissements d’avenir" programme with the
647 reference ANR-11-IDEX-0004-17-EURE-0006. This work was performed using HPC re-
648 sources from GENCI-TGCC A0090107403 and A0110107403, and GENCI-IDRIS AD011013295.
649 Guillaume Gastineau was funded by the JPI climate/JPI ocean ROADMAP project (grant
650 number ANR-19-JPOC-003).

651 **6 Open Research**

652 **Data Availability Statement**

653 The CMIP5 and CMIP6 data is available through the Earth System Grid Feder-
654 ation and can be accessed through different international nodes. For example : [https://](https://esgf-node.ipsl.upmc.fr/projects/esgf-ipsl/)
655 esgf-node.ipsl.upmc.fr/projects/esgf-ipsl/

656 Codes used in this article for the backward optimization and the figures are from
 657 Bône (2023) software available freely at <https://zenodo.org/record/8233743>.

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