

Attribution of River-Sourced Floating Plastic in the South Atlantic Ocean Using Bayesian Inference

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Key Points:

- We developed a probabilistic framework to attribute the sources of floating oceanic plastic
- The framework uses Bayes theorem to combine river plastic emissions with Lagrangian simulations
- The framework yields probability maps and age distributions of the most likely source in the region

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15 **Abstract**

16 Most marine plastic pollution originates on land. However, once plastic is at sea, it is
17 difficult to determine its origin. Here we present a Bayesian inference framework to com-
18 pute the probability that a piece of plastic found at sea came from a particular source.
19 This framework combines information about plastic emitted by rivers with a Lagrangian
20 simulation, and yields maps indicating the probability that a particle sampled somewhere
21 in the ocean originates from a particular source. We applied the framework to the South
22 Atlantic Ocean, focusing on floating river-sourced plastic. We computed the probabili-
23 ty as a function of the particle age, at three locations, showing how probabilities vary
24 according to the location and age. We computed the source probability of beached par-
25 ticles, showing that plastic found at a given latitude is most likely to come from the clos-
26 est source. This framework lays the basis for source attribution of marine plastic.

27 **Plain Language Summary**

28 Plastic is commonly found floating near the surface of the ocean but it is difficult
29 to know where it was introduced into the environment. For some large plastic items, the
30 origin can be estimated by analysing the information printed on them, but for small par-
31 ticles, this information is typically missing. To estimate the origin of particles at sea, we
32 built a framework that assigns a probability indicating the chance of finding a particle
33 that came from a particular source, found at a specific location of the ocean. The frame-
34 work uses estimates of plastic emitted by rivers, in combination with a simulation of the
35 transport of particles at the ocean surface, to compute the probability that a particle,
36 found at a particular location in the South Atlantic, comes from a certain river. Sim-
37 ilarly, we computed the probability that a particle of a certain age (defined as the time
38 it has been drifting in the ocean) comes from a particular river, showing that the prob-
39 ability changes according to the particle age. Finally, we computed the probability for
40 particles stranded at the coasts of South America and Africa, showing that plastic found
41 on beaches is most likely to come from the closest river.

42 **1 Introduction**

43 Floating plastic items have been found in all of the world's oceans (Eriksen et al.,
44 2014; Van Sebille et al., 2015), but the origins (i.e. where and when the plastic entered
45 the ocean) of these plastic items are often not obvious. For some of the larger macroplas-

tics, the origin can be attributed by careful analysis of labels (e.g. Lebreton et al. (2018); Schofield et al. (2020); Turner et al. (2021)), but most (micro)plastic particles are too small and nondescript for their origin to be identified this way. Nevertheless, it is important to assess and possibly attribute the likely source for these smaller particles too, as they are among the most harmful to marine ecosystems (Koelmans et al., 2019).

Here, we use numerical simulations to compute the pathways of virtual plastic particles that float on the surface of the ocean (Hardesty et al., 2017; Van Sebille et al., 2018). By tracking particles, it is in principle possible to connect any source with any location. However, the multitude of possible sources very quickly makes this a computationally unwieldy approach. To overcome this computational challenge, we here propose using a Bayesian inference approach to attribute sources in a probabilistic sense.

Such a probabilistic approach has been used before to locate objects lost at sea, like the submarine *Scorpio* (Richardson et al., 1971) and the (yet to be found) Malaysian Airlines flight MH370 (Davey et al., 2016). The main difference between these search & rescue applications of Bayesian inference and our application in the source attribution of floating plastic is that the sources of plastic are spatially very heterogeneous, and so is its distribution at sea.

To develop this probabilistic framework for attribution of likely plastic sources, we here focus on plastic emitted by rivers, as rivers are considered the principal pathway for mismanaged plastic waste (MPW) into the ocean (Lebreton & Andrady, 2019). We selected the South Atlantic Ocean as the study location because the South Atlantic Sub-tropical Gyre is an accumulation zone for plastic (Cózar et al., 2014; Ryan, 2014; Morris, 1980), but also because of the presence of large urban centers along the American and African coast that contribute to the plastic found at sea (Jambeck et al., 2018; do Sul & Costa, 2007), and because we plan to compare our results with samples collected during a 2019 expedition to the region.

2 Theory

Bayesian inference uses Bayes' Theorem to estimate the conditional probability of an event happening under certain conditions by combining prior knowledge about the problem with data obtained through an experiment. In particular, our objective is to estimate the probability that a particle sampled at sea would come from a certain source.

77 This can be written as the conditional probability $p(R_i|S_{loc})$: the probability of sampling
 78 a particle at a location S_{loc} from a specific source R_i .

79 Bayes' theorem offers a way of estimating $p(R_i|S_{loc})$, by combining prior knowl-
 80 edge with new observations. In our case, Bayes' theorem is

$$p(R_i|S_{loc}) = \frac{p(S_{loc}|R_i)p(R_i)}{p(S_{loc})}, \quad (1)$$

81 where $p(R_i|S_{loc})$ is the conditional probability that we aim to estimate, $p(S_{loc}|R_i)$ is the
 82 opposite conditional probability that can be estimated by performing a numerical sim-
 83 ulation (see below), $p(R_i)$ is the probability of a particle being released at a particular
 84 source and $p(S_{loc})$ is the probability of sampling a plastic particle in a specific location,
 85 regardless of the source. It is important to note that $p(R_i|S_{loc}) \neq p(S_{loc}|R_i)$. The lat-
 86 ter term namely indicates the probability of a plastic particle found at a location to come
 87 from a specific source, and the former indicates the probability of a particle coming from
 88 a specific source being at a location. Each term is commonly referred to by it's inter-
 89 pretation. For instance, $p(R_i)$ is denoted as 'the prior' because it represents the prior
 90 knowledge of the problem, $p(S_{loc}|R_i)$ is 'the likelihood', which updates our prior knowl-
 91 edge from the problem, $p(S_{loc})$ is the 'normalizing constant', and $p(R_i|S_{loc})$ is 'the pos-
 92 terior'.

93 In eq. (1), computing the normalizing constant $p(S_{loc})$ requires observations for all
 94 plastics in the ocean regardless of their source, which means that $p(S_{loc})$ also considers
 95 plastic that comes from sources that are not taken into account in the numerator of eq. (1).
 96 Therefore, the posterior probabilities at each S_{loc} would not add to one in each location
 97 but instead will add to a fraction that corresponds only to the sources of plastic consid-
 98 ered in the study. This is inconvenient when the focus is only on plastic coming from spe-
 99 cific sources such as riverine plastic. To overcome this inconvenience, we can constrain
 100 the sum of all posterior probabilities to be equal to one

$$\sum_{i=1}^N p(R_i|S_{loc}) = 1, \quad (2)$$

101 where the sum is defined for the N number of sources. Then, substituting $p(R_i|S_{loc})$ for
 102 eq. (1)

$$\sum_{i=1}^N \frac{p(S_{loc}|R_i)p(R_i)}{p(S_{loc})} = 1, \quad (3)$$

103 and by factorizing and solving for $p(S_{loc})$

$$p(S_{loc}) = \sum_{i=1}^N p(S_{loc}|R_i)p(R_i), \quad (4)$$

104 we obtain a normalizing constant that only considers the sum of all our hypotheses (i.e.
 105 products of prior and likelihoods). Finally, by substituting $p(S_{loc})$ in eq. (1) we get

$$p(R_i|S_{loc}) = \frac{p(S_{loc}|R_i)p(R_i)}{\sum_{i=1}^N p(S_{loc}|R_i)p(R_i)}, \quad (5)$$

106 which is an alternative form of Bayes' theorem (Carlin & Louis, 2008) that ensures that
 107 the sum of all posterior probabilities is one in each location. This last equation is used
 108 in this study.

109 **3 Methodology**

110 **3.1 Selecting the Sources and Computing the Prior**

111 Our prior is based on the annual amount of riverine plastic estimated by Meijer
 112 et al. (2021), who used a probability framework combined with geographical data of MPW
 113 to estimate the plastic mass emissions of the world rivers into the ocean, at the location
 114 of the river mouths. From their global data set, we selected the locations and annual emis-
 115 sions for all 1,010 rivers that emit plastic into the South Atlantic. To avoid immediate
 116 beaching, we moved the river mouth locations to the center of the closest ocean grid-cell
 117 of the model's flow field. When various rivers shared the same closest grid-cell, we summed
 118 their emissions. This condensed the number of release locations to 535 (without affect-
 119 ing the total amount of plastic released by the rivers in the South Atlantic).

120 We then clustered the rivers in 10 groups that contained the top polluting rivers
 121 and their neighboring rivers. These clusters are 2° by 2° square regions centered around
 122 ten locations that coincide with important cities or river estuaries. We used the result-
 123 ing 283 river mouth locations in these 10 clusters as the release positions for the parti-
 124 cles in the simulation. The 10 clusters (Figure 1) account for 87.9% of the riverine plas-
 125 tic emissions in the South Atlantic. There are two clusters on the African coast: around
 126 the city of Cape Town and on the Congo River estuary. The other eight clusters are on
 127 the South American coast: five near the cities of Rio de Janeiro, Porto Alegre, Santos,
 128 Salvador, and Recife; and three on the river estuaries of Rio de la Plata, Itajaí and Paraíba.

129 We defined the prior distribution $p(R_i)$ to be the fraction of plastic emitted at each
 130 cluster, normalised by the total amount of plastic emitted at the 10 clusters. Our prior
 131 thus is a 10-dimensional categorical or discrete distribution, in which each source has an



Figure 1. Map of the top 50 rivers (red dots) in the South Atlantic from Meijer et al. (2021) and the clusters (black squares) used as sources in this study. The size of the red circles is proportional to the rivers’ plastic emission. The size of the black squares exaggerates the true size of the clusters, which is 2° by 2°.

132 associated probability defined between 0 to 1, and the sum of the 10 probabilities is 1.

133 The probability associated with each source is shown in Table 1.

134 3.2 Simulation Setup and Computing the Likelihood

135 To compute the likelihood $p(S_{loc}|R_i)$, we released virtual particles from each of the
 136 sources R_i and tracked them through the South Atlantic surface flow. We performed the
 137 simulation using the Parcels framework (Delandmeter & Sebille, 2019) on the Surface
 138 and Merged Ocean Currents (SMOC) data set from the Copernicus Marine Environmen-
 139 tal Service (CMEMS) (Drillet et al., 2019). The SMOC data set is a 2D surface flow field,
 140 with a 1/12° resolution, of the sum of the velocity contributions from the Eulerian com-
 141 ponent associated with currents, the tidal component, and the Stokes Drift component
 142 associated with waves (Drillet et al., 2019). In this study we assumed that the particles
 143 were at the surface at all times.

144 The domain of the simulation was the South Atlantic Ocean, from 70°W to 25°E
 145 and between 50°S to the Equator. We used hydrodynamic data from 1 April 2016 to 31
 146 August 2020, releasing particles in the first year only and then tracking them for another
 147 3.4 years. During the simulation, if a particle left the domain, we stopped tracking that

Sources (R_i)	Proportion (%)	$p(R_i)$
Congo	1.6	0.019
Cape Town	4.2	0.051
Rio de la Plata	9.8	0.121
Porto Alegre	8.3	0.099
Santos	4.6	0.048
Paraibá	3.8	0.031
Itajaí	7.5	0.086
Rio de Janeiro	28.5	0.334
Salvador	6.8	0.078
Recife	12.7	0.133
Other rivers	12.1	-

Table 1. The proportion of the total annual plastic released to the South Atlantic and the prior probability $p(R_i)$ of a particle being released at a specific source R_i . The "Other rivers" row indicates the proportion of plastic from rivers outside the clusters and is therefore not considered in $p(R_i)$.

148 specific particle. We implemented a stochastic parametrization for beaching of buoyant
 149 particles as described in Onink et al. (2021), using a beaching timescale of $\lambda_b = 10$ days
 150 and a re-suspension timescale of $\lambda_r = 69$ days. To parametrise unresolved turbulence
 151 that acts on the floating plastic (Van Sebille et al., 2020), we implemented uniform dif-
 152 fusion defined in the whole domain with a value of $10 \text{ m}^2 \text{ s}^{-1}$, similar to Onink et al. (2021)
 153 and Lacerda et al. (2019).

154 We performed one simulation with 100,000 particles per source, with a fourth-order
 155 Runge-Kutta integration time step of 1 h. We released the particles from the 283 river
 156 mouth locations inside the 10 clusters. On average, the particles were released 10 km from
 157 the coast. The number of particles released at each location was proportional to the emis-
 158 sion of each of the rivers within the cluster, and equally distributed over one year. We
 159 stored the particles' positions every 24 h, for a total of 1,234 points per trajectory.

160 We computed the likelihood $p(S_{loc}|R_i)$ by binning the particle positions in $1^\circ \times 1^\circ$
 161 bins. For this, we counted the particles inside each bin at every time step, and then we
 162 averaged the number of particles during a time period. Then, we divided the average num-
 163 ber of particles at each bin by the sum of all the averaged counts in all bins. The $p(S_{loc}|R_i)$
 164 in each bin has a value between 0 and 1 and the sum of the probabilities of all bins is
 165 1. This yielded 10 $p(S_{loc}|R_i)$ maps, one per source R_i .

166 The likelihood was computed based on the positions of the particles according to
 167 their age. The particle age represents the transit time of particles between the source
 168 R_i and a sampling location S_{loc} (i.e., their drifting time), with each particle following
 169 a different pathway until reaching S_{loc} (Van Sebille et al., 2018).

170 **3.3 Oceanic Particles Posterior Probability**

171 We computed the posterior probability $p(R_i|S_{loc})$ using eq. (5) independently for
 172 each 1° by 1° bin, using the corresponding likelihood and the normalizing constant in each
 173 particular bin. Doing this for all the sources, we get the local posterior distribution in
 174 each bin, as a probability between 0 and 1 for each source. This results in 10 posterior
 175 probability maps (one per source) which add up to 1 for each bin.

176 **3.4 Beached Particles Posterior Probability**

177 Since we use a stochastic parametrization for simulating the beaching of particles
 178 near the coast, we can also map the probability of a beached particle coming from a spe-
 179 cific source. To compute this, we built two cumulative latitudinal histograms of the par-
 180 ticles that were beached at a specific time step: one for the American coast and the other
 181 for the African coast. The cumulative latitudinal histogram is formed by counting the
 182 particles that are beached in latitudinal bins of 1° , disregarding the longitude of those
 183 particles, and by classifying them into particles that beached either at the American or
 184 the African coast. With the counts per latitude, we computed the average at each bin
 185 for the duration of the whole simulation and normalized by the sum of all average counts
 186 per bin. As for the posterior probability maps, we computed the beached posterior prob-
 187 ability $p(R_i|S_{lat})$ using eq. (5), where S_{lat} is the latitudinal bin.

188 **4 Results**

189 **4.1 Likelihood Maps**

190 Figure 2 shows the likelihood maps for particles released at each source R_i , aver-
 191 aged over a period of 3.4 years. The values for the likelihood in the bins are between 0
 192 to 10^{-4} , as they represent the proportion of particles (in relation to the total number
 193 of particles from a source in the domain) that cross a grid cell. Each source has 100,000
 194 particles, minus the particles that exited the domain at a certain time step, so if in one
 195 bin there are 100 particles, the likelihood would be in the order of 10^{-4} .

196 In general, the dark blue areas represent regions where almost no particles were found
 197 from a specific source, while the yellow regions represent locations where it was more likely
 198 to find particles from that source. Specifically, for the South American sources, the like-
 199 lihood of finding particles from Recife, Santos, and Salvador is almost zero in the open
 200 ocean between 20°S to 40°S , that is, where the subtropical gyre is located. The parti-
 201 cles released from those sources tend to stay close to shore and beach because of the ef-
 202 fect of Stokes drift that pushes them towards the coast.

203 For the sources Itajaí, Paraíba, Porto Alegre, Rio de Janeiro, and Rio de la Plata,
 204 the likelihood to find particles released by any of those sources is the highest in the sub-
 205 tropical gyre (between 20°S to 40°S), with values ranging from 1×10^{-4} to 3×10^{-4} ,
 206 suggesting a high chance of finding particles from those sources, with Porto Alegre be-

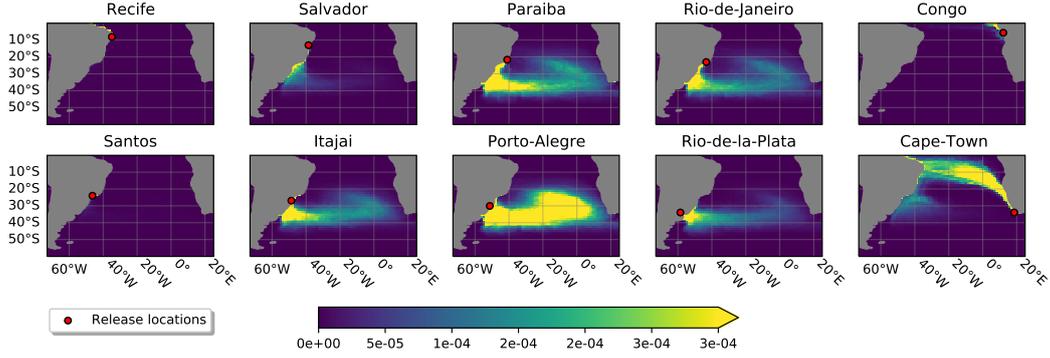


Figure 2. Likelihood maps of the spatially binned $p(S_{loc}|R_i)$ for each source. The color scale indicates the probability of finding a plastic particle coming from the source (indicated as a red point).

207 ing the largest contributor. Closer to the South American coast, the likelihood is above
 208 3×10^{-4} for all these sources. North of the gyre, from 20°S and further north, the like-
 209 lihood of finding particles from the American coast is near-zero.

210 For the African sources, shown on the right of Figure 2, we see that the likelihood
 211 of finding particles released in Cape Town is the highest in the Benguela Current. These
 212 particles are likely to reach the South American coast near the Cape of Saõ Roque, and
 213 will less likely get carried by the Brazil Current towards the coast of Argentina. The par-
 214 ticles released at the Congo get carried away northward to the Equator, outside of the
 215 domain of our simulation, and are unlikely to find these particles in other parts of the
 216 studied domain.

217 4.2 Oceanic Particles Posterior Probability

218 Figure 3 shows the posterior probability $p(R_i|S_{loc})$ maps for each source, averaged
 219 over 3.4 years. In particular, the particles in our simulation did not reach latitudes south
 220 of 50°S, leading to no defined posterior probability in the Antarctic Circumpolar Cur-
 221 rent (ACC). This is due to the generally northward Ekman drift in the ACC (Onink et
 222 al., 2019), and because we assumed that the particles only originate from ten sources placed
 223 north of 50°S. In total, 130,585 particles exited the domain: 97,926 across the Equator
 224 and 32,659 into the Indian Ocean.

225 Regarding the individual panels in Figure 3, the posterior probabilities for Recife
 226 and Santos were near-zero because only very few particles were transported into the open
 227 ocean. The probability that particles that end up between 50°W to 40°W and close to
 228 Brazil originated from Salvador was up to 30%. The posterior probabilities of Itajaí and
 229 Paraíba were below 20% everywhere, with the highest values located in the subtropical
 230 gyre, while probabilities were close to zero in the rest of the domain. For Rio de la Plata,
 231 the highest probabilities were found at the boundary with the ACC, approaching 40%
 232 in probability, with decreasing values when going from there towards the equator. The
 233 two sources that dominated in the region of the subtropical gyre, between 20°S to 50°S,
 234 were Porto Alegre and Rio de Janeiro, with probabilities around 40%. South of South
 235 Africa, particles released from Rio de Janeiro are dominant, contributing 60% to the prob-
 236 ability. The remaining 40% are mainly contributed by Porto Alegre and Rio de la Plata.
 237 In the Benguela Current region and extending northwest to the northern coast of Brazil,
 238 the probability is close to 100% that particles originate from Cape Town. Finally, the
 239 posterior probability of Congo is almost only located near the source and farther north.

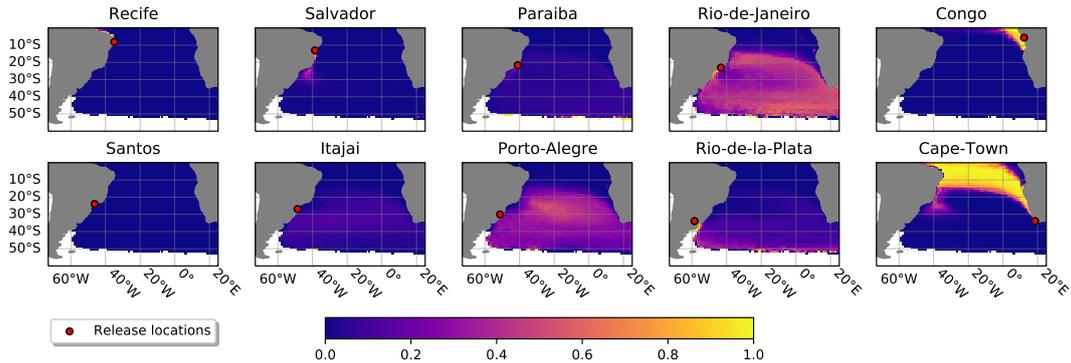


Figure 3. Posterior probability maps, averaged over 3.4 years, showing $p(R_i|S_{loc})$, the probability of finding a particle from a specific source at any point in the South Atlantic. Each map displays the probability for a specific source in all the bins of the domain. The red dots indicate the locations of the sources from which the particles entered the ocean.

240 4.3 Local Posterior Age Distributions

241 The posterior age distributions yield the probable sources of a particle of a certain
 242 age, sampled at a certain location. Figure 4 shows the posterior age distributions for three
 243 sampling locations, averaged over a time window of 30 days. The dashed line represents

244 the number (N) of particles that reach the location as a function of age. The posterior
 245 probability distributions were only computed when $N > 10$.

246 The panel for sampling location A in Figure 4, located in the western part of the
 247 subtropical gyre (32.37°S, 37.64°W), shows for example that a particle sampled at that
 248 location with age younger than 0.4 years is very unlikely to come from any of the con-
 249 sidered river sources. For particles between 0.4 years to 1.0 years, the most probable sources
 250 are Salvador and Porto Alegre. For ages older than 1.0 years, the probability from Sal-
 251 vador drops below 20% while Rio de Janeiro grows. For particles older than 1.5 years,
 252 Porto Alegre and Rio de Janeiro have the largest probabilities, with values fluctuating
 253 between 20% and 40%. The rest of the sources have values below 20%.

254 The posterior age distributions for location B (32.37°S, 5.80°E) in Figure 4, show
 255 that it is unlikely to find particles younger than 1.2 years coming from any of the con-
 256 sidered sources: only particles older than 1.4 years can reach this point. Similar to point
 257 A, the sources with the largest probability, throughout all ages, are Rio de Janeiro and
 258 Porto Alegre. For the younger particles, these probabilities oscillate around 50%, while
 259 for older particles, the two sources decrease down to 30% for 3.4-year-old particles. The
 260 remaining sources stay below 20% for all ages. The plot corresponding to point C located
 261 north of the gyre (19.19°S, 13.39°W), shows that particles reach this location two years
 262 after release. Fewer particles were present on average compared to A and B, reaching
 263 a peak N at 2.7 years. The largest probability corresponds to Rio de Janeiro and Porto
 264 Alegre. Rio de Janeiro is the predominant source of particles of all ages, although, Porto
 265 Alegre becomes significant when particles are 2.5 years old or older. The other sources
 266 remain below 10% for all ages recorded.

267 4.4 Beached Particles Posterior Probability

268 Figure 5 shows the posterior probabilities for a particle to beach at certain latitude,
 269 $p(R_i|S_{lat})$, based on its origin. The $p(R_i|S_{lat})$ for the American coast are displayed in
 270 the right panel and the ones for the African coast are shown in the left panel. On the
 271 American coast, the nearest source to the bin S_{lat} has the highest probability, which peaks
 272 at the same latitude as the source or in its vicinity. This suggests that the plastic found
 273 on beaches close to a source is most likely to come from that source. Santos is the only

274 exception to this trend because its probability is overshadowed by its proximity to Rio
 275 de Janeiro which emissions are six times larger.

276 In the right panel of Figure 5, the beached probabilities for latitudes between 25°S
 277 to 35°S on the African coastline show a dominance of particles coming from the Amer-
 278 ican coast, accounting for 90% of the beached particles. The probability of the beached
 279 particles coming from Cape Town was found to be less than 10% in this region. Between
 280 18°S to 25°S, Cape town was the only source for beached particles. There were also re-
 281 gions, namely between 12°S to 14°S and 16°S to 18°S, where no particles of any of the
 282 considered sources beached. Finally, at latitudes between 5°S to 12°S, the only proba-
 283 ble source was Congo, no particles from other sources beached that far north. The rea-
 284 son that we found 100% probability for one single source or no beached particles at all,
 285 is that we only considered two sources in Africa, which were at the borders of the stud-
 286 ied domain. In the future, more sources need to be considered, both in this region and
 287 outside, to improve these estimates.

288 5 Conclusions and Discussion

289 We introduced a Bayesian probabilistic framework that allowed us to estimate $p(R_i|S_{loc})$,
 290 the probability that a plastic particle, sampled at the surface of the South Atlantic Ocean,
 291 came from a particular source. The framework supports different types of analyses and
 292 can be used, for example, to compute spatial probabilities, compute local probability as
 293 a function of particle age, or analyse the probabilities once a physical process (such as
 294 beaching) alters the particles' state.

295 The time average window used for computing the likelihood $p(S_{loc}|R_i)$ can be ad-
 296 justed according to the aim of the study. Usually, computing the likelihood for small time
 297 windows leads to greater variability in the likelihood and for instance in the posterior
 298 probability. For these reasons, we computed the average likelihood on the whole simu-
 299 lation and from there we computed the posterior probability.

300 As we showed in Figure 3, visualizing the posterior $p(R_i|S_{loc})$ in maps allows us
 301 to identify the most important sources that pollute ocean regions that provide high ecosys-
 302 tem services and that are vulnerable to plastic, such as subtropical gyres (Helm, 2021)
 303 and marine protected areas (Krüger et al., 2017). This can be used to prioritize the re-
 304 duction of MPW in the principal sources to mitigate the problem. In particular, Porto

305 Alegre and Rio de Janeiro are the most probable sources of riverine plastic in the South
306 Atlantic Subtropical Gyre.

307 The local posterior age distributions, shown in Figure 4 further illustrate the anal-
308 ysis that can be done by selecting a location and by computing the probability distri-
309 butions as a function of the particle’s age. This can point us to the most likely source
310 if we estimate the time the plastic has been drifting in the ocean, by assessing its degra-
311 dation (Chamas et al., 2020; Gewert et al., 2015).

312 The latitudinal beached posterior probabilities, shown in Figure 5, demonstrate how
313 this framework can be used to analyse the contribution of different sources to particle
314 sinks (such as beaches) when considering certain physical processes that alter particle
315 pathways (such as the process of beaching). This can be expanded to including other ad-
316 ditional physical processes that can alter the dynamical state of the virtual particles, such
317 as sinking (Lobelle et al., 2021).

318 This study focuses on floating plastic coming from rivers that discharge plastic into
319 the South Atlantic. In our analysis, we ignore plastic entering the domain from the In-
320 dian Ocean leakage (Van der Mheen et al., 2019) and from the North Atlantic (Speich
321 et al., 2007). To consider it, we need to assume these leakages as sources, by knowing
322 how much plastic enters the domain through the boundaries, or expand the domain to
323 consider other basins.

324 One major advantage of the Bayesian nature of our framework is that it allows up-
325 dating the results when better estimates of plastic emissions are available without hav-
326 ing to redo the (computationally expensive) Lagrangian simulations. For instance, it can
327 be expanded by including a prior that accounts for seasonal variations in river-borne plas-
328 tic inputs, or by taking into account different types of land-based or sea-based sources.

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 335 .24416/UU01-90F027. The supporting figures, tables, and text can be found in the sup-
 336 porting information file associated with this article.

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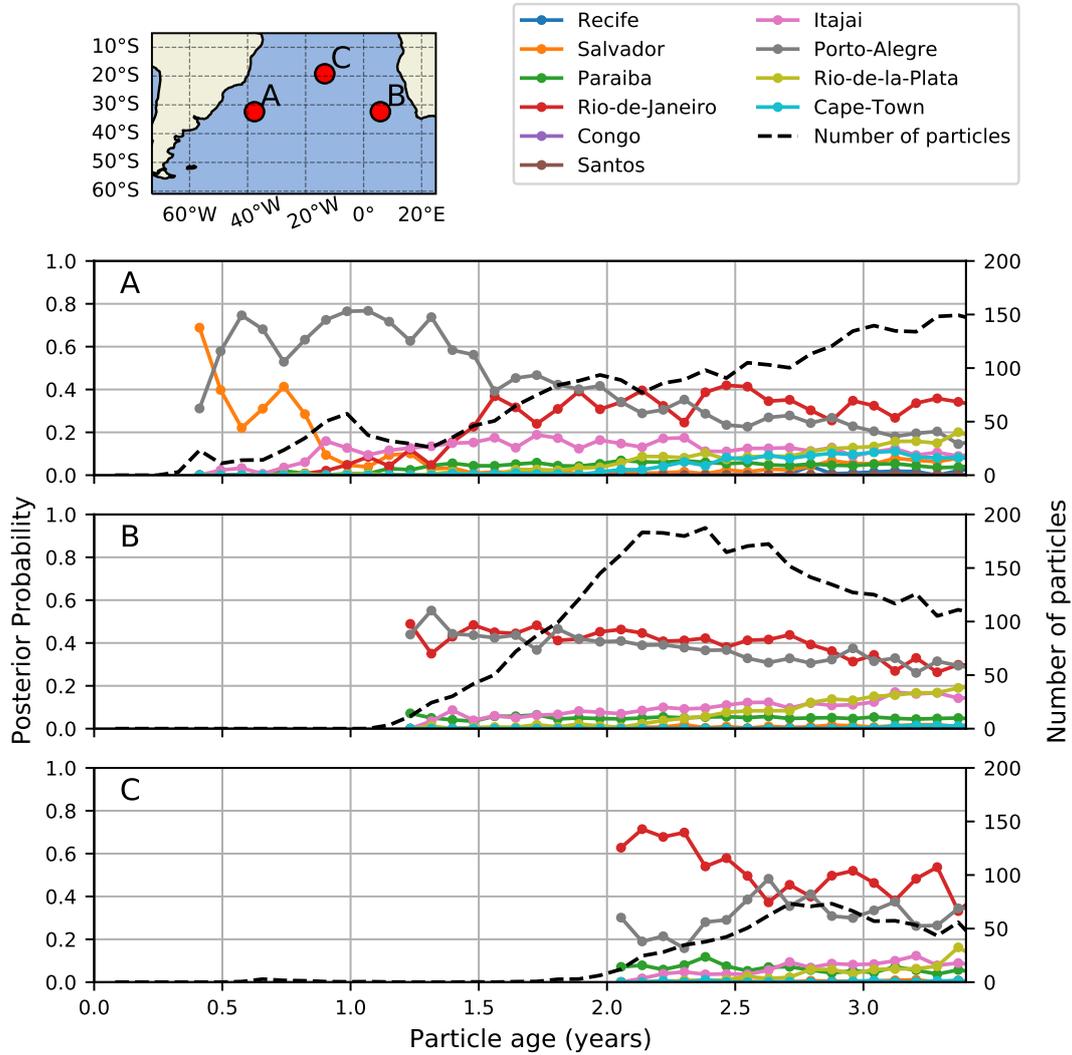


Figure 4. Local posterior age distributions at three different locations for the posterior probability. The map on the top right marks the locations A, B, and C, that correspond to the time series shown in the plots A, B, and C. Each color in A, B and C, represents the probability $p(R_i|S_{loc})$ for a particular source. The black dashed line represents the number of particles (N) at the respective location.

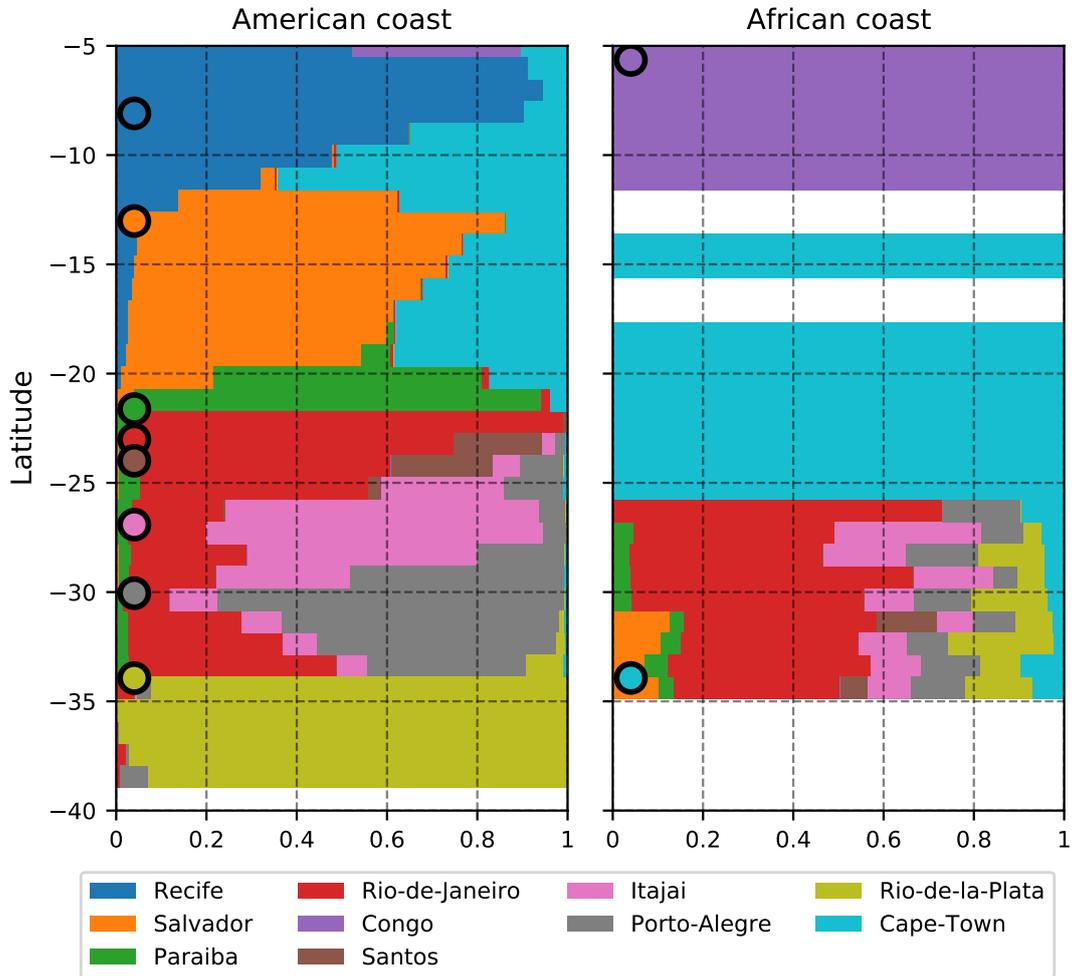


Figure 5. Horizontal bar plot for the posterior probabilities of beached particles (x -axis) at a specific latitude (y -axis). The panel on the left shows the probabilities at the American Coast and the panel on the right the probabilities at the African coasts. Each color is associated with a source, shown in the legend at the bottom. Each latitude has a corresponding horizontal bar summing the probabilities from the sources at that latitude to 1. The round markers on the left of each plot represent the latitudes of the sources. If the marker is on the left panel, the source is at the American coast, and if the marker is in the right panel, the source is located at the African coast.