

1 **Spatial Heterogeneity of the Respiratory Health Impacts of Wildfire Smoke PM_{2.5} in**
2 **California**

3 **V. Do^{1*}, C. Chen^{2*}, T. Benmarhnia^{2,3}, and J. A. Casey^{1,4}**

4 ¹Department of Environmental Health Sciences, Columbia University Mailman School of Public
5 Health, New York, NY, USA ²Scripps Institution of Oceanography, UC San Diego, La Jolla,
6 CA, USA. ³Irset Institut de Recherche en Santé, Environnement et Travail, UMR-S 1085,
7 Inserm, University of Rennes, EHESP, Rennes, France. ⁴Department of Epidemiology,
8 University of Washington, Seattle, WA, USA.

9 *Authors contributed equally to this manuscript.

10 Corresponding author: Chen Chen (chc048@ucsd.edu)

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12 **Author contribution:**

13 Conceptualization: CC, TB, JAC

14 Methodology: CC, TB

15 Formal analysis: CC

16 Investigation: VD, CC, TB, JAC

17 Visualization: VD, CC, JAC

18 Supervision: TB, JAC

19 Writing—original draft: VD, CC

20 Writing—review & editing: VD, CC, TB, JAC

21

22 **ORCID ID:** VD (0000-0001-6127-4361); CC (0000-0002-0632-946X); TB (0000-0002-4018-
23 3089); JAC (0000-0002-9809-4695).

24

25 **Key points**

- 26 • Statewide, exposure to wildfire PM_{2.5} is associated with increased odds of respiratory
27 acute care utilization in California.
- 28 • The wildfire PM_{2.5}-health association varies spatially across air basins, counties, and ZIP
29 Code Tabulation Areas.
- 30 • Areas with higher proportions of Black and Pacific Islander populations and less
31 affluence had worse wildfire PM_{2.5}-related outcomes.

32 **Abstract (word limit 250)**

33 Wildfire smoke fine particles (PM_{2.5}) are a growing public health threat as wildfire events
34 become more common and intense under climate change, especially in the Western United
35 States. Studies assessing the association between wildfire PM_{2.5} exposure and health typically
36 summarize the effects over the study area. However, health responses to wildfire PM_{2.5} may vary
37 spatially. We evaluated spatially-varying respiratory acute care utilization risks associated with
38 short-term exposure to wildfire PM_{2.5} and explored community characteristics possibly driving
39 spatial heterogeneity. Using ensemble-modelled daily wildfire PM_{2.5}, we defined a wildfire
40 smoke day to have wildfire-specific PM_{2.5} concentration $\geq 15 \mu\text{g}/\text{m}^3$. We included daily
41 respiratory emergency department visits and unplanned hospitalizations in 1,396 California ZIP
42 Code Tabulation Areas (ZCTAs) and 15 census-derived community characteristics. Employing a
43 case-crossover design and conditional logistic regression, we observed increased odds of
44 respiratory acute care utilization on wildfire smoke days at the state level (odds ratio [OR] =
45 1.06, 95% confidence interval [CI]: 1.05, 1.07). Across air basins, ORs ranged from 0.88 to 1.57,
46 with the highest effect estimate in San Diego. A within-community matching design and spatial
47 Bayesian hierarchical model also revealed spatial heterogeneity in ZCTA-level rate differences.
48 For example, communities with a higher percentage of non-Hispanic Black or Pacific Islander
49 residents had stronger wildfire PM_{2.5}-outcome relationships, while more air conditioning and tree
50 canopy attenuated associations. We found an important heterogeneity in wildfire smoke-related
51 health impacts across air basins, counties, and ZCTAs, and we identified characteristics of
52 vulnerable communities, providing evidence to guide policy development and resource
53 allocation.

54 **Keywords:**

55 Wildfire, smoke, acute care utilization, spatial heterogeneity, vulnerability, environmental justice

56 **Plain language summary (word limit 200)**

57 Wildfire smoke is a growing public health threat, one becoming more pressing as climate change
58 progresses. People are exposed to different levels of wildfire smoke. People also have different
59 abilities to protect themselves from smoke exposure based on their job, housing quality, or other
60 factors. In addition, people have different physiological responses to wildfire smoke. Therefore,
61 the relationship between wildfire smoke and health could vary across the state of California. We
62 conducted a study using modeled daily wildfire smoke fine particle concentrations and daily
63 respiratory acute care utilizations from 2006-2019 in California. We estimated area-specific
64 wildfire smoke and acute care utilization associations at state, air basin, county, and ZIP Code
65 Tabulation Areas levels. We found different associations across the state, with the strongest
66 association in San Diego air basin. San Francisco Bay air basin had the highest number of acute
67 care utilizations attributable to wildfire smoke due to their large population. We identified
68 several community characteristics that may have explained the observed spatial differences,
69 including higher proportions of Black and Pacific Islander populations and less community
70 affluence. Our findings support the allocation of scarce resources to areas and communities more
71 vulnerable to wildfire smoke to improve population health in a changing climate.

72 **1 Introduction**

73 Wildfire PM_{2.5} is a growing threat to public health. Drier conditions and warmer
74 temperatures in the Western United States (US) contribute to wildfire events that are more
75 common, intense, and expansive in scope (Abatzoglou, 2013; Littell et al., 2009; Mueller et al.,
76 2020; Westerling et al., 2006). The resulting wildfire PM_{2.5} has increased overall trends in
77 ambient air pollution, counteracting policy efforts to improve air quality (Burke et al., 2023;
78 Ford et al., 2018). Wildfire PM_{2.5} can infiltrate the lungs and precipitate respiratory events
79 through inflammation and oxidative stress (Xing et al., 2016). In previous epidemiological
80 studies, exposure to wildfire smoke has been linked to a variety of adverse health effects,
81 particularly for respiratory conditions (Aguilera et al., 2020, 2021; Gould et al., 2024; Kondo et
82 al., 2019; Reid & Maestas, 2019). Recent toxicologic and epidemiologic studies found that
83 wildfire PM_{2.5} can have a higher adverse health impact on the pulmonary system than PM_{2.5} from
84 other sources (Aguilera et al., 2021; Kim et al., 2018; Wegesser et al., 2009), and disregarding
85 the differential dose-response of wildfire PM_{2.5} led to an underestimation of PM_{2.5} related health
86 burden (Darling et al., 2023), which warrants independent studies of wildfire PM_{2.5} health
87 impacts.

88 Wildfire PM_{2.5} concentrations vary across space and time, and so do the corresponding
89 health effects. Proximity to wildfires, wind direction, and social factors determine levels of
90 wildfire PM_{2.5} exposure (Casey et al., 2023; Reid & Maestas, 2019). For example, in the past few
91 years, several cities experienced the worst 24-hour average PM_{2.5} levels recorded on Earth
92 because of nearby wildfires (Masters, 2018; Osaka, 2022). Additional spatially-varying factors
93 including meteorologic and topographic conditions such as the Santa Ana winds (Gershunov et
94 al., 2021) may shape the spatial distribution of wildfire PM_{2.5} and health outcomes (Leibel et al.,
95 2020). Furthermore, the toxicity of wildfire PM_{2.5} could change across space as the PM_{2.5} ages
96 when traveling (O'Dell et al., 2020). Few studies have accounted for the spatial dependence in
97 wildfire PM_{2.5} exposure on health and those that did focused on a single wildfire event affecting
98 a small geographical area (i.e., San Diego air basin) (Aguilera et al., 2020) or only accounted for
99 spatial autocorrelation among areas closely located (Reid et al., 2016). Evaluating how health
100 effects related to wildfire PM_{2.5} are distributed across larger geographical areas involving more
101 wildfire events could inform future mitigation efforts to target specific areas and shape
102 regulations to better prepare for wildfire PM_{2.5}-related health burden.

103 Community characteristics like socioeconomic status and racial/ethnic composition can
104 drive spatial differences in the health impacts of wildfire PM_{2.5} through both exposure disparities
105 and differential response. For example, due to historical discriminatory practices, disparities in
106 housing quality exist such that communities of color tend to have lower-quality, substandard
107 housing (Hernández & Swope, 2019; Jacobs, 2011). Given wildfire PM_{2.5}'s ability to easily
108 infiltrate the home (Mendoza et al., 2021), communities of color may be more exposed to
109 wildfire PM_{2.5}. Differences in community characteristics could also lead to spatially varying
110 physiological response and behavioral adaptations towards wildfire PM_{2.5}. Lower-income
111 communities have more constraining choices to protect themselves from wildfire PM_{2.5} (Burke et
112 al., 2022). Minoritized groups with worse baseline health conditions due to social
113 marginalization and systemic racism will likely have worse health responses to wildfire PM_{2.5}
114 (Berberian et al., 2022; Smith et al., 2022). Moreover, the effects of wildfire PM_{2.5} may be worse
115 in communities that already experience a disproportionately high burden of other environmental
116 exposures due to the potential synergistic effects of compound exposures (C. Chen et al., 2023).
117 Taken together, there is a need for further research on community characteristics as drivers of the
118 spatially varying health effects of wildfire PM_{2.5} (Marlier et al., 2022).

119 Here, we aimed to investigate the spatially-varying relationship between wildfire PM_{2.5}
120 exposure and respiratory acute care utilizations and to examine whether various community
121 characteristics explained the observed spatial heterogeneity in impact of wildfire PM_{2.5} on
122 respiratory acute care utilization. We used ZIP Code Tabulation Area (ZCTA)-level ensemble-
123 modelled daily wildfire PM_{2.5} concentrations and daily respiratory acute care utilizations in
124 California from 2006-2019 to estimate spatially-varying health effects across four spatial units:
125 state, air basin, county, and ZCTA. We also examined community vulnerability factors of such
126 health effects at the ZCTA level.

127

128 **2 Materials and Methods**

129 **2.1 Data sources and study population**

130 We restricted all analyses to 1,396 ZCTAs in California satisfying two criteria: 1) having
131 a population $\geq 1,000$ in the 2010 US Decennial census for statistical power consideration
132 (Bureau, 2021a); and 2) having at least one wildfire smoke day during the study period (2006-

133 2019). The second criterion was a requirement for this study because unexposed ZCTAs do not
134 contribute information to the case-crossover or within-community matched designs (Mittleman
135 & Mostofsky, 2014; Schwarz et al., 2021). We chose ZCTA as the main spatial unit in our
136 analyses because of the spatial resolution of health outcome.

137 **2.1.1 Wildfire smoke day**

138 We utilized a previously developed time-series dataset for daily wildfire-specific PM_{2.5}
139 concentration at the ZCTA level (Aguilera et al., 2023) to identify smoke days. Briefly, Aguilera
140 *et al.* (2023) first generated the ZCTA-specific daily PM_{2.5} concentrations (all sources) from a
141 stacked ensemble model using several data-adaptive algorithms and many predictors (e.g., air
142 monitor data, satellite-derived aerosol properties, meteorological conditions, and land-use
143 information). Then, they identified ZCTA-days exposed to smoke plumes using validated NOAA
144 Hazard Mapping Systems (HMS) products. Next, they applied a chained random forest algorithm
145 to impute counterfactual non-wildfire PM_{2.5} concentrations in ZCTA-days with wildfire smoke
146 (expected PM_{2.5} concentrations in the absence of the smoke) (Aguilera et al., 2023). The
147 wildfire-specific PM_{2.5} is the difference between the estimated daily PM_{2.5} concentrations from
148 the ensemble model and the imputed non-wildfire smoke PM_{2.5} concentrations in each ZCTA.
149 For each ZCTA, we defined a wildfire smoke day as a day with wildfire-specific PM_{2.5}
150 concentration $\geq 15 \mu\text{g}/\text{m}^3$, a threshold based on the World Health Organization guideline for 24-
151 hour PM_{2.5} (Organization, 2021).

152 **2.1.2 Health outcomes**

153 We used the Patient Discharge Data and Emergency Department Data collected by the
154 California Department of Health Care Access and Information (CA.gov, 2023). This dataset
155 contains all acute care utilizations that are not prearranged in the general population of
156 California, including unscheduled hospitalizations and emergency department visits. Emergency
157 department visits that led to hospitalizations were recorded as unscheduled hospitalizations only.
158 For each ZIP code, we identified daily respiratory acute care utilizations with primary diagnosis
159 codes recorded as diseases of the respiratory system (see the list of included *International*
160 *Classification of Diseases* codes in supplementary Text S1). The ZIP code was based on the
161 patients' residential address at the time of the visit. Since the US Census Bureau created ZCTAs

162 to represent populated areas of the ZIP code service area, with the latter being a sum of service
163 routes by the United States Postal Service, we treated them as the same in analysis and used
164 ZCTA in the remainder of this manuscript.

165 **2.1.3 Community characteristics**

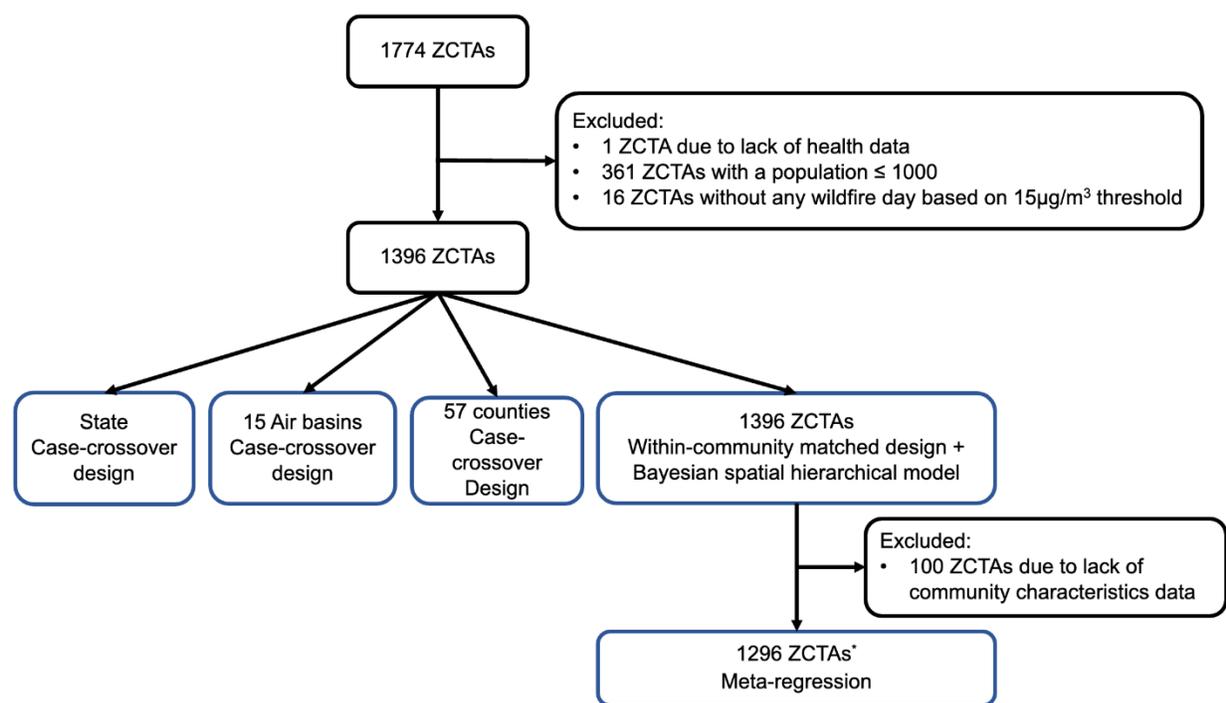
166 To explore whether the effects of wildfire smoke days varied by community
167 characteristics, we used 15 ZCTA-level variables. Communities of color have a greater risk for
168 wildfire-related health outcomes possibly due to disproportionate cumulative environmental
169 burden and systemic discrimination (Berberian et al., 2022), so we obtained the proportions of
170 self-reported race/ethnicity (separate proportions of non-Hispanic white, Black, Asian, American
171 Indian or Alaska Native, Native Hawaiian or Other Pacific Islander, and Hispanic residents)
172 from the 2010 US Decennial Census. We also collected population density from the same data
173 source (2010 Census, 2023). Additional variables were obtained from the Public Health Alliance
174 of Southern California Healthy Places Index report version 3.0 (Healthy Places Index, 2023;
175 Maizlish et al., 2019), which are mostly based on averages of the American Community Survey
176 data from 2015 to 2019. Included variables are the proportion of employment among those ages
177 20 to 64, the proportion of 25 and older with a bachelor's degree or higher, the proportion of
178 insured among those aged 18-64, the proportion of the population with an income that is greater
179 than 200% of the federal poverty level, per capita income in the US. dollars, the percentage of
180 households with access to an automobile, and the population-weighted percentage of area with
181 tree canopy. We also obtained the ZCTA-level percentage of households with access to central
182 air conditioning (A/C) from the California Residential Appliance Saturation Study survey
183 (KEMA, Inc., 2010) because air conditioning access may buffer against air pollution exposure
184 (Liang et al., 2021). Table S2 provided detailed descriptions and sources for each variable of
185 community characteristics. All variables other than race/ethnicity and population density were
186 coded such that a higher value corresponds to a higher proportion of economically advantaged
187 subpopulations.

188 **2.2 Statistical analyses**

189 We estimated the health impacts of wildfire PM_{2.5} concentrations on respiratory acute
190 care utilizations at four geographical levels: state, air basin, county, and ZCTAs. The California

191 Air Resources Board designates 15 air basins, geographies with distinct meteorological
 192 conditions to regionally distribute resources to address emissions. Each air basin contains
 193 between one and 11 counties (California Air Resources Board, 2023). We assigned ZCTAs to a
 194 county and an air basin based on the location of their population-weighted centroids. Counties
 195 and air basins with no ZCTAs that had a population ≥ 1000 and experienced a wildfire smoke
 196 day were excluded from analyses (Figure 1). In meta-regression to investigate the influence of
 197 community characteristics on ZCTA-specific effect estimates, we further excluded 100 ZCTAs
 198 without complete community characteristics data. All analyses were conducted in R version
 199 4.1.0 (R Core Team, 2021) and the analytic code is available at GitHub:
 200 https://github.com/benmarhnia-lab/cal_wildfire_spatial.git.

201



202

203 Figure 1. Flowchart of the California study population and exclusion criteria (black boxes) and
 204 method utilized in each set of analyses (blue boxes).

205 *For analysis of air conditioning prevalence, we further excluded 274 ZCTAs (1122 in meta-
 206 regression) due to data missingness.

207 **2.2.1 Case-crossover design for health analyses at state-, air basin-, and county-level**

208 We implemented the time-stratified case-crossover design to evaluate the effects of
209 wildfire PM_{2.5} on daily respiratory acute care utilization at the state level, air basin level, and
210 county level (Maclure, 1991; Mittleman, 2005). In the time-stratified case-crossover design, we
211 matched each day when an acute care utilization occurred (case) to other days of the same
212 weekday during other weeks of the same month in the same ZCTA (controls). This study design
213 compared exposures of a case to themselves at different times and accounts for individual-level
214 confounders (e.g., age, race/ethnicity and sex) and temporal trends of the exposure beyond a
215 month (Maclure, 1991; Mostofsky et al., 2018). For state-level analysis, we ran a weighted
216 conditional logistic regression to account for the matching procedure and included matched case
217 and control sets from all 1,396 ZCTAs to estimate the odds ratio (OR) of exposure to wildfire
218 smoke and respiratory acute care utilizations, with weight equal to the number of acute care
219 utilizations in the case day. For air basin-level and county-level analyses, we ran the same
220 conditional logistic regressions using only the matched sets in ZCTAs whose population-
221 weighted centroids fall within the corresponding air basin or county. These stratified analyses
222 assume that wildfire smoke has the same effect across all ZCTAs within the same air basin or
223 county. We used the “survival” package for conditional logistic regression (Therneau et al.,
224 2023).

225 To incorporate the total acute care utilization counts during wildfire smoke days and
226 provide estimates of the health burden, we calculated the population attributable number of acute
227 care utilizations due to wildfire PM_{2.5} during the study period at the county, air basin, and state
228 levels. For each geographical area, we calculated the population attributable number as the
229 product of area-specific attributable fraction (one minus the inverse of area-specific OR) (Lash et
230 al., 2021) and the area-specific total number of acute care utilizations among all wildfire smoke
231 days during the study period.

232 **2.2.2 Within-community matched design coupled with spatial Bayesian hierarchical model**
233 **for ZCTA-level health analyses**

234 To explore finer scale spatially varying effects, we used a previously developed within-
235 community matched design to estimate the ZCTA-specific effect of wildfire PM_{2.5} on the risk of
236 daily respiratory acute care utilization (C. Chen et al., 2023). Specifically, we identified matched

237 controls for each day exposed to wildfire smoke as non-wildfire smoke days of the same year
238 and ZCTA, and within the window of 30 calendar days before or after the wildfire smoke day.
239 We excluded days within the window of 3 calendar days before or after any wildfire smoke day
240 from the controls to avoid spillover effects from other wildfire days. To estimate rate differences,
241 we calculated the difference between the acute care utilization rate on the exposed case day and
242 the weighted averages of acute care utilization rates among non-exposed control days. Acute care
243 utilization rates on exposed case days were the count of acute care utilizations divided by ZCTA
244 population size from the 2010 US Decennial Census. Weighted averages for non-exposed control
245 days were weighted acute care utilization rates based on inverse temporal distance to exposed
246 day (i.e., one divided by number of days to the matched exposed day). We used the average rate
247 difference of all exposed days within a ZCTA to represent the ZCTA-specific rate difference and
248 scaled the rate difference to per 100,000 person-day.

249 Since ZCTAs closer together might exhibit similar effects from a wildfire smoke day
250 compared to ZCTAs farther away, we used a spatial Bayesian hierarchical model (BHM) to
251 leverage this spatial autocorrelation and increase the precision of our rate difference estimates
252 (Schwarz et al., 2021). We included a covariance structure to leverage this spatial autocorrelation
253 across ZCTAs and used an empirical semivariogram to identify the shape and starting values of
254 the covariance structure (spherical shape and 2, 16, and 8 for sill, nugget, and range parameters
255 respectively) (Bivand et al., 2013). We also used flat priors to introduce minimal prior
256 information into the Bayesian model: inverse gamma distribution with scale and shape equal to
257 0.001 for the sill and nugget parameters, and uniform distribution from 0.001 to 6 for the range
258 parameter. We used 10,000 Monte Carlo Markov chain samples with 75% burn-in to estimate the
259 ZCTA-specific rate differences after spatial pooling. Additionally, we calculated the signal-to-
260 noise ratio to represent the precision of the estimates, which is equal to the ratio between the
261 mean of the rate differences in the recovered samples and the corresponding standard deviation.
262 The signal-to-noise ratio allows us to have a mappable measure of statistical precision and values
263 higher than 2 are considered precise. We used the “spBayes” package in R for the spatial BMH
264 (Finley et al., 2015).

265 **2.2.3 Effect modification by community characteristics at the ZCTA level**

266 We evaluated potential effect modification by community characteristics on the effect of
267 a wildfire smoke day on acute care utilization at the ZCTA level using meta-regression. For each
268 community characteristic, which was selected *a priori*, we ran a meta-regression of the pooled
269 ZCTA-specific rate difference on the community characteristic. To preserve statistical power, we
270 excluded 100 ZCTAs without complete data for 14 community characteristics other than A/C
271 prevalence, and we excluded 274 ZCTAs for meta-regression of the A/C prevalence. Our
272 estimates are reported as rate difference per interquartile range increase of the community
273 characteristic. We used the “meta” package for meta-regression (Balduzzi et al., 2019).

274 **2.3 Sensitivity analyses**

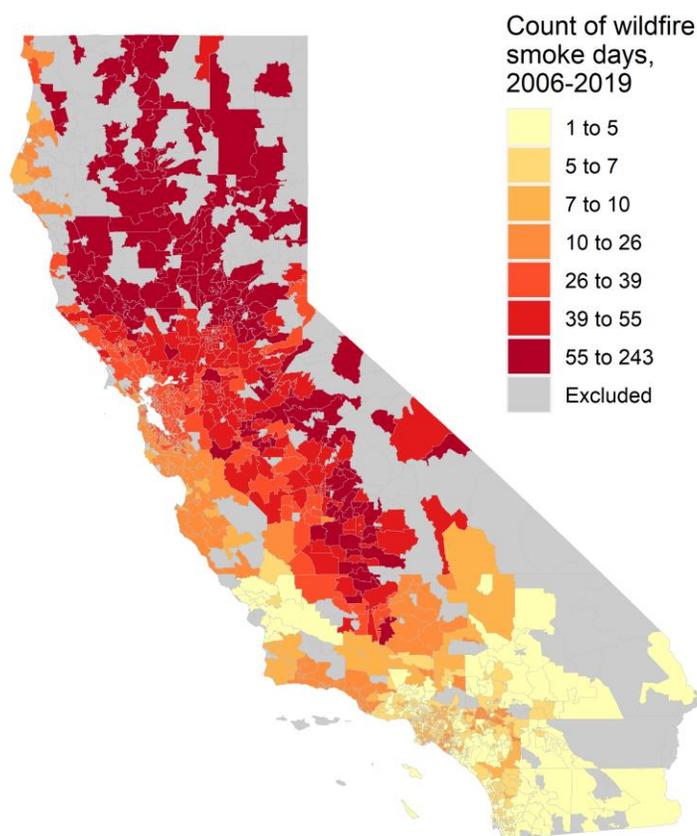
275 Since atmospheric aridity might affect the probability of wildfire occurrence and ambient
276 temperature is a known risk factor for respiratory acute care utilization, we conducted sensitivity
277 analyses for the state-level case-crossover analyses by including two forms of daily ambient
278 temperature as a linear term or a natural cubic function with six degrees of freedom. We
279 calculated daily ambient temperature at the population-weighted centroid of each ZCTA based
280 on an existing 4km×4km temperature surface (Daly et al., 2008). We also evaluated the
281 individual 1-day lagged effect of wildfire smoke on acute care utilization in a case-crossover
282 analysis.

283 To evaluate the robustness of the within-community matched design and spatial BHM,
284 we conducted a sensitivity analysis using informative priors employed in previous studies for
285 the sill and nugget in the spatial BHM, which are inverse gamma distributions (2 for shape and
286 1/starting value for scale) (C. Chen et al., 2023). This sensitivity analysis tested the robustness of
287 the spatial BHM towards prior specification and the informative priors used here give more
288 weight to our interpretation of the empirical semivariogram while the flat priors in main analysis
289 were more data-driven. We also used community-level socioeconomic information from the
290 Healthy Places Index report version 2.0 in the meta-regression, which is based on averages of
291 2011 to 2015, earlier than the averages of 2015 to 2019 in the main analysis (Delaney et al.,
292 2018).

293 **3 Results**

294 **3.1 Characteristics of ZCTAs, wildfire smoke days, and respiratory acute care utilizations**

295 Our study spanned 2006-2019 and included 1,396 California ZCTAs (99.1% of
296 California population) that had a population $\geq 1,000$ people and experienced at least one wildfire
297 smoke day (wildfire PM_{2.5} concentrations $\geq 15 \mu\text{g}/\text{m}^3$). In total, we observed 40,065 wildfire
298 smoke ZCTA-days in the 1,396 ZCTAs (0.6% of all ZCTA-days) during the study period. The
299 median number of ZCTA wildfire smoke days was 17 (1st and 3rd quartiles: 6 and 43), with
300 higher numbers in the Central Valley and Northern California (Figure 2). Most of the wildfire
301 smoke days occurred between June and November (96.7%), with more wildfire smoke days in
302 2007, 2008, 2017 and 2018 (Figure S1). We observed 18,049,797 non-scheduled respiratory
303 acute care utilizations in the study area between 2006 and 2019, with 75,175 occurring in
304 wildfire smoke days.



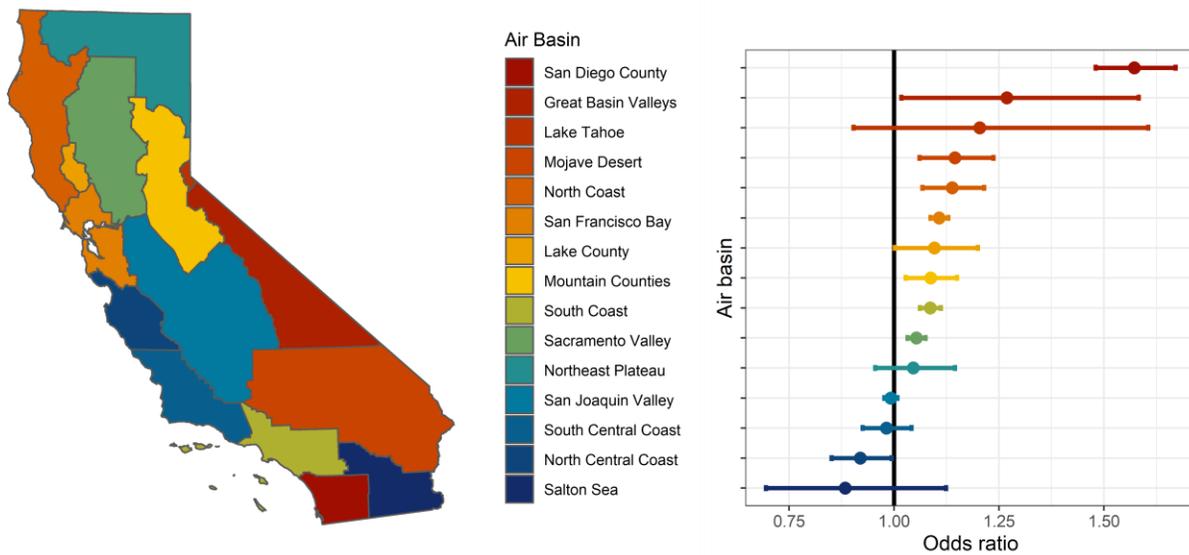
305

306 Figure 2. Spatial distribution of total ZCTA-level wildfire days in septiles between 2006-2019
307 among 1,396 ZCTAs included in the study. We considered wildfire days to be days with wildfire
308 $PM_{2.5}$ concentrations $\geq 15 \mu\text{g}/\text{m}^3$.

309 **3.2 Spatial heterogeneity of wildfire smoke day effects**

310 We first conducted a state-level analysis that did not consider spatial heterogeneity and
311 observed increased odds of respiratory acute care utilizations on wildfire smoke days (OR =
312 1.06, 95% confidence interval (CI): 1.05, 1.07), corresponding to 4122 (95% CI: 3491, 4747)
313 counts of acute care utilizations attributed to wildfire smoke between 2006 and 2019 (Table S1).
314 We then conducted three analyses considering spatial heterogeneity. In our air basin-level
315 analysis, the median OR point estimate was 1.09 (minimum and maximum: 0.88, 1.57) across
316 the 15 air basins (Table S1). We observed higher point estimates in San Diego as well as Great
317 Basin Valley, and lower point estimates in Salton Sea and North Central Coast (Figure 3). After
318 incorporating total acute care utilization counts during wildfire smoke days, air basins
319 experienced the highest acute health burden are San Francisco Bay and Sacramento Valley, with
320 1616 (95% CI: 1325, 1901) and 798 (95% CI: 490, 1099) counts of acute care utilizations
321 attributed to wildfire smoke between 2006 and 2019, respectively (Figure S2 and Table S1). In
322 our county-level analysis, the median point estimate for ORs was 1.06 (minimum and maximum:
323 0.45, 1.57) across 57 counties (Table S1). The direction of point estimates for air basins were
324 similar to those in their respective counties with a few exceptions (Kings County in the San
325 Joaquin air basin, Plumas County in Mountain Counties air basin) (Figure S3). San Diego
326 County and Los Angeles County experienced the highest acute care utilizations attributed to
327 wildfire smoke between 2006 and 2019 (Figure S4).

328 In the third analysis, we used a within-community matched design coupled with a spatial
329 Bayesian hierarchical model to assess spatial heterogeneity at the ZCTA level. We observed the
330 median point estimates for rate differences was -0.07 (minimum and maximum: -19.87, 29.61)
331 across 1,396 ZCTAs after accounting for spatial autocorrelation. We observed more spatial
332 heterogeneity in the ZCTA-level point estimates than across air basin or county, with higher and
333 more precise values observed in coastal metropolitan areas of San Diego, Mojave Desert, and
334 Great Basin Valleys, and lower and more precise values observed in the Salton Sea, North Coast
335 and Central Coast (Figure S5).



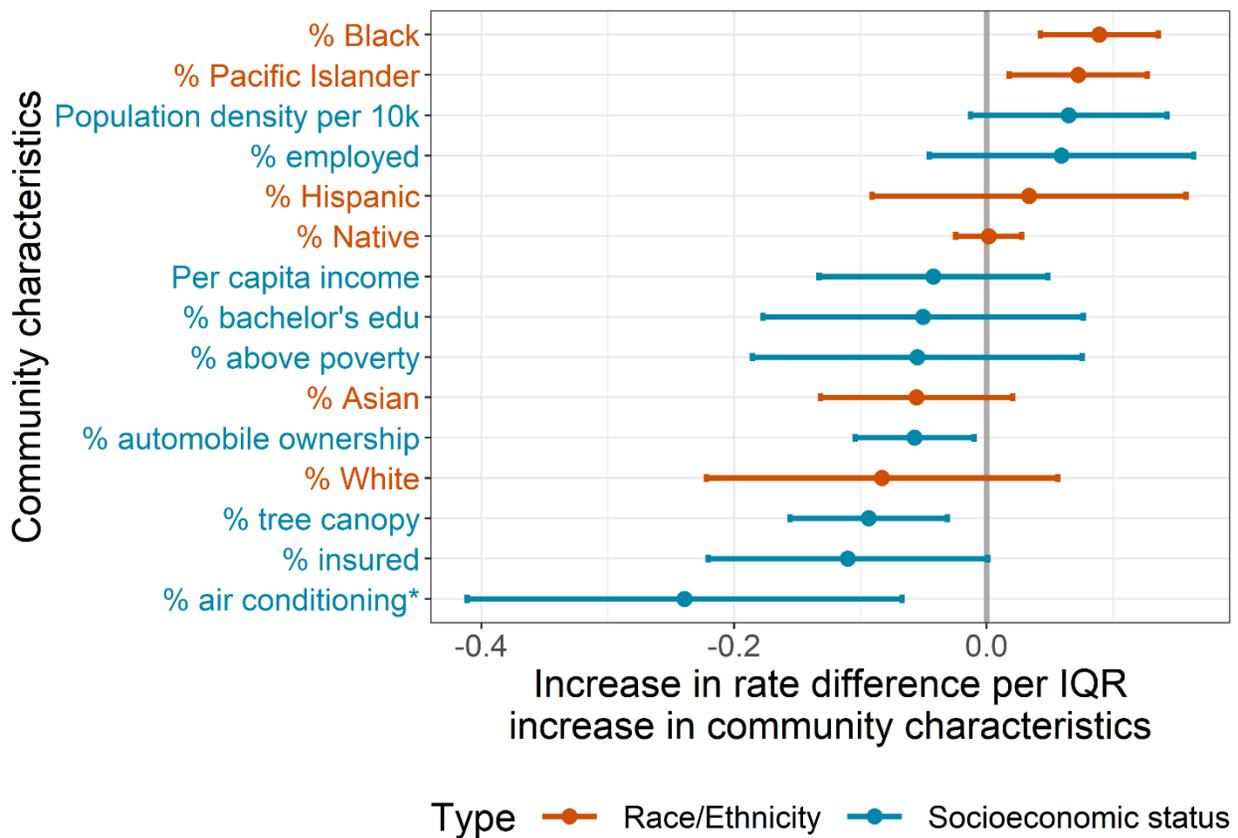
336
 337 Figure 3. The air basin specific effect estimates (odds ratio) of wildfire smoke day on same-day
 338 respiratory acute care utilization, 2006-2019. Left: spatial distribution of the point estimates;
 339 Right: point estimates and 95% confidence intervals. We employed conditional logistic
 340 regressions in a time-stratified case-crossover design, matching on ZCTA, day of week, month,
 341 and year.

342 **3.3 Effect modification of wildfire smoke day effects by community characteristics**

343 We evaluated effect modification by community characteristics as measured by 14
 344 variables in 1,296 ZCTAs with rate difference and complete community characteristics (Figure
 345 1). Analysis of A/C prevalence was only available in 1,122 ZCTAs. We found that a higher
 346 proportion of Black residents and Pacific Islander residents was associated with higher rate
 347 differences for respiratory acute care utilizations between wildfire smoke days and non-wildfire
 348 smoke days. ZCTAs with a higher proportion of white residents and Asian residents were
 349 associated with lower rate differences (Figure 4). Communities with a higher proportion of
 350 economically advantaged subpopulations were associated with lower rate differences for
 351 respiratory acute care utilizations between wildfire and non-wildfire smoke days. Effect
 352 modification was more pronounced for proportions of automobile ownership, tree canopy, and
 353 A/C prevalence (Figure 4).

354 **3.4 Sensitivity analyses**

355 At the state level, adding daily ambient temperature as a potential confounder in the
 356 evaluation of wildfire smoke day effect did not meaningfully change effect estimates, regardless
 357 of the form of temperature in the model (linear or nonlinear) (Figure S6). Our ZCTA-specific
 358 effect estimates were also robust to the choice of priors in spatial BHM (Figure S7). The effect
 359 modification results did not change meaningfully when utilizing ZCTA-level sociodemographic
 360 information from earlier years (2011-2015) among 1,235 ZCTAs (Figure S7).



361
 362 Figure 4. Effect modification of community characteristics on the effect of wildfire smoke (i.e.,
 363 days with wildfire $PM_{2.5} \geq 15 \mu g/m^3$) on same-day respiratory acute care utilization rate among
 364 1,296 CA ZCTAs. Race/ethnicity data was obtained from the 2010 US Decennial Census and
 365 socioeconomic information was obtained from the Healthy Place Index 3.0, and air conditioning
 366 was obtained from the California Residential Appliance Saturation Study survey.

367 *We included 1,122 ZCTAs for % air conditioning meta-regression because of data missingness.

369 **4 Discussion**

370 It is imperative to determine areas that experience the worse health outcomes after
371 wildfire PM_{2.5} exposure to reduce their associated burden. In our study, we found that wildfire
372 smoke days (i.e., days with wildfire PM_{2.5} $\geq 15 \mu\text{g}/\text{m}^3$) were associated with increased same-day
373 respiratory acute care utilizations in a statewide California model. However, the amplitude of
374 this relationship differed spatially across air basins, counties, and ZCTAs. Additionally, we
375 found that the impact of wildfire smoke days was worse for ZCTAs with higher proportions of
376 Black and Pacific Islander residents and less pronounced in more affluent areas with buffering
377 resources like tree canopy and A/C. Taken together, our study found that the health
378 consequences of wildfire PM_{2.5} exposure vary across space and community characteristics,
379 providing valuable evidence to guide the development of effective policies and the allocation of
380 resources.

381 Identifying areas experiencing the worse health effects is crucial for resource allocation,
382 public health response, and preparedness directives. In California, we observed higher health
383 impacts from wildfire PM_{2.5} in certain air basins including San Diego, Great Basin Valleys, and
384 Lake Tahoe. As air basins were created to originally manage and control non-wildfire pollution
385 emissions, wildfire PM_{2.5} and its health impacts may still differ within these air basins. As
386 climate change progresses, an estimated 82 million individuals in the Western US are predicted
387 to experience some wildfire smoke waves (at least two consecutive days with $>98^{\text{th}}$ quantile of
388 wildfire-specific PM_{2.5}) by the middle of the 21st century (Liu et al., 2016), making wildfire an
389 increasingly important source of total PM_{2.5}. Prior work found that PM_{2.5}-related health burdens
390 are under-estimated when wildfire PM_{2.5} is not explicitly considered in health impact
391 assessments (Darling et al., 2023). Thus, it is critical to revisit air pollution problems with an eye
392 to wildfire PM_{2.5} and to consider spatial differences in these exposures and effects.

393 When considering community characteristics, we found that the effects of wildfire PM_{2.5}
394 were worse for historically marginalized racial groups and less-resourced communities. These
395 community characteristics may also be key drivers of the observed spatial heterogeneity of
396 health effects. Prior work evaluating health disparities in the context of wildfire smoke observed
397 that socially and economically disadvantaged subgroups faced worse health effects (H. Chen et

398 al., 2021; Reid et al., 2016, 2023). In our study, we identified Black and Pacific Islander
399 residents as minoritized racial groups experiencing worse consequences at the same level of
400 exposure. Structural racism has given rise to disparities in environmental exposures, quality of
401 housing stock, access to economic and material resources, and baseline health (Bailey et al.,
402 2017). Such racially patterned disparities may worsen the health effects of exposure to wildfire
403 PM_{2.5}. We also found that ZCTAs with greater material resources had a dampened health
404 response to wildfire PM_{2.5} exposure. Access to material resources may indicate greater wealth,
405 which has been linked to improved capacity to mitigate and cope with wildfire PM_{2.5} (Burke et
406 al., 2022; deSouza & Kinney, 2021). Our findings contribute to prior research focused on
407 examining vulnerability to wildfire PM_{2.5} across subgroups (Vargo et al., 2023). Additionally,
408 current air quality management plans can make an effort to protect the most vulnerable. For
409 example, clean air centers in California may be expanded to serve additional communities of
410 color and economically disadvantaged areas (Bay Area Air Quality Management District, 2021;
411 US EPA, 2021).

412 This study had a few limitations. First, the modeled wildfire-specific PM_{2.5} (Aguilera et
413 al., 2023) may underestimate extreme exposure values given the training sample. However, our
414 use of a binary exposure definition dichotomized at $\geq 15 \mu\text{g}/\text{m}^3$ would correctly classify extreme
415 values as wildfire smoke days. The binary definition meant that we assumed health risks were
416 the same for any exposure level exceeding the threshold, and thus we could not capture any
417 exposure-response relationships that may occur particularly at the higher wildfire PM_{2.5} values
418 (Heft-Neal et al., 2023). Second, we utilized spatial units based on administrative borders, which
419 may not be the most relevant unit to assess spatial heterogeneity in the effect of wildfire PM_{2.5}
420 exposure. In addition, these units are of irregular shapes and sizes, with uneven population
421 densities across them. However, we centered our exposure estimates to the population-weighted
422 centroids of ZCTAs to improve the spatial alignment of health outcome and exposure. Another
423 limitation is that we assigned wildfire PM_{2.5} exposure at individuals' residential ZCTAs but
424 people may move across ZCTAs, which can result in exposure misclassification. However, for
425 days with high wildfire PM_{2.5}, individuals who can stay home would likely remain at indoors and
426 reduce the possibility of exposure misclassification.

427 With the increasing severity of wildfires, it is crucial to improve our understanding of
428 wildfire PM_{2.5}-related health impacts. We have a few recommendations for future research

429 endeavors in the area. First, we only evaluated spatial variation in the health impacts of wildfire
430 $PM_{2.5}$ in California, and future studies should extend to other US states and countries. Such
431 consideration could facilitate early identification of vulnerable areas and populations, and it can
432 guide subsequent targeted intervention efforts. Second, given the heterogeneity that we and
433 others have observed by community characteristics, future studies should identify the most
434 salient characteristics that modify the relationship between wildfire $PM_{2.5}$ and health. We tested
435 how community characteristics in isolation modified the effect of wildfire $PM_{2.5}$ on health but
436 these characteristics likely act synergistically, and future studies should endeavor to identify the
437 combination of characteristics that leads to the highest vulnerability. Third, we evaluated short-
438 term associations between wildfire $PM_{2.5}$ and health but climate change will likely lead to
439 increases in repeated wildfire $PM_{2.5}$ exposure and thus we must improve our understanding of the
440 health impacts of long-term wildfire $PM_{2.5}$ exposure. Last, we summarized ZCTA community
441 characteristics using a combination of Decennial Census Survey data and American Community
442 Survey-based Healthy Places Index data, which may miss important sub-populations. For
443 example, although the 2010 Census enumerated people in emergency and transitional shelters
444 (Bureau, 2021b), those experiencing homelessness—likely a highly vulnerable group (Ramin &
445 Svoboda, 2009)—may still be missed. We encourage an inclusive future research agenda that
446 prioritizes potentially vulnerable and understudied populations.

447 Most previous wildfire epidemiological studies assume that the effect of wildfire $PM_{2.5}$ is
448 consistent across geographies and populations. Our results suggest that instead, spatial
449 heterogeneity exists in the relationship between short-term wildfire $PM_{2.5}$ exposure and
450 respiratory acute care utilizations in California. We identified several community characteristics
451 that may have explained the differences observed; these included higher proportions of Black
452 and Pacific Islander populations and more affluent community. Allocating scarce resources
453 based on differential response to wildfire $PM_{2.5}$ could help reduce health disparities.

454

455

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458
459 **Availability statement:** the ZCTA-level wildfire-specific PM_{2.5} data used to identify wildfire
460 smoke days in the study are available at [https://github.com/benmarhnia-](https://github.com/benmarhnia-lab/Wildfire_PM25_California_ZIP)
461 [lab/Wildfire_PM25_California_ZIP](https://github.com/benmarhnia-lab/Wildfire_PM25_California_ZIP). The ZCTA-level community characteristics are available
462 from the Public Health Alliance of Southern California Healthy Places Index report version 3.0
463 (<https://www.healthyplacesindex.org>) and the 2010 Census
464 (<https://web.archive.org/web/20100320084325/http://2010.census.gov/2010census/>). The
465 respiratory acute care utilization data is not publicly available to protect patients' privacy but
466 access of the health outcome data could be requested directly at the California Department of
467 Health Care Access and Information website ([https://hcai.ca.gov/data-and-reports/research-data-](https://hcai.ca.gov/data-and-reports/research-data-request-information/)
468 [request-information/](https://hcai.ca.gov/data-and-reports/research-data-request-information/)). The analytic code is available at GitHub: [https://github.com/benmarhnia-](https://github.com/benmarhnia-lab/cal_wildfire_spatial.git)
469 [lab/cal_wildfire_spatial.git](https://github.com/benmarhnia-lab/cal_wildfire_spatial.git).

470
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472
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