# Seismic discrimination of controlled explosions and earthquakes near Mount St. Helens using P/S ratios

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**Key Points:**

1. Influences on local P/S ratios are analyzed with 23 shallowly buried explosions and 406 $M_L \geq 1$ earthquakes near Mount St. Helens.

2. Optimized array-median P/S ratios discriminate explosions with 100% true positives and 4.93% false positives.

3. P/S ratio variations at individual stations are interpreted as site effects that correlate with shallow crustal velocity structure.

## Abstract

Explosions and earthquakes are effectively discriminated by P/S amplitude ratios for moderate magnitude events (M≥4) observed at regional to teleseismic distances (≥200 km). It is less clear if P/S ratios are effective explosion discriminants for lower magnitudes observed at shorter distances. We report new tests of P/S discrimination using a dense seismic array in a continental volcanic arc setting near Mount St. Helens, with 23 single-fired borehole explosions ($M_L$ 0.9-2.3) and 406 earthquakes ($M_L$ 1-3.3). The array provides up to 95 three-component broadband seismographs and most source-receiver distances are <120 km. Additional insight is provided by ~3,000 vertical component geophone recordings of each explosion. Potential controls on local distance P/S ratios are investigated, including: frequency range, distance, magnitude, source depth, number of seismographs, and site effects. A frequency band of 10-18 Hz performs better than lower or narrower bands because explosion-induced S-wave amplitudes diminish relative to P for higher frequencies. Source depth and magnitude exhibited negligible influences. Site responses for earthquakes and explosions are correlated with each other and with shallow crustal Vp and Vs from travel-time tomography. Overall, the results indicate high potential for local distance P/S discrimination in a continental volcanic arc setting, with 100% true positives and 4.93% false positives using a frequency band of 10-18 Hz for array-median P/S ratios.

## Plain Language Summary

Methods to remotely classify seismic sources as either shear-slip earthquakes or shallow subsurface explosions are important geophysical tools. They are used for investigations of anthropogenic processes such as underground nuclear explosion tests and mining activity as well as to obtain pure catalogs of tectonic and volcanic earthquakes. The ratio of compressional (P) wave amplitude to shear (S) wave amplitude is effective for classifying large earthquakes and explosions observed hundreds to thousands of kilometers away. Application of such methods at local distances is a topic of growing interest because smaller magnitude sources are only observed at close distances. We tested influences on P/S amplitude ratios and their effectiveness for source discrimination using a data set of tectonic and volcanic earthquakes and controlled shallow borehole explosions near Mount St. Helens. Most observations were within 120 km. We found that use of high frequencies and array-median statistics is largely effective for source

47    discrimination in this setting, successfully classifying 100% of explosions while falsely

48    attributing about 5% of earthquakes to explosions. Variability of P/S ratios across the seismic

49    array has a highly similar pattern for explosions and earthquakes and that pattern appears to be

50    controlled by the seismic velocity structure of the shallow crust.

# 1. Introduction

Quantifying the distinctive properties of explosive seismic sources is an important capability for identification and characterization of underground nuclear explosion tests and a key component of verifying compliance with nuclear test-ban treaties (Bowers and Selby, 2009). Seismic identification and characterization of explosions are also important in other contexts, such as investigations of chemical explosions that may be the subjects of judicial investigations (e.g., Holzer et al., 1996; Koper et al., 2002, 2003) and volcanic explosions studied in the interest of natural hazard mitigation (e.g., Neuberg et al., 1994; Rowe et al., 1998; Johnson and Lees, 2000). Early research on seismic discrimination of nuclear explosions and earthquakes focused on events that are observable at regional to teleseismic distances (≥200 km), which generally requires M≥4 (Richards and Zavales, 1990). One useful seismic approach to discriminating nuclear explosions and earthquakes at such distances is the measurement of P/S amplitude ratios. Abundant empirical evidence confirms that large explosions (M≥4) produce greater P/S amplitude ratios compared to tectonic earthquakes, although the observed ratios are regionally heterogeneous (Taylor et al., 1989; Kim et al., 1993; Walter et al., 1995; Hartse et al., 1997; Xie and Patton, 1999; Jenkins and Sereno, 2001).

Potential expansion of seismic discrimination techniques to local distances and lower magnitudes (or explosive yields) has become a subject of increasing interest (Zeiler and Velasco, 2009; O'Rourke et al., 2016; Kolaj, 2018; Pyle and Walter, 2019; Holt et al., 2019; Kinter et al., 2020). At local distances and lower magnitudes, methods are usually developed to discriminate chemical explosions from earthquakes and, in some cases, to differentiate single-fired explosions from delayed-fired explosions that are common in mining activity (Zeiler and Velasco, 2009; O'Rourke et al., 2016). Single-fired chemical explosions in boreholes provide a useful proxy for underground nuclear tests (Stump et al., 1999) and well-documented examples are available for study on account of geophysical experiments targeting seismic source physics (e.g., Pyle and Walter, 2019) or imaging of geological structures (e.g., O'Rourke et al., 2016). There is mixed evidence regarding the effectiveness of P/S discrimination for ≤M3 events at distances of ≤200 km. Some empirical investigations report that failure to discriminate between earthquakes and explosions may be restricted to the shortest distances (e.g., ≤50 km in O'Rourke et al., 2016). Others report sporadic failures across the local distance range, potentially on account of lithologic variability of the crust causing complex source and path effects that are more subdued

82    at regional to teleseismic distances (e.g., Pyle and Walter, 2019). Further research is needed to

83    address questions such as: Does the complexity of crustal wave propagation make P/S ratios less

84    diagnostic of explosive sources below a common distance threshold? Can diminished explosion

85    discrimination performance be overcome by larger seismic arrays or enhanced local calibration

86    using earthquakes to constrain site and/or path effects? What are the underlying differences in

87    explosion sources that control discrimination capability at local distances?

88         We use a data set with dense local distance recordings of earthquakes and single-fired

89    borehole explosions near Mount St. Helens to gain further insight into the seismic source and

90    structural properties that control the effectiveness of the P/S ratio explosion discriminant for low

91    yields (Fig. 1). Important aspects of the data set are its diverse source-receiver paths with most

92    distances <120 km and its combination of single-fired borehole explosions and earthquakes,

93    including shallow volcanic seismicity at Mount St. Helens (e.g., Ulberg et al., 2020; Glasgow et

94    al., 2018). The source-receiver distribution was optimized for 3D structural imaging, so the

95    source locations are not limited to one isolated test site or linear transect. The inclusion of

96    volcanic earthquakes near Mount St. Helens offers an interesting comparison because volcanoes

97    often produce more frequent shallow earthquakes and more variable source mechanisms

98    compared to other geologic settings.

99    ## 2. Data

100    This study uses seismic recordings of earthquakes and single-fired chemical explosions in

101    shallow boreholes (Table S1). The explosions were detonated for the controlled source

102    component of the iMUSH project to image the magmatic system underlying Mount St. Helens

103    (Kiser et al., 2016, 2018; Hansen et al., 2016). Eight of the 23 boreholes were loaded with 907 kg

104    of explosive and the other 15 were loaded with 454 kg (Fig. 1b). The explosives were buried in

105    shallow boreholes with maximum depths of ~22-28 m and explosives occupying the bottom ~5

106    m of the boreholes. Earthquake source information from 2014 - 2016 was provided by the Pacific

107    Northwest Seismic Network, which contributes to the USGS Comprehensive Catalog (ComCat;

108    USGS, 2020). Earthquakes within 75 km of Mount St. Helens with $M_L \geq 1$ were considered for

109    the study. Lower magnitude earthquakes were not clearly recorded across most of the study area

110    so most seismograms would be removed by the requirement of signal-to-noise (SNR) >2.

111    Additionally, using $M_L \geq 1$ provides a close match to the magnitude range of the explosions ($M_L$

112    0.9-2.3).

113        The primary seismic waveform data for this study are from the broadband part of the

114    iMUSH project (Han et al., 2018; Crosbie et al., 2019; Mann et al., 2019; Ulberg et al., 2020),

115    which temporarily increased the density of local seismic monitoring. The iMUSH broadband

116    network (XD 2014-2016) was combined with permanent regional three-component broadband

117    stations to create a composite network of up to 95 stations within 75 km of Mount St. Helens

118    (Fig. 1a). Sampling rates vary among the networks. Based on the minimum sample rate of 40 Hz

119    we limited our analysis to frequencies ≤18 Hz. Network information for all the broadband data is

120    provided in Table S1. Examples of explosion and earthquake record sections from the broadband

121    array are shown in Figure 2. The example events have equal local magnitudes ($M_L$=1.9) and they

122    are plotted after application of bandpass filtering from 2-6 Hz. The explosion source exhibits a

123    stronger P-wave than the earthquake, while the earthquake exhibits a more prominent S-wave

124    than the explosion (Fig. 2).

125        A secondary source of seismic waveform data is vertical-component geophones deployed

126    during the iMUSH controlled source imaging experiment (Kiser et al., 2016, 2018; Hansen et al.,

127    2016). Our analysis of source discrimination using P/S energy focuses on three-component

128    recordings so we only present record sections from the vertical geophones to illustrate local

129    phase propagation (Fig. 3). In Figure 3, the geophone seismograms for all 23 explosions (70,944

130    source-receiver pairs) were sorted into 1 km distance bins and the median envelope in each

131    distance bin is plotted with trace-normalized amplitude. The high density of recordings illustrates

132    the existence of the S-wave starting at small distances and extending across the array.

133    Comparison of the 2-6 Hz envelopes in Figure 3a and 6-18 Hz envelopes in Figure 3b shows

134    diminished S-wave amplitudes relative to P for higher frequencies. Due to the short duration of

135    the geophone deployment and movement of most geophones during the experiment, there are

136    few opportunities for comparison with earthquakes that were recorded across most of the array.

137    Consequently, we focus on the broadband data from 2014 - 2016 for evaluation of P/S ratio as an

138    explosion versus earthquake discriminant.

# 3. Methods

## 3.1 P/S amplitude ratios

P/S ratios were measured with an approach similar to that used by O'Rourke et al. (2016). Below we focus on the subtle differences with respect to use of three-component data and phase windowing for amplitude measurements. O'Rourke et al. (2016) omitted the transverse components for P measurements because a 1D isotropic medium should not produce P-wave energy on the transverse component. In contrast to O'Rourke et al. (2016), we use all three components for measurements of amplitude in P, S, and pre-event noise windows for both earthquakes and explosions. P-wave energy is commonly observed on transverse components in our dataset (Fig. 2), presumably due to out-of-plane scattering.

Phase windowing at local distances presents a challenge due to the decreasing separation of P and S with decreasing distance. To maximize use of the data we scaled the window length by the S-P time, so that the time window length increases with distance until reaching a fixed duration. The time window starts at -5% of the S-P time before the phase arrival time (P, S or Noise) and ends at 50% of the S-P time after the arrival time for shorter source-receiver distances. At larger distances (i.e., beyond 40 km) a fixed value of 3 sec after the arrival time (Fig. 4) is used for the end of the measurement window. Stations that are too close to an event (i.e., produce a S-P time window less than 1 sec) are discarded because they prevent reliable separation of P and S amplitudes and their arrival times are more variable. Pre-event noise time windows were sampled starting at 10 sec before the origin time of each explosion or earthquake source. Our goal is to ensure the maximum phase energy is captured within the windows while minimizing the contamination from noise or interfering phases. Although the optimization of phase-windowing is somewhat arbitrary, consistent P/S ratio results were achieved in additional tests with the maximum window length increased to 4, 5, and 6 sec. We focus on the 3 sec window results for the remainder of the text.

The phase windows shown in Figure 4 were calculated using a local 1D P and S velocity model derived from the controlled source survey (Fig. 1b and Fig. S1; Kiser et al., 2016). Events with P-wave root-mean-square amplitude SNR>2 were retained for P/S ratio measurements (Fig. S2). Stability in the results for different phase window lengths from 3-6 sec indicates that the

168 deviations from the predicted 1D arrival times have little effect. This is not surprising because

169 phase picks for the explosion sources from Ulberg et al., (2020) have small travel time residuals

170 compared to the 3 sec window length (Fig. 4). Residual times for 1309 manually picked P phases

171 from explosions recorded by the broadband array have a standard deviation of 0.3 sec (average

172 0.01 sec), much smaller than our window width. Manual picks are unavailable for many P and S

173 phases at individual stations because only the clearest arrivals are picked. Thus, we prefer the

174 generic phase windowing approach using the 1D velocity model and 3 sec phase windows for

175 consistency and to avoid biasing amplitude measurements by under-representing weak S-waves

176 that are important to source discrimination.

177    After phase windowing the three-component P/S ratio ($R$) at a given station (n) was

178 calculated as:

$$R_n = \frac{\sqrt{P_i^2 - N_i^2}}{\sqrt{S_i^2 - N_i^2}}, \ \text{Eq(1)}$$

180 where the subscript i denotes summation of the three seismogram components. For source

181 discrimination tests, the array median and scaled median absolute deviation (SMAD) of all the

182 individual P/S ratios are used. P/S ratios were measured with waveforms bandpass filtered in

183 several frequency bands. We tested narrow frequency bands of 2-4 Hz, 4-6 Hz, 6-8 Hz, 8-10 Hz,

184 10-12 Hz,12-14 Hz, and 14-16 Hz, following O'Rourke et al., (2016). We also tested two

185 broader frequency bands of 6-12 Hz and 10-18 Hz. In each case a two-pole Butterworth filter

186 was applied to the response-corrected waveforms after demeaning and detrending.

## 3.2 Bootstrap resampling for different array sizes

188 The temporary broadband array (XD), in combination with the permanent networks (UW, CC),

189 provides better seismic sampling than is available for many areas of interest for explosion

190 monitoring (Fig. 1a). Although up to 95 stations are available during the 2-year period, not all

191 stations/components are available the entire time and meet the SNR>2 threshold. Within our

192 preferred frequency range (10-18Hz), the number of stations that are finally used for P/S ratio

193 calculation for explosions ranges from 36 to 78, with an average of 66. The average number of

194 stations for $M_L \geq 2$ and $M_L$ 1-2 earthquakes are 72 and 47, respectively. To fully evaluate the

195 potential of our preferred method and parameter set, we used bootstrap resampling test to

196     simulate results for six subset array sizes: 2, 4, 8, 16, 32, and 64. For each subset, the cumulative

197     array was randomly resampled 1,000 times (without replacement) from the 74 common stations

198     that recorded all earthquakes and explosions without any prerequisite (i.e., SNR). Within each

199     sampled set, the array-median P/S ratio for earthquakes and explosions are calculated after

200     applying the SNR control. In the other words, the bootstrapping approach attempts to simulate

201     scenarios with smaller arrays and an unconditioned distribution of SNR recordings.

## 3.3 Site corrections for P/S ratios

203     We estimated station-based site corrections using P/S ratios of earthquakes. This was chosen to

204     mimic the circumstance in which P/S corrections could be estimated using only earthquake data

205     and then applied to subsequent events of unknown physical origin. The correction is calculated

206     as a normalization of event-averaged P/S ratio at each of the N stations:

207     $$C_{1,2,3...N} = R_{eq@station(1,2,3,..n)}/\text{mean}(R_{eq@station(1,2,3,..n)}), \quad \text{Eq(2)}$$

208     Where, $R_{eq@station(1,2,3,..n)}$ is a vector that contains N values, each corresponding to an event-

209     averaged P/S ratio (i.e., calculated from up to 406 earthquakes) at a given station after the SNR

210     control. For each event (explosion or earthquake) recorded at a subset of m stations, its P/S ratio

211     is then corrected by dividing $C_{1,2,3...m}$. The outcome of site correction is discussed in section 4.5

212     and $R_{eq}$ for the entire array is presented in Figure 9b.

## 4. Results

## 4.1 Optimal frequency content

215     Array median P/S ratios for explosions are greater than for earthquakes in all frequency bands

216     (Fig. 5). However, the one SMAD intervals of P/S ratios overlap for lower frequency bands and

217     most narrow frequency bands. Distributions of P/S ratios are typically not Gaussian (i.e., long-

218     tailed), so the SMAD is preferred instead of the standard deviation. In general, higher bands

219     provide larger separation in the median P/S ratios between earthquakes and explosions, while

220     broader bands help to decrease the variance and stabilize discrimination of individual events

221     (Fig. 5). With respect to source discrimination, 10-18 Hz is the preferred frequency band for this

222     local distance data set. The 14-16 Hz band also performs well but for individual events it is more

223  variable and may reduce the number of SNR>2 phase windows for analysis. The 10-18 Hz array-
224  median P/S ratios and SMAD values for earthquakes and explosions are 1.72($\pm$0.34) and
225  0.46($\pm$0.19), respectively. P/S ratios for both earthquakes and explosions increase with
226  frequency, however the increase with frequency is greater for the explosions (Fig. 5). This result
227  from analysis of the three-component broadband array data is consistent with the observation
228  from the vertical-component geophones showing that S-wave amplitude diminishes relative to P
229  for higher frequencies (Fig. 3).

## 4.2 Distance, magnitude, and depth effects

231  We focus on results from distances <120 km because most of the available source-receiver pairs
232  are within that distance range (Fig. 4b). Median P/S ratios for earthquakes and explosions sorted
233  into six 20 km wide distance bins show little change at shorter distances and then increasing P/S
234  ratio with distance beyond ~60 km (Fig. 6a). The SMAD values are also larger at greater
235  distances due to the decreased number of samples within each distance bin. Given that distance
236  distributions of earthquakes and explosions are very similar in this study (Fig. 4b), we do not
237  apply distance corrections to the P/S ratios.

238      P/S ratios for earthquakes and explosions do not vary strongly with local magnitude (Fig.
239  5b). The median P/S ratios in most 0.2 $M_L$ bins overlap within one SMAD. There appears to be
240  an increase in P/S ratio for the largest explosion magnitude bin, but it is heavily influenced by
241  one explosion source with a particularly high P/S ratio resulting in a relatively large SMAD for
242  the highest magnitude bin among explosions (Fig. 5b). Earthquake P/S ratios remain low across
243  the local magnitude range. There is a slight decrease in the median earthquake P/S ratios for the
244  largest magnitudes ($M_L$>2.5), which is within the range of one SMAD. Given the weak distance
245  and magnitude dependences found in the P/S ratios, we did not apply the magnitude and distance
246  amplitude correction (MDAC) approach that is commonly used in regional distance studies
247  (Anderson et al., 2009).

248      Estimated earthquake source depths also do not exhibit a clear correlation with P/S ratios
249  (Fig. 7). As the area near Mount St. Helens is characterized by variable surface topography (i.e.,
250  ranging from 0 - 4 km, Fig. 1), we also take the topography into account and calculate the
251  relative distance to the surface (instead of absolute depth to sea level). Binning the earthquake
252  sources into four depth ranges (0-5, 5-10, 10-15, and 15-20 km) results in median P/S ratios that

253  all overlap with one SMAD. Depth uncertainties are highly variable depending on local array

254  coverage, but reported uncertainties from the Pacific Northwest Seismic Network for $M_L \geq 1$

255  earthquakes near Mount St. Helens are typically <1 km. In that context, results from binning

256  within 5-km depth intervals should not be strongly biased by inaccurate depth estimates. There is

257  a slight decrease in median P/S ratio from 0.3 at 0-5 km to 0.26 at 15-20 km, which is within the

258  SMAD (i.e., 0.07-0.14) for any of the source depth bins.


## 4.3 Number of stations in the seismic array

260  Bootstrap resampling of the three-component seismographs simulates the potential average

261  source discrimination performance with smaller seismic arrays (see section 3.2). The cumulative

262  distribution functions (CDFs) from resampling show that greater variations are observed for

263  explosion P/S ratios compared to earthquake P/S ratios (Fig. 8). Increasingly large seismic arrays

264  help reduce the overlap of earthquake and explosion P/S ratio distributions near a value of ~1

265  (Fig. 8). Separation of earthquake and explosion P/S ratios is subject to diminishing

266  improvement for arrays with over 16 stations. The improvement with increasing array size

267  appears modest in this context because the curves shown in Figure 8 are generated as averages

268  from 1,000 bootstrap resampling subsets (for 2, 4, 8, 16, 32 and 64 station cases). The variance

269  of P/S ratios is much larger for smaller arrays and its effect on discrimination performance is

270  discussed further in section 5.1.


## 4.4 Site response effects and shallow crust structure

272  Variations in station-averaged P/S ratios from earthquakes and borehole explosions exhibit

273  similar spatial patterns, despite the generally higher P/S ratios for explosions (Fig. 9). A scatter

274  plot of the earthquake station-averages versus the explosion station-averages shows an

275  approximately linear trend, with a correlation coefficient of 0.625. If the earthquakes are subset

276  by magnitude, the $M_L 1$-2 and $M_L \geq 2$ subsets yield correlations of 0.625 and 0.641 (these two

277  earthquake subsets are correlated at 0.91), respectively. These correlations indicate that

278  observations of small (~$M_L 1$-3) earthquakes may provide a useful basis for determining

279  empirical P/S amplitude site corrections that are applicable to explosions. However, the

280  improvement for the purpose of source discrimination appears to be subtle in this case (Fig. 10).

281 Application of site corrections increased the difference between the median explosion and
282 earthquake P/S ratios from 1.29 to 1.41, but the SMAD values also increased slightly for
283 explosions (0.38 to 0.43) and earthquakes (0.21 to 0.23). Furthermore, we found the site
284 correction terms are positively correlated with shallow crustal velocities. We adopted a velocity
285 model from a recent local travel time tomography study around Mount St. Helens (Ulberg et al.,
286 2020) and compared our site corrections with the Vp, Vs and Vp/Vs values beneath all the
287 station (site) locations. At 1km depth, the cross-correlation coefficients for the site correction
288 terms are 0.397 for Vp and 0.479 for Vs. Site corrections exhibit a much weaker correlation of -
289 0.086 with Vp/Vs (Fig. 9 d-f). The correlation coefficients remain comparable for all layers
290 around sea level (±2km) and they decrease at deeper depths, suggesting that the site correction is
291 influenced by uppermost crustal structure.

## 5. Discussion

### 5.1 Effectiveness for explosion discrimination

294 Receiver operating characteristic (ROC) curves show the diagnostic performance of binary
295 classifiers, in this case for explosions versus earthquakes, using a variety of discrimination
296 thresholds (Fig. 11). For each P/S ratio used as a discrimination threshold the point along the
297 ROC curve is determined by the rate of true positives and false positives. The optimal
298 discrimination threshold is chosen by maximizing the area under the curve (AUC), thus an ideal
299 classifier would plot in the upper left corner corresponding to 100% true positive and 0% false
300 positives. Application to P/S ratios with the preferred frequency band of 10-18 Hz and the full
301 data set of earthquakes and explosions shows 100% true positives and 4.93% (20 out of 406)
302 false positives for a P/S ratio threshold of 1.2 (Fig. 11a). Half of the events that create false
303 positives (i.e., P/S>1.2) are only recorded with <6 stations, reflecting the disadvantage of
304 insufficient station coverage for low-magnitude sources (Fig. S3). For comparison, a lower and
305 narrower frequency band of 6-8 Hz results in 86.95% true positives and 5.67% false positives for
306 a P/S ratio threshold of 0.9.
307 Discrimination performance generally improves with increasing event magnitude and
308 array size (Fig. 11c, d). Using all explosions and only $M_L \geq 2$ earthquakes, a threshold P/S ratio of
309 1.2 results in 100% true positives and 0% false positives. The $M_L$ 0.9-2.3 explosions produced an

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310   average of 66 stations with SNR>2, while the $M_L$ 1-2 earthquakes only produced an average of
311   47 stations with SNR>2. So, the diminished discrimination performance for $M_L$1-2 earthquakes
312   is likely influenced by SNR at individual stations and the total number of stations available for
313   analysis. ROC curves for seismic arrays composed of N=4, 16, and 64 seismographs illustrate
314   the effects of array size (Fig. 11c). For N=64, the bootstrap resampling results show that a
315   maximum AUC of 1 is consistently achieved for discrimination of explosions and $M_L \geq 2$
316   earthquakes. For N=16, the average of the AUC maxima is 0.98. For N=4, the average of the
317   AUC maxima reduces to 0.94 (Fig. 11d). However, the ROC curves for N=4 are much more
318   variable, including a few samples with near-random performance (Fig. 11c) and some extreme
319   cases reaching the AUC maximum (Fig. 11d). Generally, larger array size is expected to better
320   mitigate source/path variations. A separate bootstrap resampling test with an earthquake
321   epicenter nearly co-located with an explosion shows that <16 stations are sufficient for robust
322   discrimination (Fig. S4), suggesting that minimum number of stations could be much lower if the
323   earthquake and suspected explosion are closely located with depth difference (i.e., as an ideally
324   case). In this example the horizontal locations are separated by 2.4 km and the depths are
325   separated by 11.1 km.

326       Comparison of the new results from near Mount St. Helens with other local distance P/S
327   ratio discrimination tests may be influenced by different local propagation effects related to
328   geological settings as well as measurement approaches. A similar scale study using data covering
329   the Bighorn mountain range and adjacent sedimentary basins in Wyoming found effective P/S
330   discrimination at ~50-200 km distance but failure at <50 km distance (O'Rourke et al., 2016).
331   Below we consider two minor methodological differences between this study and O'Rourke et
332   al. (2016) that might contribute to the difference in performance. We acknowledge that different
333   local propagation effects may also be the cause of the difference, but a thorough re-analysis of
334   their data would be required to isolate that influence.

335       The first methodological difference is that we include the transverse component for P-
336   wave measurements (see section 3.1). Explosion recordings in this study commonly show
337   SNR>2 (our threshold for analysis) even for P-waves recorded on the transverse component (Fig.
338   2). If scattering of P energy onto the transverse component is important at distances <50 km, then
339   the inclusion of transverse component data may be beneficial for explosion discrimination. A
340   comparison of array-median P/S ratios with and without the transverse component shows that

341 separation of the explosion and earthquake ratios improves with inclusion of the transverse

342 component (Fig. 12). The increased separation is primarily due to increases in the explosion P/S

343 ratios, while changes in the earthquake P/S ratios are smaller. This is consistent with the

344 possibility that out-of-plane P-wave scattering is stronger for explosions than earthquakes. In this

345 case, the benefit of including transverse component P-wave energy spans our entire local

346 distance range (Fig. 12), but given that P/S ratios tend to increase with distance for both

347 approaches the incremental improvement is more important at smaller distances (≤60 km).

348 The second difference is that we did not apply the MDAC method for regional seismic

349 discriminants (Anderson et al., 2009) to correct for distance and magnitude effects. We chose not

350 to use MDAC because the distance and especially magnitude trends in our raw P/S ratios are

351 modest compared to the SMAD values for explosion and earthquakes, respectively (Fig. 6).

352 Additionally, O'Rourke et al. (2016) note that stations at their shortest distances (~20-40 km)

353 exhibit abnormally high MDAC corrections for explosions. We speculate that the amplitude

354 correction techniques developed for regional distances may need to be revised for shorter local

355 distances. Similar to the weak distance and magnitude trends found in this study, Zeiler and

356 Velasco (2009) reported negligible distance dependence of explosion discriminants (including

357 P/S ratios) within the local distance range. We suggest that the necessity and design of the

358 amplitude correction techniques for explosions observed at short distances (≤50 km) should be

359 further investigated.

360 In a recent local to regional distance study, Pyle and Walter (2019) investigated P/S ratio

361 discrimination with six chemical explosions from the Source Physics Experiment (Snelson et al.,

362 2013) and nearby tectonic earthquakes. Similar to O'Rourke et al. (2016) and this study, Pyle

363 and Walter (2019) found that P/S discrimination performance improved with increasing

364 frequency (≥6 Hz) and that array averaging of ~10 or more stations was needed to achieve clear

365 separation of explosion and earthquake P/S ratios. Four of the six explosions used in their study

366 had larger explosive yields than all the explosions in this study and the receiver spacing within

367 ~100 km was sparser, so relative to O'Rourke et al., (2016) there is less opportunity for

368 comparison of results at similar scales. The fact that we observe optimal performance for

369 somewhat higher frequencies in this study (e.g., 10-18 Hz) likely reflects our focus on shorter

370 source-receiver distances of <120 km compared to distances up to ~450 km in Pyle and Walter,

371 (2019).

## 5.2 Source effects

The geophone record sections show that S-waves emerge within the first few kilometers from the source, acknowledging that there is not sufficient sampling available to resolve S-wave emergence in detail at scales of hundreds of meters (or less) from the source (Fig. 3). Assuming the S-wave is effectively generated at the source for this data set, increasing P/S ratio with distance is expected if S-wave attenuation is greater than P-wave attenuation in the crust (e.g., Pyle et al., 2017). In contrast, a difference in source spectra may be needed to explain why explosion P/S ratios increase with frequency more than earthquake P/S ratios (Fig. 5).

Prior studies of $M \gtrsim 3$ chemical explosions and nuclear explosions indicate that explosion-induced crustal S-waves (Lg) exhibit lower corner frequencies than explosion-induced P-waves (local Pg & regional Pn phases; Xie, 2002; Fisk, 2006, 2007). In these studies, the S-wave corner frequency is typically about half of the P-wave corner frequency. Our results with lower yield chemical explosions are similar based on the increased separation of local distance explosion and earthquake P/S ratios at higher frequencies. The 10-18 Hz band may perform relatively well for P/S ratio discrimination because it partially overlaps the interval between the lower S-wave corner frequency and the higher P-wave corner frequency (e.g., Fisk, 2007). Using the Mueller-Murphy model corner frequency scaling from Fisk, (2007) the explosions in this study are expected to produce local P-wave corner frequencies of ~35-45 Hz and S-wave corner frequencies of ~17-23 Hz (Fig. S5). The 40 Hz sampling rate for most stations limited our analysis to frequencies below ~18 Hz, so only the highest frequencies in this study are near the expected S-wave corner. Extension of P/S ratio measurements to higher frequencies by collecting higher sample-rate data may be beneficial, especially for very short distances ($\leq$50 km) and low magnitudes ($M_L \leq 2$).

The physical controls on explosion-induced S-wave spectra remain debated. If a difference in source spectra explains the observed frequency dependence of P/S ratios, the physical cause may be rooted in the different P and S elastic length scales of sensitivity to near-source rock damage (Taylor, 2009). However, a variety of complicated and locally specific effects such as spall and free-surface topographic scattering are thought to influence explosion-induced S-wave spectra so there may not be one dominant explanation (e.g., Xie and Lay, 1994; Patton and Taylor, 1995). An alternative to explaining the frequency dependence of P/S ratios

402  with different source spectra is that explosions may preferentially excite shallower shear modes

403  that attenuate more rapidly with distance (e.g., Baker et al., 2004, 2012).

404       One explosion source in this study that produced a relatively low P/S ratio resulted in

405  clear surface evidence of shear slip at the source. The 'X4' explosion (Fig. 1b; also see Table S1)

406  detonated in a borehole with maximum depth of ~22 m depth resulted in a small normal fault

407  scarp with a length of ~15 m and vertical offset of up to ~20 cm. Consequently, it is unsurprising

408  that this explosion exhibited a relatively low P/S ratio of 1.27 (Fig. 10), which is near the optimal

409  binary classification cutoff of 1.2 for 10-18 Hz the data set including all $M_L{\geq}1$earthquakes (Fig.

410  11b). It is unknown if 'blind' shear failures that did not rupture the surface are common for other

411  low P/S ratio explosions.


412  ## 5.3 Combining local distance seismic discriminants

413  Ratios of different seismic magnitude metrics such as mb & Ms have long been used as an initial

414  or 'screening' step for source discrimination of M${\gtrsim}$3 events (Stevens and Day, 1985; Selby et

415  al., 2012; Ford and Walter, 2014). At local scales there is growing evidence that the difference

416  between the local (Richter) magnitude and the coda duration magnitude ($M_L$-$M_C$) can help

417  discriminate explosions, primarily because the magnitude difference is sensitive to source depth

418  (Holt et al., 2019; Voyles et al., 2020). Very shallow sources such as borehole explosions

419  preferentially produce longer coda leading to negative $M_L$-$M_C$. P/S amplitude ratios ignore the

420  extended coda and we find negligible depth dependence for P/S amplitude ratios in this study

421  (Fig. 7). Thus, the capabilities of the two discriminants appear to be rooted in complementary

422  aspects of the local seismic wavefield. A simple workflow to leverage the two approaches would

423  be using $M_L$-$M_C$ to screen for unusual events given that routine catalog generation can commonly

424  provide both parameters. Then events identified by screening could be further investigated with

425  P/S ratios. More thorough approaches to jointly classifying events based on $M_L$-$M_C$ and P/S

426  ratios should also be considered. P/S ratio measurements could also be automated to enhance

427  screening capabilities once travel time relationships and site corrections are developed for an

428  array. Additional use of more time-consuming moment tensor analysis of low magnitude local

429  distance events might be reserved for events flagged by $M_L$-$M_C$ and P/S screening (e.g., Alvizuri

430  and Tape, 2016).

## 6. Conclusions

P/S ratio discrimination of earthquakes and shallow single-fired borehole explosions at distances less than 120 km was tested with a dense broadband seismic array near Mount St. Helens. Taking advantage of the excellent coverage of both stations and sources, we evaluated the effects of frequency bands, source-receiver distances, source magnitudes and focal depths. Optimal separation of P/S ratios for the two sources types was found using a 10-18 Hz frequency band, which achieved explosion discrimination with 100% true positives and 4.93% false positives using the entire array. Randomly resampling to simulate smaller arrays shows performance of ≥98% true positives and ≤6.3% false positives for ≥16 stations. Performance becomes highly variable using ≤4 stations. Successful separation using the dense array and optimized frequency band left limited room to improve the performance, consequently our site corrections derived from 406 local earthquakes had a negligible effect on discrimination statistics. Despite the limited improvements in this case, we found that the site corrections correlate with Vp and Vs of the uppermost crust from a recent tomography study. Future studies using P/S ratios may benefit from local site corrections derived from small earthquake recordings or local tomography models.

## Acknowledgements

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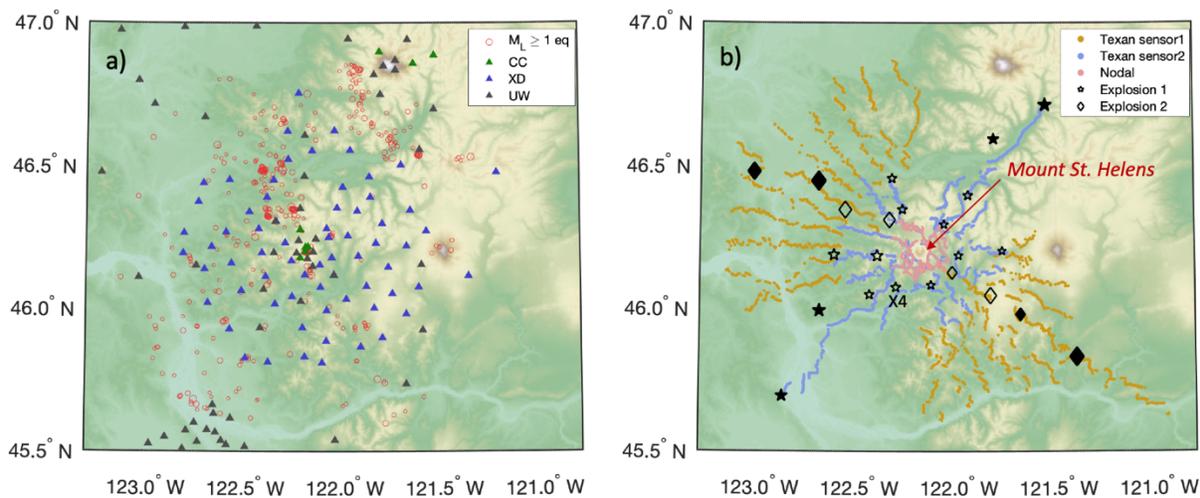
Figure 1. Study area maps with sources and receivers. A) Broadband stations used in the study are plotted along with the locations of $M_L \geq 1$ earthquakes that occurred between June 20, 2014 to August 31, 2018 (i.e., time span for XD array) within 75 km of Mount St. Helens. B) Vertical geophone locations are plotted along with single-fired explosion locations. Explosion locations loaded with 907 kg of explosive are denoted by filled symbols and locations loaded with 454 kg are denoted by hollow symbols. The symbol sizes for explosions are scaled by their $M_L$. Explosions labeled with stars were detonated during controlled source phase 1 (recorded by Texan 1 array; ~2370 nodes) and diamonds were detonated during phase 2 (recorded by Texan 2 array; ~2220 nodes). 'X4' denotes one explosion that is discussed in the text. Geophone locations in pink (~900 nodes) recorded both sets of explosions.
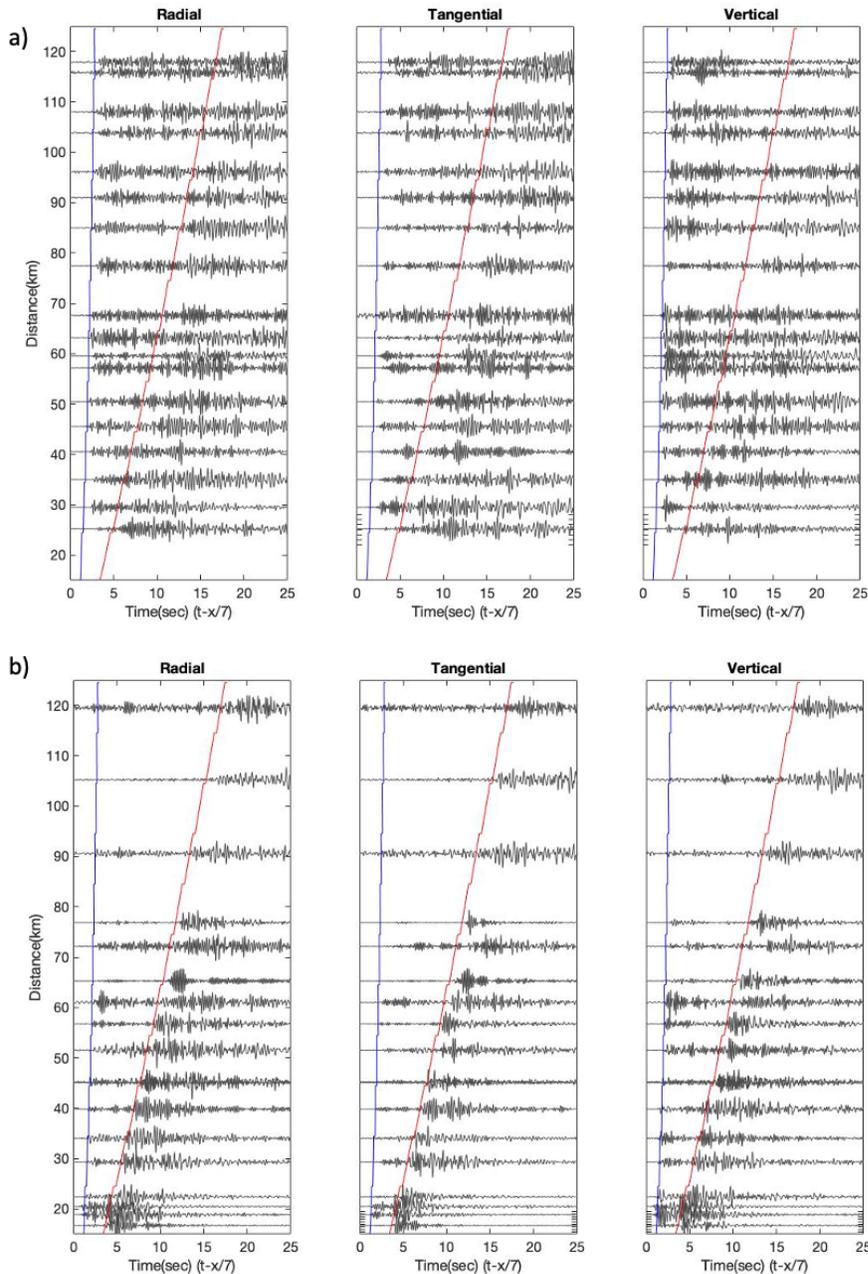
Figure 2. A) A record section is plotted for an explosion event with 616kg equivalent TNT load (Y8, located at the NW of array, $M_L$=1.9; see Table S1 for details). The event showed relatively strong S energy among explosions. Only 18 seismograms are shown for clarity. B) Record section for an equal-magnitude earthquake (2015 Feb 9, 05:45:35) with waveforms from 19 stations. Waveforms are filtered between 2-6 Hz at plotted with a reduced velocity of 7 km/s in both figures. The blue and red lines mark the predicted P and S wave arrivals, respectively, using a locally optimized 1D velocity model (Kiser et al., 2016; also see Fig. S1).
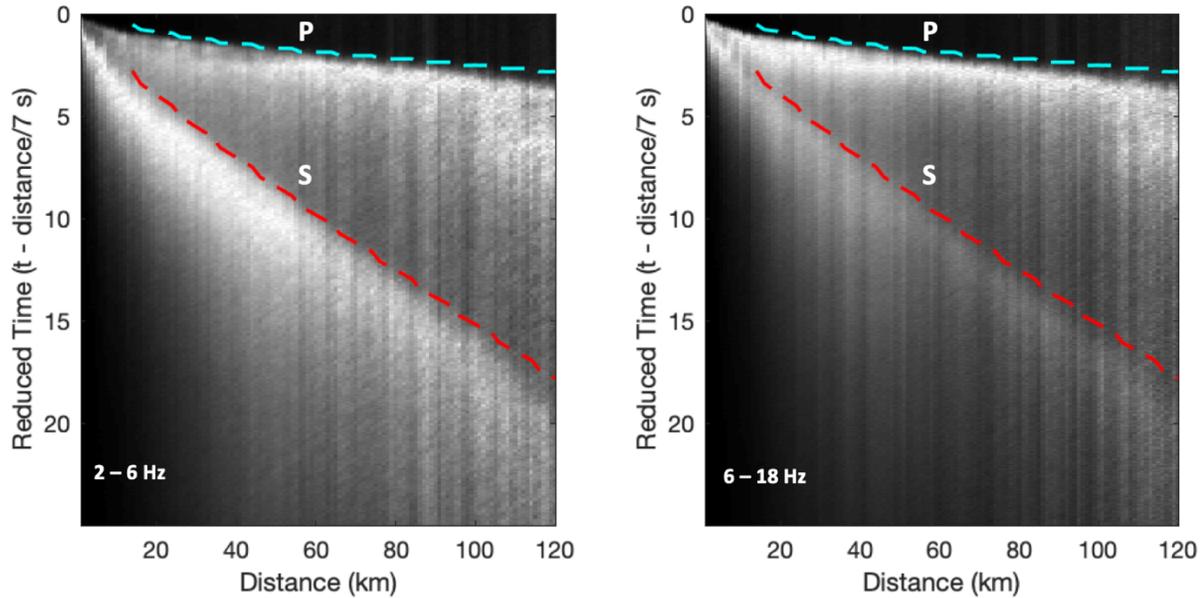
Figure 3. Stacked geophone recordings of explosions. A) Median envelopes of geophone seismograms from all 23 explosions in 1 km distance bins (70,944 source-receiver pairs). Seismograms were bandpass filtered from 2-6 Hz before envelope calculation and median-stacking. Median envelopes at each distance are normalized by their maximum. The cyan dashed line marks the beginning of the P window and the red dashed line marks the beginning of the S window. B) Same as A, except that seismograms were bandpass filtered from 6-18 Hz. The higher frequency band exhibits diminished S-wave amplitudes relative to P.

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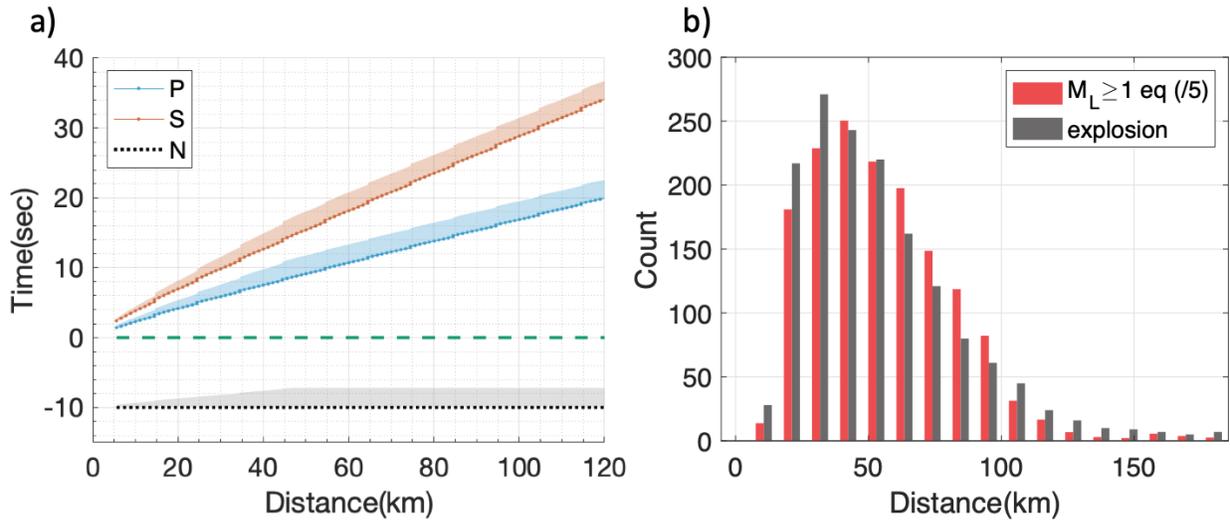


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611      Figure 4. Phase windowing and source-receiver distances. A) Phase windowing based on travel

612      times from a local 1D P and S velocity model (Kiser et al., 2016). B) Histogram of source-

613      receiver distances using 10 km bins up to 180 km. Note that the number of earthquake source-

614      receiver pairs in each distance bin is divided by 5 to better enable comparison of the two

615      distributions.

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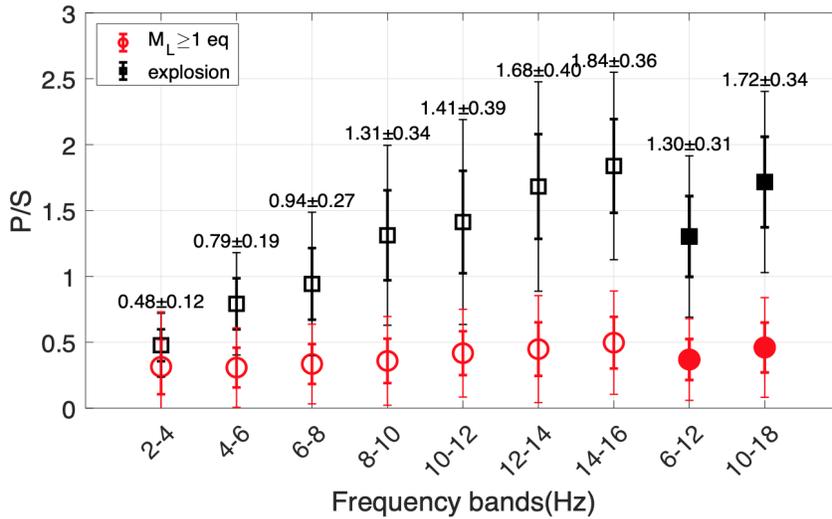


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618      Figure 5. Median P/S ratio for explosions and earthquakes at nine different frequency bins: seven

619      2-Hz narrow bins same as O'Rourke et al., 2016 and two wide bins (filled symbols). Bold error

620      bars are SMAD and thin bars are 2-scaled median absolute deviation (i.e., 2*SMAD). The

621      median±SMAD for explosions are labeled on top.
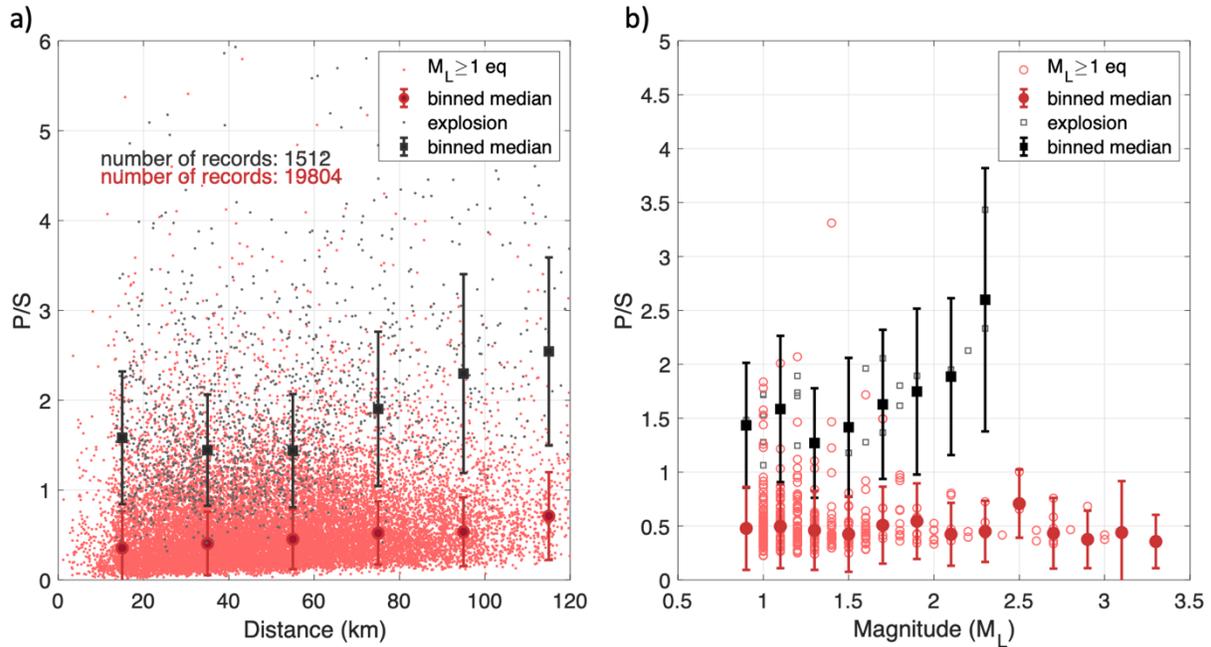
Figure 6. Distance and magnitude effects. A) P/S versus distance averaged over 20 km with step size of 20 km for all source-station pairs of earthquakes and explosions. B) P/S ratio versus magnitudes with bin size of 0.2. Error bars indicate the SMAD calculated from station-based ratios of all events within each magnitude bin.
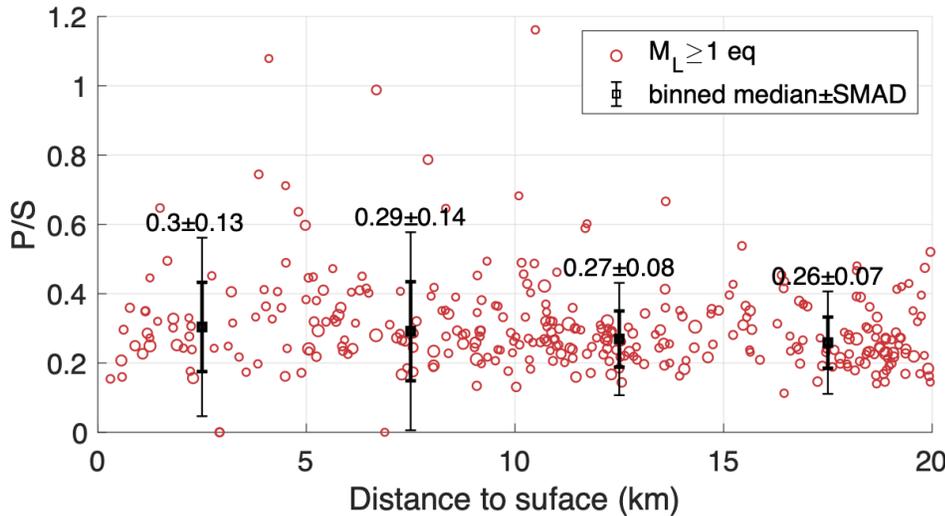


Figure. 7. Earthquake P/S ratios versus source depth below surface. Red circles show array-averaged P/S ratios for individual earthquakes, with circle size scaled by $M_L$ (1-3.3). The four bold (thin) bars show median SMAD (2*SMAD) within depth bins: 0-5 km, 5-10 km, 10-15 km, and 15-20 km.

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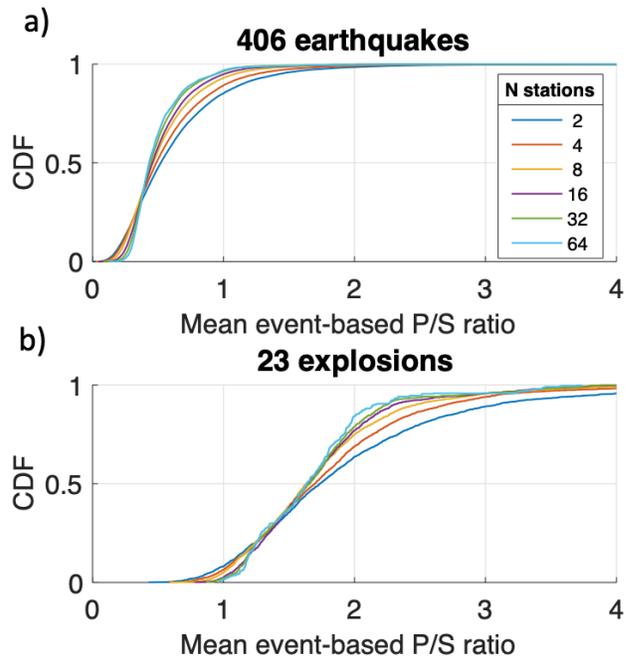
Figure 8. Effects of array size on P/S ratios. A) Cumulative distribution functions (CDF) for P/S ratios of earthquakes as a function of array size for 2, 4, 8, 16, 32, and 64 stations. Each event ratio is calculated from the median of P/S ratio measured at N stations while the final CDF curves are generated from the mean of all events. B) Same as (A) except for explosions rather than earthquakes.
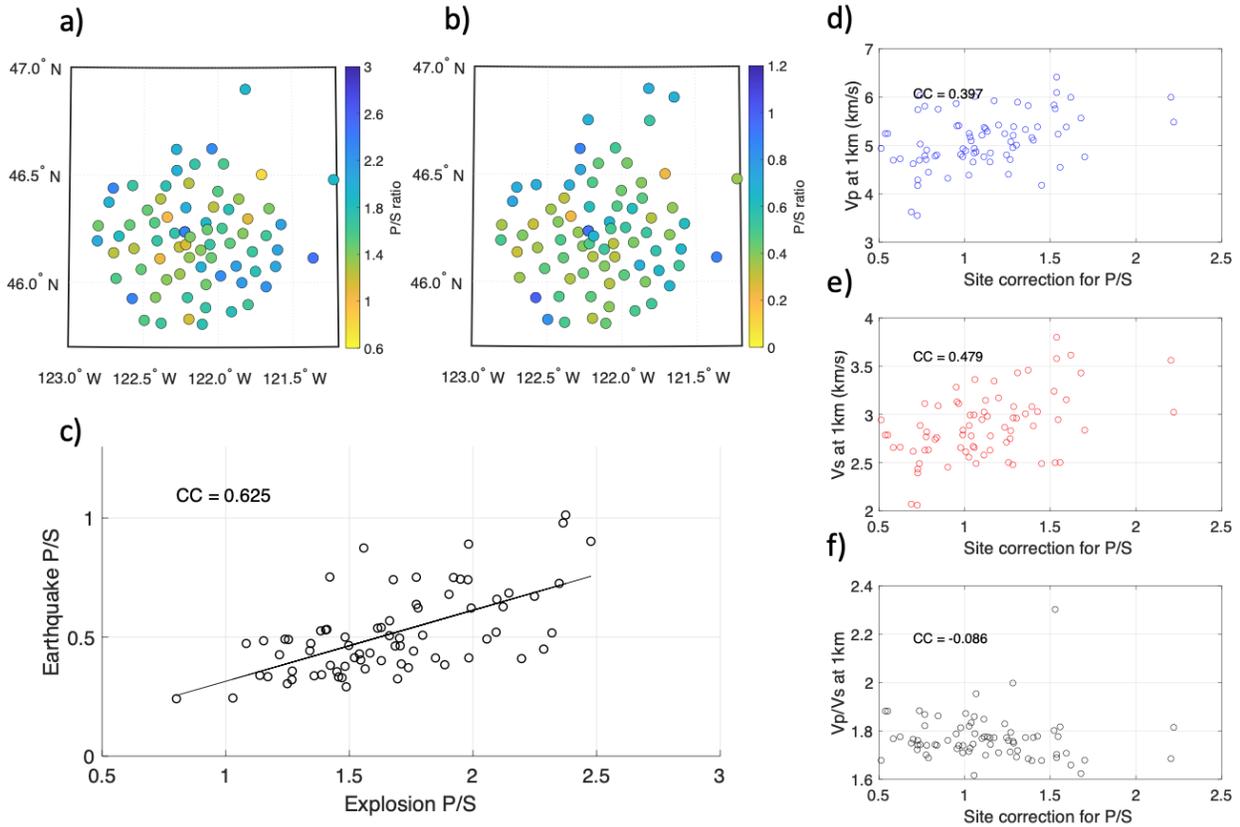
Figure 9. Site corrections and crustal seismic structure. A) Map of station-averaged P/S ratios for explosions. B) Map of station-averaged P/S ratios for earthquakes. Note that the color scale is different from A so that the similarity in spatial pattern is easier to identify. C) Scatter of station-averaged P/S ratios for earthquakes versus explosions. D-F) Scatter plots of site correction versus Vp, Vs, and Vp/Vs at 1 km depth from Ulberg et al., 2020. CC: cross-correlation coefficient.
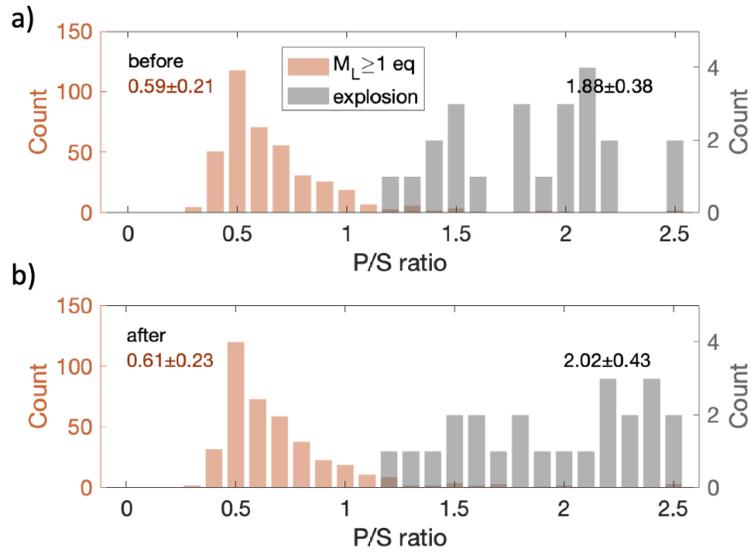
647
648   Figure 10. Site corrections and P/S ratios. A) Histograms of array-averaged P/S ratios for all

649   explosions and earthquakes. B)  Same as A, except the P/S site correction term is applied. The

650   median±SMAD of the P/S ratio distributions are labeled in corresponding colors.
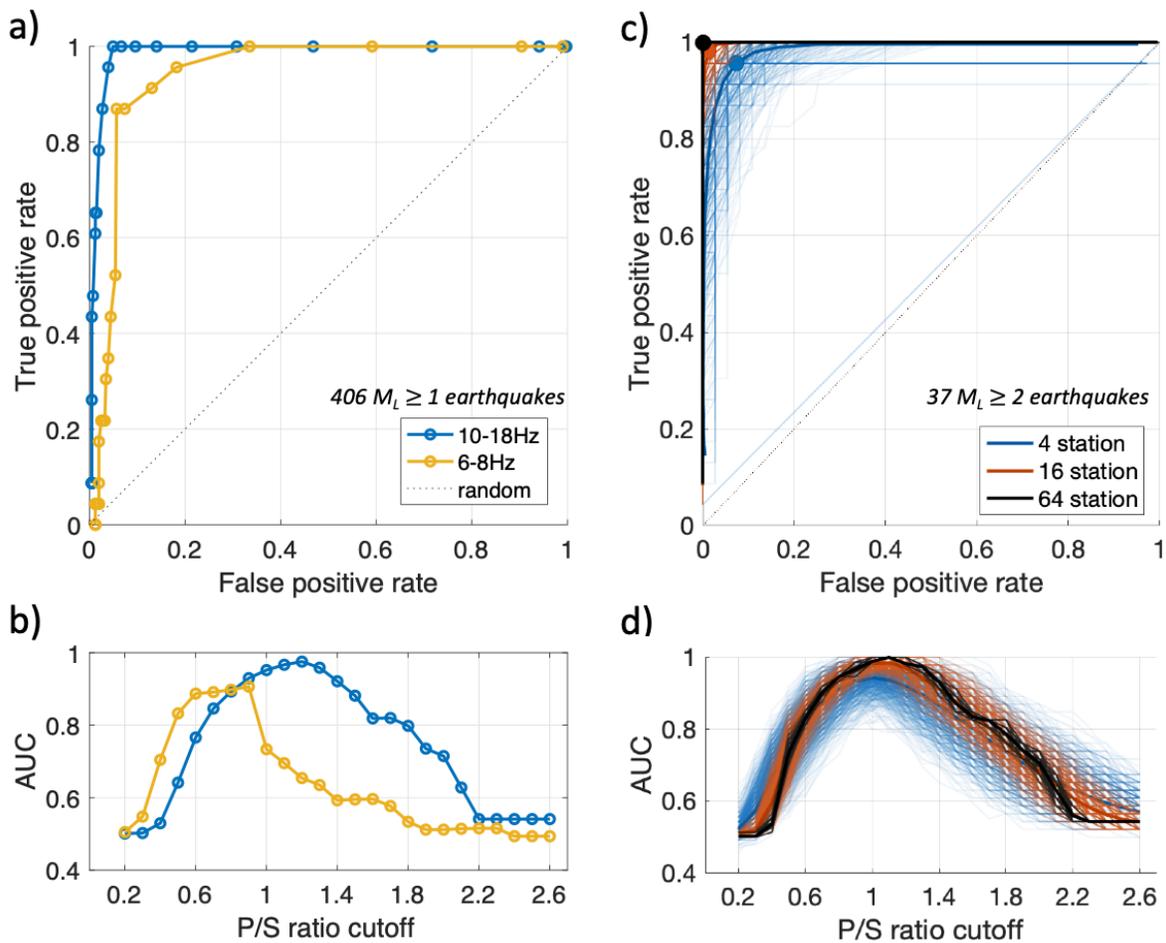
Figure 11. A) ROC curves for P/S ratio threshold optimization at 10-18 Hz and 6-8 Hz. B) Area under curve (AUC) for searched P/S ratio thresholds (0.1 step size from 0.2 to 2.6). The preferred P/S ratio cutoff is 1.2 at 10-18Hz (AUC=0.98), which leads to a true positive rate of 100% and false positive rate of 4.93%. C&D) Similar to A&B) but showing bootstrapping results using only $M_L \geq 2$ earthquakes. The bold lines are averaged curves from 1000 subsets and the three dots mark the highest AUC. Note that a few outliers exhibit near-random performance for the 4-station case (light blue near the diagonal in C) and a few others may reach AUC=1 while the averages (bold curves) are below 1 for 4-station and 16-station cases (in D).
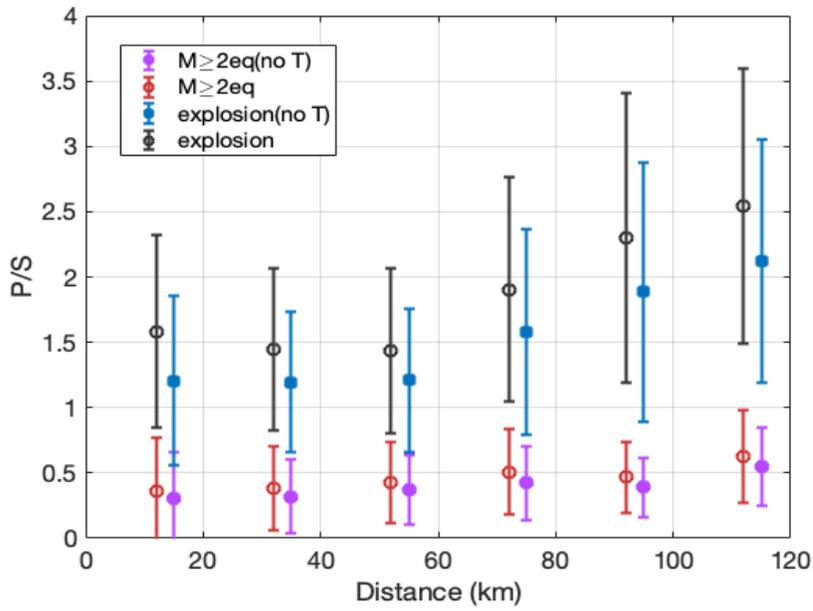
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661    Figure 12. Effect of transverse component P-wave energy on P/S ratios. The error bars show

662    medians and SMAD values of 10-18 Hz P/S ratios in 20 km bins for all source-station pairs

663    (similar to Figure 6a but using only $M_L \geq 2$ earthquakes and all explosions).

664